Welfare Effects of Property Taxation

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Motivation

- Property taxes account for about one third of total capital tax revenues across Western countries (Zucman, 2015)
- We know little about the real effects of property taxation
- Particularly true for the welfare effects of the property tax (Oates and Fischel, 2016)
- OECD: "Taxes on immovable property among the least distortive tax instruments"
- If true, it is (mostly) about the distributional effects of property taxation

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WHO BEARS THE BURDEN OF THE PROPERTY TAX?

This Paper

- 1. Sufficient statistics approach to study welfar effects of property taxation
 - Add distributional perspective to efficiency-centered sufficient statistics approach
 - Approach nests standard textbook incidence model as well as full-fledged spatial equilibrium model
 - Allows for local public goods (benefit view) as well as the capital tax view
- 2. Reduced form evidence exploiting institutional setting of property taxation in Germany
 - Substantial variation: about 5,200 reforms of the property tax
 - Micro housing price data from ImmobilienScout24 combined with administrative data
- 3. New approach to simulate welfare effects
 - Simulate welfare effects across income distribution using household micro data
 - ▶ Household allowed to be various agents simultaneously: worker, renter, firm owner, landlord

Contribution to Literature

Preview of Findings

- Theory: Equity effects depend on a price responses on housing and labor market, amenity response captures efficiency margin
- 2. Empirics: Close to full pass-through of tax to total tax-inclusive rents in the medium-run
- 3. Simulation: Property tax is regressive, efficient use of tax revenues can reduce adverse distributional effects



Introduction

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Welfare Simulations

Basic Set-Up

- Derivation of distributional effects of property taxes using a sufficient statistics approach (Chetty, 2009)
- Household
 - Household derives utility from housing, a composite consumption good (numéraire), and local public goods/amenities of city c
 - Household pays the property-tax-inclusive, total rent q_i per square meter of housing: $q_c = p_c + t_c$
 - ► Household may receive income y_i from various sources; here: wages, profits, rental income: $y_i = w_c l_i + \pi_c e_i + p_c s_i$
 - Households maximize utility $u_i(x_i, h_i, l_i, e_i, s_i, A_c)$ s.t. to the budget constraint $q_c h_i + x_i = y_i$
- Using simple envelope conditions, we can derive welfare effects

Welfare Predictions

Proposition (Household Welfare)

The money-metric effect of a small increase in city c's property tax t_c on household i's utility is given by:

$$\Delta W_i := \frac{\mathrm{d}u_i/\mathrm{d}t_c}{\partial u_i/\partial x_i} = -h_i^* \frac{\mathrm{d}q_i^*}{\mathrm{d}t_c} + I_i^* \frac{\mathrm{d}w_c^*}{\mathrm{d}t_c} + e_i^* \frac{\mathrm{d}\pi_c^*}{\mathrm{d}t_c} + s_i^* \frac{\mathrm{d}\rho_c^*}{\mathrm{d}t_c} + \delta_i \frac{\mathrm{d}A_c^*}{\mathrm{d}t_c} \qquad \text{with } \delta_i = \frac{\partial u_i/\partial A_c}{\partial u_i/\partial x_i}.$$
(1)

Welfare depends on

- the pass-through of tax increases on total rent
- households pre-reform behavior (consumption and income streams)
- the change in local amenities
- ► Individual welfare effects ΔW_i can be aggregated to social welfare effects $\Delta W := \sum_i g_i \Delta W_i$ using marginal social welfare weights g_i (Saez and Stantcheva, 2016).

Efficiency and Distribution

- Most sufficient statistics approaches focus on efficiency (Kleven, 2021)
- Our proposition captures this margin via the amenity effect
- We extend common approach by introducing distributional effects, which are governed by price responses
- Adding the equity dimension is important as the proposition is implementable at household level using microdata
- Level of disaggregation depends on data quality and identification of price/amenity responses
- In our analysis: price/amenity responses at least at city-level, pre-reform quantities observed at individual level

Relation to Other Modeling Approaches

- ▶ Partial Equilibrium: Nests simplest textbook case: one renter, one landlord, no public goods
- Capital Taxes and General Equilibrium: Incorporates interaction with other sectors, such as capital market (Mieszkowski, 1972) or other markets, e.g. construction sector (Ahlfeldt et al., 2015)
- ► Tiebout Sorting: Accounts for benefit view through effect on local public goods (Tiebout, 1956, Hamilton, 1976)
- Quantative Spatial Equilibrium: Proposition is a short-cut to the welfare prediction of structural spatial equilibrium models

(Epple et al., 2001, Kline, 2010, Moretti, 2011, Redding and Rossi-Hansberg, 2017) • Spatial Eq. Results

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Local Property Taxes in Germany (*Grundsteuer B*)

- Set by each of the 11,500 German municipalities
- Most important tax on property in Germany, revenue 12 billion EUR in 2013, accounts for 21% of directly controlled local taxes Municipal Finances
- ► Tax Liability = Federal Tax Rate × Municipal Scaling Factor × Assessed Value

Local Property Tax Rate

- Assessed Value: fixed, determined by states in 1964, or at 1964 prices (West Germany)
- Federal Tax Rate: stable, differentiated by house type, 0.32% on average Federal Tax Rates
- Municipal Scaling Factor: varies, city councils votes on it annually <u>Tax Variation</u>
- \rightarrow Decreasing effective tax rates (wrt market value) in absence of reforms
- Only systematic reason for municipalities to raise property tax: improve long-run fiscal sustainability

(Romer and Romer, 2010, Guajardo et al., 2014, Alesina et al., 2015, Fuest et al., 2018, Lichter et al., 2024) 🕨 Effect on municipal finances

