Implied Volatility Surface Dynamics: from Parameter-Driven to Observation-Driven Models and Beyond

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³ [Empirical study](#page-36-0)

Xia Zou (VU & TI) **[IV Surface Dynamics](#page-0-0)** August 27, 2024 2 / 20

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A typical approach for the dynamics in factors: state-space framework (parameter driven).

- Challenge: state-space models beyond the linear Gaussian is computationally expensive.
- An alternative: score-driven approach (observation-driven) with an explicit expression for the likelihood function.

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- Koopman et al. [\(2017\)](#page-56-1) in a multivariate framework:
	- Point forecasts: comparable performance.
	- Density forecasts: much worse performance of score-driven models.
- We explore the origins of this difference in a multivariate framework between the two model classes in more detail.
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- \bullet Koopman et al. [\(2017\)](#page-56-1): by assuming the error terms in the score-driven have an equicorrelation structure, the performances become more comparable.
- We pinpoint the origin of performance difference in the state-space and the score-driven framework: a (too) restrictive assumption on the covariance structure of the measurement noise.
- We introduce a simple adaptation of the measurement equation in the score-driven model \rightarrow comparable density forecast performance of score-driven models with their state-space counterparts.
- After closing the performance gap, the score-driven approach can easily be adapted to accommodate non-Gaussian features, without any complication to the ML estimation and inference procedures.
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- We apply our findings to model the dynamics of IV surfaces of S&P500 index options using data from January 2010 to December 2022.
- We find that a linear Gaussian state-space model outperforms a plain-vanilla score-driven model by a large margin, both in terms of density fit, and Value-at-Risk (VaR) violation rates.
- After adapting the score-driven model with the adjusted covariance structure for the measurement errors, the score-driven model behaves roughly at par with the state-space model.
- \bullet Adding Student t error terms to the score-driven model even increases the density fit beyond that of its state-space counterpart.
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Xia Zou (VU & TI) **[IV Surface Dynamics](#page-0-0)** August 27, 2024 7/20

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\boldsymbol{\beta}_{t+1} = (\mathbf{I} - \mathbf{B}) \boldsymbol{\theta} + \mathbf{B} \boldsymbol{\beta}_t + \boldsymbol{\xi}_t
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- \bullet We gather all static parameters of the model into a vector ψ that needs to be estimated.
- This set-up accommodates both a state-space and a score-driven framework, depending on our choice of $\xi_t.$

We assume that IV_t evolves as follows:

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Linear Gaussian state-space : $(\varepsilon_{t}^{\top},\xi_{t}^{\top})^{\top}$ is normally distributed.

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\ell(\psi) = -\frac{1}{2} \sum_{t=1}^{T} \left(\ln |2\pi F_t| + \left(\mathbf{IV}_t - \mathbf{IV}_{t|t-1} \right)^{\top} F_t^{-1} \left(\mathbf{IV}_t - \mathbf{IV}_{t|t-1} \right) \right).
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Score-driven framework: $\boldsymbol{\xi} _t$ is chosen as the derivative (with respect to ${{\beta} _t})$ of the log predictive density of $\boldsymbol{IV_t}.$

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\begin{aligned} \xi_t &= \mathbf{A} \left(\mathbf{M}_t^\top \mathbf{H}_t^{-1} \mathbf{M}_t \right)^{-1} \mathbf{M}_t^\top \mathbf{H}_t^{-1} \varepsilon_t, \\ \ell(\psi) &= -\frac{1}{2} \sum_{t=1}^T \left(\ln|2\pi \mathbf{H}_t| + \varepsilon_t^\top \mathbf{H}_t^{-1} \varepsilon_t \right), \end{aligned}
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Consider a typical specification with diagonal error covariance matrix \boldsymbol{H}_{t} .

Assume the DGP to be a state-space model and fit a score-driven model to it:

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\textit{IV}^{\text{ssf}}_t-\widehat{\textit{IV}}^{\text{sd}}_t = \varepsilon_t+\textit{M}_t\,\xi^{\text{ssf}}_t+\textit{M}_t\,\left((\textbf{I}-\textit{B})\left(\theta - \hat{\theta}\right)+\textit{B}\left(\beta^{\text{ssf}}_{t-1}-\widehat{\beta^{\text{sd}}}_{t-1}\right)\right)
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where β_t^{sd} is the ${\beta_t}$ from the score-driven model, and β_t^{ssf} is that from the state-space.

- The first component has an uncorrelated covariance structure.
- The final term is typically small given the good forecast performance of score-driven models even for a state-space DGP. (see Koopman et al., [2016\)](#page-56-0).
- \bullet It is the second term that results in a cross-correlated prediction errors, which contrasts with the assumed uncorrelated structure.

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 $IV_t = M_t \beta_t + \varepsilon_t,$ $\varepsilon_t \sim h \left(\varepsilon_t \mid H_t + M_t C M_t^{\top}, \vartheta \right),$

- In our single level factor example above, the adjusted measurement equation induces an equicorrelation structure.
- This explains the huge improvements in density fit that Koopman et al. [\(2017\)](#page-56-1) find when imposing equicorrelation structures.
- Our approach above can, however, also be used for richer factor structures.