Statutory Incidence

Landlord to submit property tax to fiscal authorities

- By regulation property tax is included in annual ancillary cost (together with fees for garbage/sewage, fresh water, sometimes central heating costs)
- Renters pay a monthly advance on these ancillary cost
- Settlement of ancillary costs after the end of the year based on actual costs
- Effective statutory incidence is on the renter

Data Sources and Sample

Housing market microdata

- Source: ImmobilienScout24 provided by the research data center FDZ Ruhr at RWI
- Offered rents for new leases
- Main variable: total tax-inclusive rent per square meter after standard hedonic correction
- Other variables: Offered sales prices, main characteristics of buildings, location (municipality)

Administrative municipality data

- Sources: Various official statistics
- Data on property taxes, population, GDP, municipal expenditures, business profits, housing permits, unemployed, average wages, commuting zones, etc.

Estimation sample

Descriptive Statistics

- Focus on West Germany
- ▶ Years 2008–2015: Data availability housing price data and event study design
- Calculate average annual municipal house prices

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Empirical Model: Event Study Design

- Event study designs with 4 leads and 4 lags using a 1st-diff distributed-lag model (Suárez Serrato and Zidar, 2016)
- **•** Regress change in outcome $Y_{m,t}$ in muni *m*, year *t* on leads and lags of property tax rate changes:

$$\Delta \ln Y_{c,t} = \sum_{3}^{4} \gamma_j \Delta Property TaxRate_{c,t-j} + \psi \Delta X_{c,t} + \theta_{r(c),t} + \varepsilon_{c,t}$$
(2)

- First difference takes out municipal fixed effects μ_c
- ▶ MSA-by-year fixed effects $\theta_{r,t}$ to pick up secular shocks
- **•** Depending on the specification: time-varying municipal controls $X_{c,t}$
- Standard errors $\varepsilon_{m,t}$ clustered at municipal level
- **Dynamic treatment effects** β_i are calculated as follows (Schmidheiny and Siegloch, 2023)

$$\widehat{\beta}_{j} = \begin{cases} -\sum_{k=j+1}^{-1} \widehat{\gamma}_{k} & \text{if } -4 \leq j \leq -2 \\ 0 & \text{if } j = -1 \\ \sum_{k=0}^{j} \widehat{\gamma}_{k} & \text{if } 0 \leq j \leq 4. \end{cases}$$
(3)

Baseline Effect of Property Taxes on Total Rents (including Taxes)



Notes: This figure illustrates the estimated treatment effect of a one percentage point increase in the property tax rate on total tax-inclusive rents (in logs) relative to the pre-reform year. The underlying econometric model is described in equations (2) and (3). The specification accounts for lags in the local business tax rate and MSA-by-ger fixed freets. Observations are weighted by average population levels over the sample period. Vertical bars indicate 95% confidence intervals. Standard errors are robust to clustering at the municipality level. Dashed gray lines indicate the implied estimates for either zero or full shifting of taxes from landlords to tenants based on the corresponding average tax-to-rent ratio reported in the German Income and Expenditure Survey (EVS, 2013).

Mechanism: Reporting – Effects on Net-of-Tax Rents



Notes: This figure illustrates the estimated treatment effect of a one percentage point increase in the property tax rate on net-of-tax rents (in logs) relative to the pre-reform year. The underlying econometric model is described in equations (2) and (3). The specification accounts for lags in the local business tax rate and MSA-by-year fixed effects. Observations are weighted by average population levels over the sample period. Vertical bas indicate 95% confidence intervals. Standard errors are robust to clustering at the municipality level. Dashed gray lines indicate the implied estimates for either zero or full shifting of taxes from landlords to tenants based on the corresponding average tax-to-rent ratio reported in the German Income and Expenditure Survey (EVS, 2013).

Mechanism: Salience - Share of Ads including Property Tax Due



Notes: This figure illustrates the estimated treatment effect of a one percentage point increase in the property tax rate on share of ads that includes information on the property taxes to be paid relative to the pre-reform year. The underlying econometric model is described in equations (2) and (3). The specification accounts for lags in the local business tax rate and MSA-by-year fixed effects. Observations are weighted by average population levels over the sample period. Vertical bars indicate 95% confidence intervals. Standard errors are robust to clustering at the municipality level.

Probing Identification

- 1. Regional confounders: Account for local shocks more flexibly, e.g., CZ-by-year fixed effects 💽
- Municipal confounders: No pre-trends in local business cycle variables
 ; control for local business cycles
- 3. Selection on unobservables: Calculate Oster (2019) bounds 💽
- 4. IV strategy: Use state-level changes in reference tax rate VIV Model First Stage Reduced Form
- 5. Lag lengths: Test long-run effects, change assignment to T and C group 💽
- 6. Large event : Event dummy specifications with (large) tax increases 💽
- 7. Heterogeneous treatment effects: Apply estimators of Sun and Abraham (2021), de Chaisemartin and D'Haultfœuille (2020, 2024) 💌

Probing Identification



Notes: This figure presents the results for seven of our the eight identification checks outlined before. Estimate depict the estimated treatment effect of a one percentage point increase in the property tax rate on total tax-inclusive errors (in log) relative to the pre-efform ways. The baseline estimates are shown in mice, results from alternative specifications are depicted in blue. The Parel A presents tummary stimated of per-treatment trends, i.e., the warge coefficient in the four years prior to a tax reform. Panel B shows the long-run effect measured as the average estimate are shown in the four years prior to a tax reform. Panel B shows the long-run effect measured as the average estimate of the third and fourth lag in the poperty tax rate. All regressions also account for leads and lags in the local basiness tax rate. Observations are weighted by average population levels over the sample period. Horizontal bars indicate 9% conditione: interval. Standed errors are robust to culturely low level.

Further Robustness Checks

- 1. Accounting for local business tax changes
 Results
- 2. Fixed effects instead of first differences
 Results
- 3. Minimum number of ads per municipality-year cell
 Results
- 4. Weights: population, ads **Presults**
- 5. Including East Germany Results
- 6. Hedonic correction Results

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Data and Empirical Implementation

- ▶ We use Proposition 1 to simulate the welfare effects of a property tax change on the household level
- Rely on household microdata (EVS, German Income and Expenditure Survey) to observe pre-reform quantities (incomes and expenditures) at household-level
- Use long-run reduced form effects on total and net-of-tax rents presented above
 - Introduce heterogeneity by estimating the rent effect for different apartment types (size, age)
 - Account for owner-occupied housing by simulating changes in tax payments based on imputed rents
- Estimate (heterogeneous) wage and business income effects using administrative municipal data
- Estimate amenity effect using the change in property tax revenue
- Evalibrate local public good preference δ to different benchmark values

Estimates of Prices Effects for Simulation

	All Rental Ads		By Apartment Size			By Construction Year		
	Total	Net	Below	Between	Above	Before	Between	After
	Rent	Rent	60m ²	60–80m²	80m²	1949	'49–'90	1990
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Long-Run Effect	0.034***	0.001	-0.000	0.033**	0.050***	0.041	0.017	0.003
	(0.011)	(0.011)	(0.011)	(0.011)	(0.011)	(0.032)	(0.015)	(0.022)
Muni.×Year Obs.	23,303	23,236	23,303	23,303	23,303	23,303	23,303	23,303

Panel A – Effects on Total and Net Rents

Panel B - Effects on Wages, Profits, and Public Goods

	Wage Earnings					Public Goods		
	Average Wage (9)	P25 Wage (10)	P50 Wage (11)	P75 Wage (12)	Business Profits (13)	Prop Tax Reven. (14)	Munic. Expenses (15)	
Long-Run Effect	-0.007 (0.011)	-0.056 (0.047)	-0.021 (0.018)	0.013 (0.010)	-0.035 (0.132)	0.508*** (0.026)	-0.067* (0.037)	
Muni.×Year Obs.	37,781	38,782	38,782	38,782	8,347	41,190	41,286	

Notes: This table depicts long-run price effects for various outcomes (in logs) in response to a one percentage point increase in the property tax rate relative to the pre-reform year. The underlying econometric model is described in equations (2) and (3). The specification also accounts for leads and lags in the local business tax rate and MSA-by-year fixed effects. Observations are weighted by average population levels over the sample period. Full event-study estimates are depicted in Appendix Figure 29. Standard errors in parentheses are robust to clustering at the municipality level (* p < .1, ** p < .05, *** p < .01.) See Appendix ?? for detailed information on all variables.

Simulation Results: Welfare Effects of Property Taxes Due to Price Changes



Meter: This figure Illustrates the relative walfare consequences of a one percentage point increase in the local poperty tax over the household communptics distribution (in percent). We calculate relative walfare tissues as monos metric edity changes in europer year of ideed by annual household communptics of the household and remembranes can percentage and the house appendix and remembranes are percented in Paral A, we introduce more hereageneity in our walfare simulations steps by targets as we move to Pauld D, Paral A reports walfare in a partial-aquilibrium model with toe representative agents (landiced and remembrane). Pauld B additionally accounts for differences in the numbers of the house agents, their difference in postionic in the community of the house appendix percentage postion in the community of the house appendix percent particle and remembranes are presented (in a larget a data) water (larget agent administic) and percent). Pauld B additionally accounts for differences in the numbers of the house appendix percent particle and remembranes (larget agent administic) and percent particle and the step and percent particle and remembranes are presented by the step and percent particle and remembranes are percented and remembranes and percent particle and remembranes and percent particle and remembranes and percent particle and remembranes areas

▶ absolute welfare changes ↓ ▶ Cls for GE models

Simulation Results: Welfare Effects of Property Taxes including Public Goods



Notes: This figure illustrates the estimated treatment effect of a one procentage point increase in the property tax rate on property tax revenues (in logs) relative to the pre-reform year (Panel A). The underlying commentic model and described in equations (2) and (3). The sequeficiation also accounts for lads and lags in the local business tax rate and MSA-byser rate of effects. Observations are weighted by average population levels over the sample period. Vertical bars indicate 9% confidence intervals. Standard errors are rebut to clustering at the municipality levels. The dashed gray line indicates the implied estimates for tax revenues a business. Desting the sample period. We considered and the same resonance is the indicate start in the local period in the case relative to a source population of the same resonance is the indicate start in the local period in the case relative to a source provide and the same relative to a source population of the same revenues. The same relative to a source provide communities of the same revenues. The same relative to clustering at the municipality levels in the local period in the case relative to a source provide more dashed communities of the same relative to a source provide more dashed communities of the same relative to a source provide more dashed on average for the same relative to a source provide more dashed on the case relative to a source provide more dashed on the case relative to a source provide more dashed on the case of the same relative to a source provide more dashed on the case of the same relative to a source provide more dashed on the same relative to a the same relative to a source provide more dashed on the same relative to a the relative same relative to a source provide more dashed on the same relative to a source provide more dashed on the same relative to a source provide more dashed on the same relative to a source provide more dashed on the same relative to a source provide more dashed on the same relative to a source provide more dashed on

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- 2. The property tax can be approximately welfare-neutral if public goods are provided

"ARE LOCAL PROPERTY TAXES REGRESSIVE, PROGRESSIVE, OR WHAT?"