Adjusted measurement equation

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Xia Zou (VU & TI) **[IV Surface Dynamics](#page-0-0)** August 27, 2024 12/20

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 299

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- Goncalves and Guidolin [\(2006\)](#page-56-4): simple model linear in coefficients and nonlinear in moneyness and time-to-maturity achieves a good fit to the IV surface of S&P 500 index options.
- To illustrate our main findings, we also employ the simple 5 factor models of Goncalves and Guidolin [\(2006\)](#page-56-4) to the IV surface of S&P 500 index options.
- For robustness checks, we also consider a more flexible factor presentation that includes a nonparametric factor.

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Parametric factor specification

$$
IV_t(m,\tau)=M_t\beta_t+\varepsilon_t,
$$

- \bullet IV_t(m, τ) is a q_t-dimensional vector
- $\bm{M}_t=(\bm{m}_1,\bm{m}_2,\dots,\bm{m}_{q_t})$ is a matrix $(q_t\times 5)$ with $\bm{m}_j=(1,m_j,m_j^2,\tau_j,m_j\tau_j)'$
- β_t is vector of latent factors (5×1) ,
- ε_t is the disturbance $(q_t \times 1$),
- q_t is the number of option contracts at time t.
- We use a dataset comprising European call options on the S&P 500 index, encompassing all call and put options traded on the Chicago Board Option Exchange (CBOE).
- We follow van der Wel et al. [\(2016\)](#page-57-1) to filter option contracts.
- \bullet We restrict our analysis to out-of-the-money options, defined by a Δ less than 0.5 in absolute value.
- We exclude observations characterized by time-to-maturity periods exceeding 360 days or shorter than 7 days.
- Options with implied volatilities greater than 0.7 and option prices below 0.05 are omitted from the dataset to mitigate the effect of potential data errors.

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- We follow van der Wel et al. [\(2016\)](#page-57-1) to filter option contracts.
- \bullet We restrict our analysis to out-of-the-money options, defined by a Δ less than 0.5 in absolute value.
- We exclude observations characterized by time-to-maturity periods exceeding 360 days or shorter than 7 days.
- Options with implied volatilities greater than 0.7 and option prices below 0.05 are omitted from the dataset to mitigate the effect of potential data errors.

- We compare the performance of the linear Gaussian state-space and socre-driven models in forecasting out-of-sample IV, both statistically and economically.
- For the score-driven models, we consider four variants: normal and student-t distributions without and with adjusted measurements, of which we denote as $GAS(\mathcal{N}), GAS(\mathcal{D},\mathcal{N}),$ $GAS(t)$, and $GAS(\Sigma, t)$ respectively.
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Table 1: Statistical Accuracy Measures for various estimators of the implied volatility surface: Out-of-sample

Note: Diebold and Mariano test is conducted with the state space model as a benchmark on MSE and MAE, and the null hypothesis is that both models are equally accurate.

Xia Zou (VU & TI) **[IV Surface Dynamics](#page-0-0)** August 27, 2024 17/20

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Difference in the state-space and score-driven models

We compare the VaR forecasts for the equally-weighted average implied volatility of S&P 500 options.

Table 2: Out-of-sample evaluation for 99% Value-at-Risk Estimation

We compare the VaR forecasts for the equally-weighted average implied volatility of S&P 500 options.

	Violation ratio ($\times 10^2$) Tickloss ($\times 10^3$)	
The whole period: 2012-2022		
State Space	0.04	3.68
$GAS(\mathcal{N})$	41.07	9.99***
$GAS(\Sigma, \mathcal{N})$	0.04	$5.16***$
GAS(t)	50.88	14.60***
$GAS(\Sigma,t)$	0.00	$6.14***$

Table 2: Out-of-sample evaluation for 99% Value-at-Risk Estimation

- We compare state-space and score-driven models for option implied volatility surface dynamics.
- We find that point forecasts of both models behave similarly, but density forecasts of the plain-vanilla score-driven model are substantially worse.
- We show how a simple adjustment of the measurement density of the score-driven model can put both models back on an equal footing.
- After this correction, the score-driven model can easily be extended with non-Gaussian features without complicating parameter estimation, unlike its state-space counterpart.

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- Andreou, Panayiotis C, Chris Charalambous, and Spiros H Martzoukos (2010). "Generalized parameter functions for option pricing". In: Journal of Banking $&$ Finance 34.3, pp. 633–646.
- 量 Dumas, Bernard, Jeff Fleming, and Robert E Whaley (1998). "Implied volatility functions: Empirical tests". In: Journal of Finance 53.6, pp. 2059–2106.
- 暈 Goncalves, Silvia and Massimo Guidolin (2006). "Predictable dynamics in the S&P 500 index options implied volatility surface". In: Journal of Business 79.3, pp. 1591-1635.
- Koopman, Siem Jan, André Lucas, and Marcel Scharth (2016). "Predicting time-varying 量 parameters with parameter-driven and observation-driven models". In: Review of Economics and Statistics 98.1, pp. 97–110.
- 量 Koopman, Siem Jan, André Lucas, and Marcin Zamojski (2017). "Dynamic term structure models with score-driven time-varying parameters: estimation and forecasting". In: Narowdy Bank Polski, NBP Working Paper No. 258.
- Pena, Ignacio, Gonzalo Rubio, and Gregorio Serna (1999). "Why do we smile? On the determinants of the implied volatility function". In: Journal of Banking & Finance 23.8, pp. 1151–1179.

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F van der Wel, Michel, Sait R Ozturk, and Dick van Dijk (2016). "Dynamic Factor Models for the Volatility Surface". In: Dynamic Factor Models. Vol. 35. Emerald Group Publishing Limited, pp. 127–174.

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