(Oates and Fischel, 2016)

- 1. The property tax is not progressive
- 2. The property tax can be approximately welfare-neutral if public goods are provided efficiently and individuals indifferent b/w marginal private and public good
- 3. The property tax is, hence, regressive in most settings

Appendix

Contributions to the literature

1. Theoretical approach combines ...



Chetty (2009), Kleven (2021)

... spatial equilibrium and local labor market models

Rosen (1979), Roback (1982), Albouy (2009), Glaeser (2009), Kline (2010), Moretti (2011), Kline and Moretti (2014), Ahlfeldt et al. (2015), Diamond (2016), Suárez Serrato and Zidar (2016), Fajgelbaum et al. (2019), Brülhart et al. (2024)

... property tax views 💌

Marshall (1890), Edgeworth (1897), Bickerdike (1902), Simon (1943), Tiebout (1956), Oates (1969), Mieszkowski (1972), Hamilton (1976), Mieszkowski and Zodrow (1989), Fischel (1992), Zodrow (2001b,a)

2. Reduced-form evidence on ...

... housing market effects of property taxes

Orr (1968, 1970, 1972), Heinberg and Oates (1970), Hyman and Pasour (1973), Dusansky et al. (1981), Carroll and Yinger (1994), Palmon and Smith (1998), de Bartolomé and Rosenthal (1999), Buettner (2003), Lyytikäinen (2009), Ferreira (2010), Shan (2010), Lutz (2015)

… local public goods/amenities

Bradbury et al. (2001), Bayer et al. (2007), Cellini et al. (2010), Boustan (2013), Suárez Serrato and Wingender (2016), Schönholzer and Zhang (2017), Fajgelbaum et al. (2019), Brülhart et al. (2024)

3. Welfare/distributional effects ...

... over the income distribution

... complementing structural heterogeneous agent models

Suárez Serrato and Zidar (2016), Diamond (2016), Fajgelbaum and Gaubert (2020), Brülhart et al. (2024)

Property tax views

Traditional view (Edgeworth, 1897, Simon, 1943, Netzer, 1966)

- Tax introduced in single municipality, perfectly elastic capital supply
- Tenants bear the full tax burden of property taxation

Capital tax view (Mieszkowski, 1972, Mieszkowski and Zodrow, 1989)

- Extends "old view" with Harberger general equilibrium model
- Capital owners bear the national average burden of property taxes

Benefit tax view (Tiebout, 1956, Oates, 1969, Hamilton, 1976, Fischel, 1992)

- Property taxes finance local public goods, work as a non-distortive fee
- Households are mobile and choose "optimal municipality"

Hard to provide systematic empirical evidence for different views (Oates and Fischel, 2016)

- General equilibrium aspects hard to pin down empirically
- Empirics focused on partial analyses (as in corporate tax literature)

Proof of Proposition **??** (Household Welfare)

• Using indirect utility $v_i(r_i^{C*}, y_i^*, g_c^*)$, we are interested in a small increase of the property tax t_c

$$\frac{\mathrm{d}\mathbf{v}_i}{\mathrm{d}\mathbf{t}_c} = \frac{\partial \mathbf{v}_i}{\partial \mathbf{r}} \frac{\mathrm{d}\mathbf{r}_i^{C*}}{\mathrm{d}\mathbf{t}_c} + \frac{\partial \mathbf{v}_i}{\partial \mathbf{y}} \frac{\mathrm{d}\mathbf{y}_i^*}{\mathrm{d}\mathbf{t}_c} + \frac{\partial \mathbf{v}_i}{\partial \mathbf{g}} \frac{\mathrm{d}\mathbf{g}_c^*}{\mathrm{d}\mathbf{t}_c} \tag{4}$$

Envelope theorem and Roy's identity allows to simplify the first term on the right-hand side

$$\frac{\mathrm{d}\mathbf{v}_i}{\mathrm{d}\mathbf{t}_c} = -\mathbf{h}_i^* \frac{\partial \mathbf{v}_i}{\partial \mathbf{y}} \frac{\mathrm{d}\mathbf{r}_i^{C*}}{\mathrm{d}\mathbf{t}_c} + \frac{\partial \mathbf{v}_i}{\partial \mathbf{y}} \frac{\mathrm{d}\mathbf{y}_i^*}{\mathrm{d}\mathbf{t}_c} + \frac{\partial \mathbf{v}_i}{\partial \mathbf{g}} \frac{\mathrm{d}\mathbf{g}_c^*}{\mathrm{d}\mathbf{t}_c}.$$
(5)

> Normalizing by the marginal utility from income, we get the money-metric welfare effect

$$\Delta W_i = \frac{\mathrm{d} v_i / \mathrm{d} t_c}{\mathrm{d} v_i / \mathrm{d} y} = -h_i^* \frac{\mathrm{d} r_i^{c\,*}}{\mathrm{d} t_c} + \frac{\mathrm{d} y_i^*}{\mathrm{d} t_c} + \delta_i^g \frac{\mathrm{d} g_c^*}{\mathrm{d} t_c} \qquad \text{with } \delta_i^g = \frac{\partial v_i / \partial g}{\partial v_i / \partial y},\tag{6}$$

Comparison to spatial equilibrium model

Multiply with marginal utility from income:

$$\frac{\mathrm{d}\mathbf{v}_i}{\mathrm{d}t_c} = -\frac{\partial\mathbf{v}_i}{\partial y_i^*} h_i^* \frac{\mathrm{d}\mathbf{r}_i^{C*}}{\mathrm{d}t_c} + \frac{\partial\mathbf{v}_i}{\partial y_i^*} \left(\frac{\mathrm{d}y_i^{W*}}{\mathrm{d}t_c} + \frac{\mathrm{d}y_i^{F*}}{\mathrm{d}t_c} + \frac{\mathrm{d}y_i^{L*}}{\mathrm{d}t_c} \right) + \frac{\partial\mathbf{v}_i}{\partial g_c^*} \frac{\mathrm{d}g_c^*}{\mathrm{d}t_c}.$$

Impose assumptions that are necessary in structural models, e.g.

- Marginal utility of private income: $\frac{\partial V_i}{\partial y_i^*} = (1 \delta)$
- Marginal utility of public good: $\frac{\partial V_i}{\partial g_c^*} = \delta$

• Worker *i* is neither firm owner nor landlord $(y_i^F = y_i^L = 0)$

$$\frac{\mathrm{d}V_i}{\mathrm{d}t_c} = -(1-\delta)\left[h_i^*\frac{\mathrm{d}r_i^{C*}}{\mathrm{d}t_c} - \frac{\mathrm{d}y_i^{W*}}{\mathrm{d}t_c}\right] + \delta\frac{\mathrm{d}g_c^*}{\mathrm{d}t_c}.$$

which is equivalent to formula for renter welfare in a fully specified structural spatial model

(Kline and Moretti, 2014, Suárez Serrato and Zidar, 2016)

$$\frac{dW^{H}}{d\ln\tau_{c}} = -(1-\delta)\left[\alpha + \alpha\frac{d\ln r_{c}^{H*}}{d\ln\tau_{c}} - \frac{d\ln w_{c}^{*}}{d\ln\tau_{c}}\right] + \delta\frac{d\ln G_{c}^{*}}{d\ln\tau_{c}}$$
Municipal Revenues and Local Public Goods

- In 2013, the average (median) annual revenues of municipalities was 2,691 (2,353) euro per capita
- 28% of the revenues are coming from local taxes (thereof business tax 79% and property tax 21%)
- Share of total municipal revenues from property taxes 5% (US: 12%)
- Local public goods/services less important in Germany (no schools, police, highways)
- About 80% of total municipal expenditures are spent on keeping up the usual administrative duties (employees, buildings, subsidize public firms, welfare costs); 15% of the expenditures are used for investment projects

Federal Tax Rates for West Germany (in %)



Variation in Local Property Taxes



Notes: The left panel of this figure shows the local property tax rates in 2013 for all West German municipalities, assuming a federal tax rate of 0.32 percent. The right panel depicts the number of local property tax changes by municipality in the period 1990–2018. Municipalities are grouped into population-weighted quintiles and shaded according to the tax rate or the number of tax changes, respectively. Jurisdictional boundaries are as of December 31, 2015. Gray lines indicate federal state borders. White areas indicate unpopulated unincorporated areas (gemeindefreie Gebiete). See Appendix ?? for detailed information on all variables. Mags: © Geossis-DE / BKC 2019.

Simple Example for Rents

▶ Introduction of a 20 EUR property tax on apartments in Cologne

	Pre Tax	After Tax Introduction			
Invoice/rent bill (in EUR)		Full pass-through to tenant	No pass-through to tenant		
Net Rent (producer price)	400	400	380		
Operating costs	30	30	30		
Property tax	0	20	20		
Gross Rent (consumer price)	430	450	430		

Descriptive Statistics <

Variable	Mean	SD	P10	P25	P50	P75	P90
Panel A – Housing Prices							
Average Rent (in €/m²)	9.09	2.06	7.03	7.71	8.57	10.08	11.73
Panel B – Fiscal Variables							
Local Property Tax Rate (in %)	1.30	0.30	0.96	1.09	1.28	1.48	1.71
Standard Tax Rate (in %)	0.97	0.28	0.59	0.70	0.99	1.22	1.32
Local Business Tax Rate (in %)	13.85	1.89	11.55	12.25	13.82	15.40	16.62
Expenditures per Capita (in €)	2,519.20	1,212.27	1,488.05	1,773.01	2,263.85	3,047.11	3,799.63
Revenues per Capita (in €)	2,621.72	1,351.18	1,518.37	1,817.87	2,345.59	3,188.67	3,930.63
Property Taxes per Capita (in €)	155.11	48.96	94.80	119.89	150.23	183.73	221.41
Panel C – Economic Indicators							
Average Daily Wages (in €)	71.67	14.83	53.67	61.94	70.62	80.53	91.35
Business Profits per Capita (in €)	4,215.87	5,390.21	1,362.86	2,118.60	3,256.91	5,072.34	7,454.63
Local Population Levels	153,801	283,491	5,428	11,960	32,030	124,577	548,547
Local GDP per Capita (in €)	37,753	16,385	24,088	27,247	32,599	41,573	59,925
Local Unemployment Rate (in %)	7.03	2.92	3.70	4.74	6.55	8.66	11.25

Notes: This table provides descriptive statistics for the baseline estimation sample. Means are weighted by average population levels over the sample period. See Table ?? for detailed information on all variables.

Importance of Regional Confounders <



Notes: This figure illustrates the estimated treatment effect of a one percentage point increase in the property tax rate on total tax-inclusive rents (in logs) relative to the pre-reform year using various alternative regional time trend specifications. The underlying econometric model is described in equations (2) and (3). The specification also accounts for leads and lags in the local lausines tax rate and year food effects at various regional levels (see legend). Observations are weighted by average population levels over the sample period. Vertical law indicate 9% confinence intervals. Standard errors are robust to clustein; at the municipality level.

Effect of Property Taxes on Local Business Cycle Outcomes



Notes: This figure illustrates the estimated treatment effect of a one percentage point increase in the property tax rate on local business cycle measures relative to the pre-reform year. The underlying econometric model is described in equations (2) and (3). All specifications also account for leads and lags in the local business tax rate and MSA-by-year fixed effects. Observations are weighted by average population levels over the sample period. Vertical basi indicate 95% confidence intervals. Standard errors are robust to clusteing at the municipabily level.

Accounting for Local Business Cycle Confounders



Notes: This figure illustrates the estimated treatment effect of a one percentage point increase in the property tax rate on total tax-inclusive rents (in logs) relative to the pre-reform year using different sets of control variables for local business cycles. The underlying econometric model is described in equations (2) and (3). All specifications also account for leads and lags in the local business tax rate and MSA-by-year fixed effects. Observations are weighted by average population levels over the sample period. Vertical bars indicate 95% confidence intervals. Stander dervours are hout to clusting at the multipositive level.



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Table: Bounded Estimates Following Oster (2019)

	(1) Uncontrolled Estimate	(2) Controlled Estimate	(3) Bounded Estimate
Panel A – Using Contemporaneous Controls			
Medium-Run Effect	0.034***	0.033***	0.031
	(0.011)	(0.011)	
Number of Observations	23,303	23,295	
Adjusted <i>R</i> -squared	0.003	0.004	
Panel B – Using Lagged Control Variables			
Medium-Run Effect	0.034***	0.034***	0.036
	(0.011)	(0.011)	
Number of Observations	23,303	23,294	
Adjusted <i>R</i> -squared	0.003	0.003	

Notes: This table illustrates the bounded estimates for the treatment effect of a one percentage point increase in the property tax rate on total rents (in logs) relative to the pre-reform year. Bounds have been obtained using the approximation in Oster (2019) and calibrating $\delta = 1$ and $R_{max}^2 = 1.3 \cdot R_{controlled}^2$. Panel A presents bounds using contemporaneous business cycle control variables (population, unemployment, county-level GDP) for the controlled model. Panel B, relies on the same control variables (larged by two wars. The underlying the same control variables (larged by two wars.

Instrumental variables approach 🜑

- Each state s set a standard tax rate (Fiktiver Hebesatz), which is increasing over time
- One can show that municipalities have incentive to (Buettner, 2003, Egger et al., 2010, Baskaran, 2014, Rauch and Hummel, 2016)
 - increase own property tax rate when standard rate increase
 - incentive is increasing in the difference between new standard rate and own tax rate
- We translate this into the following instrumental variable

$$IV_{m,t} = StandardTaxRateIncrease_{s,t} \cdot \frac{StandardTaxRate_{s,t} - PropertyTaxRate_{m,t-1}}{PropertyTaxRate_{m,t-1}}$$
(7)

where $StandardTaxRateIncrease_{s,t}$ is a dummy indicating a standard rate hike in state s and year t

The first stage equation is given by

$$\Delta Property TaxRate_{m,t} = \sum_{j=-\underline{j}+1}^{\overline{j}} \eta_j IV_{m,t-j} + \delta \Delta X_{m,j} + \zeta_{r,t} + \epsilon_{m,t},$$
(8)



First Stage Results Using Standard Tax Rate Increases



Figure: First Stage Results Using Standard Tax Rate Increases

Notes: This figure shows the estimated treatment effects of state-level reforms in the standard tax rate on local property tax rates using alternative specifications to account for regional confounders. Formally, we

Sebastian Siegloch (Cologne)

Effect of Standard Tax Rate Changes on Gross Rents



Netex: This figure illustrates the setimated reduced-form effect of the standard tax rate IV defined in equation (?) on total tax-inclusive rents (in (logs) relative to the year before a standard tax rate increase. The underlying economics model is analogued to equation (2) and (2). The specification also accounts for teaks and lags in the local banies tax rate and MSA-by-year field effect. Observations are weighted by average population lovels over the sample period. Municipalities in the states Baden-Württemberg and Saarland are occluded from the estimation sample. Verical ban indicate 95% confidence intervals. Standard errors are notes to clustering at the municipality level.

Identification





Notes: This figure illustrates the estimated treatment effect of a one percentage point increase in the property tax rate on total tax-inclusive rents (in logs) relative to the pre-reform year for different lag lengths. The underlying econometric model is described in equations (2) and (3). All specifications account lags in the local business tax rate and MSA-by-year fixed effects. Observations are weighted by average population levels over the sample period. Vertical bars indicate 95% confidence intervals. Standard errors are robust to clustering at the municipality level.

Effect of Property Taxes on Gross Rents: Increase Dummies (re-scaled)



Notes: This figure illustrates the estimated treatment effect of property tax changes on gross rents (in logs) relative to the pre-reform year using alternative definitions of event study variables. The baseline specification scales any change in property tax rates with the size of the tax change. The other specifications shown in the legend use simply event indicators for any tax increases (grazer or equal to the P50, P75 or P90 of the tax increase (grazer or equal to the P50, P75 or P90 of the tax increase (grazer or equal to the P50, P75 or P90 of the tax increase (grazer or equal to the P50, P75 or P90 of the tax increase) ender the underlying econometric model is described in equations (2) and (3). All specifications also account for leads and lags in the local business tax rate and MSA-by-year fixed effects. Observations are weighted by average population levels over the sample period. Vertical bars indicate 95% confidence intervals. Standard errors are robust to clustering at the municipality level.

Effect of Property Taxes on Gross Rents: Heterogeneous Treatment Effects 🜑



Notes: This figure illustrates the estimated treatment effect of large property tax increases on gross rents (in (log) relative to the pre-reform year using alternative definitions of event study variables. Large property tax changes are defined as being above the median of the property tax distribution. The sample is restricted to municipalities with either on one large change. The underlying econometric model is similar to equation (2) with the exception that we use a dichotomous treatment variable indicating a tax change instead of a continuous one. The baseline two-way fixed effects (TWPE) model is contrasted with models that account for heterogeneous treatment effects as indicated in the legend: CD2020 stands for de Chaisemartin and D'Haultfocuille (2020), SA2021 for Sun and Abraham (2021); and BJS 2022 for Borusyak et al. (2023). All specifications also account for leads and lags in the local business tax rate and MSA-by-year fixed effects. Observations are weighted by average population levels over the sample period. Vertical bars indicate 95% confidence intervals. Standard errors are robust to clustering at the municipality level.

Robustness: Accounting for Local Business Tax Changes



Notes: This figure illustrates the estimated treatment effect of a one percentage point increase in the property tax rate on total tax-inclusive rents (in logs) relative to the pre-reform year using different sets of control variables for local business cycles. The underlying econometric model is described in equations (2) and (3). All specifications also account for leads and lags in the local business tax rate and MSA-by-year fixed effects. Observations are weighted by average population levels over the sample period. Vertical bars indicate 95% confidence intervals. Stander dress are enbut to clustering at the multipolity level.

Robustness: First Differences vs. Fixed Effects



Notes: This figure illustrates the estimated treatment effect of a one percentage point increase in the property tax rate on total tax-inclusive rents (in logs) relative to the pre-reform year using different sets of control variables for local business cycles. The underlying econometric model is described in equations (2) and (3). All specifications also account for leads and lags in the local business tax rate and MSA-by-year fixed effects. Observations are weighted by average population levels over the sample period. Vertical bars indicate 9% confidence intervals. Standard energy are Routs to Calenting at the multipolity level.

Robustness: Number of ads per muni-year



Notes: This figure illustrates the estimated treatment effect of a one percentage point increase in the property tax rate on total tax-inclusive rents (in logs) relative to the pre-reform year using different sets of control variables for local business cycles. The underlying econometric model is described in equations (2) and (3). All specifications also account for leads and lags in the local business tax rate and MSA-by-year fixed effects. Observations are weighted by average population levels over the sample period. Vertical bars indicate 95% confidence intervals. Stander dervours are hout to clusting at the multipositive level.

Robustness: Weights **C**



Notes: This figure illustrates the estimated treatment effect of a one percentage point increase in the property tax rate on total tax-inclusive rents (in logs) relative to the pre-reform year using different sets of control variables for local business cycles. The underlying econometric model is described in equations (2) and (3). All specifications also account for leads and lags in the local business tax rate and MSA-by-year fixed effects. Observations are weighted by average population levels over the sample period. Vertical bars indicate 99% confidence intervals. Standard errors are robust to clustering at the municipality level.

Robustness: Including East Germany



Notes: This figure illustrates the estimated treatment effect of a one percentage point increase in the property tax rate on total tax-inclusive rents (in logs) relative to the pre-reform year using different sets of control variables for local business cycles. The underlying econometric model is described in equations (2) and (3). All specifications also account for leads and lags in the local business tax rate and MSA-by-year fixed effects. Observations are weighted by average population levels over the sample period. Vertical base indicate 95% confidence intervals. Stander dervours are hown to clusting at the multipositive level.

Robustness: Hedonic Correction of Rents <



Notes: This figure illustrates the estimated treatment effect of a one percentage point increase in the property tax rate on total tax-inclusive rents (in logs) relative to the pre-reform year using different sets of control variables for local business cycles. The underlying econometric model is described in equations (2) and (3). All specifications also account for leads and lags in the local business tax rate and MSA-by-year fixed effects. Observations are weighted by average population levels over the sample period. Vertical bars indicate 95% confidence intervals. Standard energy are houst to clusting at the multipolity level.

Robustness Checks **CO**



Estimated Treatment Effects

Reduced-form effects on Wages



Notes: This figure illustrates the estimated treatment effect of a one percentage point increase in the property tax rate on various measures of municipal wages relative to the pre-reform year. The underlying econometric model is described in equations (2) and (3). The specification accounts for lags in the local business tax rate and MSA-by-year fixed effects. Observations are weighted by average population levels over the sample period. Vertical bars indicate 95% confidence intervals. Standard errors are robust to clustering at the municipality level.

Reduced-Form Effects on Business Incomes



Notes: This figure illustrates the estimated treatment effect of a one prenatage point increase in the property tax rate on business income setup to the pre-perform year. The underlying treates of the estimated treatment of the property and the property tax rate on business income setup to the pre-perform MSA-by-year fixed effects. Observations are weighted by average population levels over the sample period. Vertical bars indicate 95% confidence intervals. Standard errors are inducted cutatring at the municipability level.

Reduced-form effects on Local Public Goods



Notes: This figure illustrates the estimated treatment effect of a one percentage point increase in the property tax rate on two measures of local public goods (the property tax revenues and the municipal expenses) relative to the pre-reform year. The underlying econometric model is described in equations (2) and (3). The specification accounts for lags in the local business tax rate and MSA-by-year fixed effects. Observations are weighted by average population levels over the sample period. Vertical bars indicate 95% confidence intervals. Standard errors are robust to clustering at the municipality level.

Effect of Property Taxes on Municipal Finances <



Notes: This figure illustrates the estimated treatment effect of a one percentage point increase in the property tax rate on municipal revenues and expenditures relative to the pre-reform year. The underlying econometric model is described in equations (2) and (3). All specifications also account for leads and lags in the local business tax rate and MSA-by-year fixed effects. Observations are weighted by average population levels over the sample period. Vertical bars indicate 95% confidence intervals. Standard errors are robust to clusteing at the municipality level.

Simulation Results: Absolute Welfare Effects of Property Taxes



Meters: This figure illustrates average absolute moneymetric utility changes for each percentile along the comuniption distribution (in euro per year). Starting from a stylized benchmark case presented in Panel A, we introduce more heterogeneity in our weights immulation starsplaystarp as we more to Panel-D. Panel A Pener to Panels - Panel A percentile approximation and percentage approximation percentage approximation and percentage approximation approximation and percentage approximation approxim

Additional Inference Tests **CO**



Notes: This figure illustrates the relative welfare consequences of a one percentage point increase in the local property tax over the household consumption distribution (in percent). We calculate relative welfare losses as money-metric utility change in euro per year divided by annual household consumption. Absolute welfare changes are depicted in Appendix (pure 32. This graph complements Panels C and D of Figure 37 by depicting empirical 90 and 90°, conditiones bounds using LOUD botters prelications (calaber and lighter blue hadde areas). Set the toxing for a first blue hadde areas). Set the toxing in the first blue hadde areas).

- Relatively large standard errors on business income estimate (small sample) leads to large CIs in simulation
- Bootstrap inference on distributional effects: 987/1000 replications yield a regressive pattern (positive slope) when adding GE effects

Model – Workers (=Tenants)

- Choose city c where they supply one unit of labor, earning wage w_c
- Maximize over (homogeneous) housing h & composite good x
- Pay rental price r and local property tax $\tau_c = (1 + t_c)$
- ▶ Utility also depends on a city-specific consumption amenities with exogenous part, \overline{A}_c and an endogenous local public good G_c , financed out of property tax revenues
- Indirect worker utility:

$$V_{ic}^{H} = a_{0} + (1 - \delta) \left[\ln w_{c} - \alpha \ln r_{c}^{H} - \alpha \ln \tau_{c} \right] + \ln \overline{A}_{c} + \delta \ln G_{c} + \ln e_{ic}$$

▶ Have a individual-location specific idiosyncratic preference $e_{ic} \sim$ EV-I with scaling parameter σ^W

> Set-up gives rise to conditional logit model, which yields labor supply N_c^S and housing demand H_c^D

Model – Firms 🜑

- Firms operate under monopolistic competition and locate in a city c (Suárez Serrato and Zidar, 2016)
- ▶ Produce a trabable output good Y_{jc} using labor N_{jc} and commercial floor space M_{jc}
- Firm profits in city c depend on
 - city-specific production amenities B_c
 - Firm-city specific productivity $\omega_{jc} \sim \text{EV-I}$ with scaling parameter σ^F
- Firms pay a property tax $\kappa \tau_c$, with $\kappa \geq 0$

Firm value:

$$\Pi_{jc}^{F} = b_0 + \ln B_c - \beta \ln w_c - (1-\beta) \ln r_c^{M} - (1-\beta) \ln (\tau_c \kappa) + \omega_{jc}$$

Set-up gives rise to conditional logit model, which yields aggregate labor demand N^D_c and aggregate commercial floor space demand M^D_c

Model – Construction Sector **CO**

- Competitive construction sector, producing residential and commercial floor space, $S_c = H_c + M_c$
- ▶ In equilibrium, arbitrage condition between two types of floor space, such that $r_c^M = \phi r_c^H$ with $\phi \ge 1$ capturing stricter regulations, and $H_c = \mu S_c$ (Ahlfeldt et al., 2015)
- Cobb-Douglas constant returns production function with land L_c and capital K_c as inputs (Epple et al., 2010)

$$S_c = H_c + M_c = L_c^{\gamma} K_c^{1-\gamma}$$

▶ Profit maximization yields land demand L_c^D and capital demand K_c

Model – Land Owners (and Capitalists)

- Standard assumption: land owners are absent, i.e., income from land not spent within city (Kline and Moretti, 2014, Ahlfeldt et al., 2015, Diamond, 2016)
- **b** Land supply is given by simple function $\ln L_c = \theta \ln I_c$ with supply elasticity $\theta > 0$ and land price I_c

Capital is traded at the world market so supply is perfectly elastic with fixed price s

Model – Local Governments 🕚

- Municipal government may use property tax revenues to finance local public good G_c
- For now, we assume the PG only affects consumptive amenities A_c
- ▶ For simplicity, assume $r_c^H = r_c^M$, $\kappa = 1$
- **>** Given share ψ of property tax revenues spent on local public good, rest distributed lump-sum
- Budget constraint

$$G_c = \psi (H_c + M_c) r_c^H t_c = \psi S_c r_c^H t_c$$

Model – Example: Effect of property taxes on rents

Zoom in on effect of a property tax increase on rents

$$\frac{d \ln r_c^{H*}}{d \ln \tau_c} = \frac{\frac{d \ln H^D}{d \ln \tau_c}}{\frac{d \ln H^S}{d \ln r_c^H} - \frac{d \ln H^D}{d \ln r_c^H}}$$

- effective housing supply, $\frac{d \ln H^S}{d \ln r_c^H}$, accounts for equilibrium effects of land via output elasticity and supply elasticites of land (γ, θ)
- effective housing demand $\frac{d \ln H^D}{d \ln r_c^H}$ accounts for labor and commercial floor space equilibrium effects via location preferences of workers and firms (σ^W, σ^F), production technology (β), product demand (ρ)
- Standard textbook result: incidence determined by relative supply and demand elasticities

Model – Quantity Effects of the Property Tax Increase



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