Cast out the pure? Inflation and relative prices on both sides of the Atlantic

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Aggregate and item inflation variability in HICPs



Headline HICP inflation and HICP item level inflation rates, rescaled year on year monthly rates.



Pure inflation vs relative price changes

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Policy relevance:

- ▶ How do sectoral and aggregate shocks affect pure inflation and relative prices?
- Should monetary policy respond to sectoral shocks?

Relation to existing literature

Aggregate inflation and relative prices:

- Old debate on relationship between aggregate inflation and higher moments of cross-sectional distribution: Ball and Mankiw 1995, Balke and Wynne 2000, Bryan and Cecchetti 1999
- Explicitly disentangling "pure" inflation: Reis and Watson 2010, recently updated for USA by Ahn and Luciani 2021
- Estimating the common trend or cyclical ("core") component of inflation: Cristadoro et al. 2005, Mertens 2016, Stock and Watson 2016 (using dynamic factor models)
- Diagnosing the sources of the post-2020 inflation surge: Ruge-Murcia and Wolman 2022, Luo and Villar 2023, diGiovanni et al. 2023

An illustrative example of pure inflation

To fix ideas, consider a model with

- i multiple sectors, differing by price stickiness and productivity processes
- ii money growth rule and aggregate technological growth

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- i multiple sectors, differing by price stickiness and productivity processes
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Then persistent shocks to (i) will move relative prices, persistent shocks to (ii) will move all *updating* prices in the same way:

absent other shocks, (ii) will progressively affect all prices in the same way

Data and Method

Data

- Harmonised, CPI-basket item level data 4-digit COICOP classes
- ▶ About 120 countries, average coverage of 55 classes each
- Monthly frequency, over 2000-2023

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Method

- Starting point: extract common components from large panel of time series
- > Put restrictions on a Bayesian DFM to identify a "pure inflation" component
- > Project inflation on this component, common relative prices and an idiosyncratic part

Why this approach

Building on Reis and Watson 2010:

- Estimate a dynamic factor model on many inflation sub-categories for the US end of quarter on end of quarter
- Ad hoc factor: cross-sectional inflation
- > Project factors into *pure inflation*, which loads uniformly, and *aggregate relative prices*.

Too many sample-specific specification choices to apply to 120 countries (our end goal).

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A more flexible approach:

- Monthly frequency and year-on-year inflation to account for heterogeneous price rigidity
- Loading restriction is a prior, so data can reject it

A Bayesian DFM with a "uniformity" restriction

Barebone state space representation of a dynamic factor model:

 $\pi_t = \Lambda F_t + \varepsilon_t$ $F_t = \Gamma F_{t-1} + \eta_t$

where F_t is 5×1 and $\Lambda F_t = \mathbf{1}a_t + \gamma R_t$

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- **•** Top $K \times K$ part of Λ is lower triangular and has strictly positive diagonal
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From 5 factors to pure (ν_t) and relative prices inflation (ρ_t)

$$\begin{array}{l} \nu_t = \mathsf{a}_t - \mathsf{E}\left(\mathsf{a}_t|\mathsf{F}_{t-1}\right)\\ \rho_t = \mathsf{E}\left(\mathsf{F}_t|\mathsf{R}_{t-1}\right)\\ \text{so that} \quad \pi_t = \nu_t + \beta\rho_t + u_t \end{array}$$



Estimation details

 $\pi_t = \Lambda F_t + \varepsilon_t$ $F_t = \Gamma F_{t-1} + \eta_t$

Priors

 $\pi_{t} \sim N(\Lambda F_{t}, Q_{\epsilon})$ $F_{t} \sim N(\Gamma F_{t-1}, \mathbb{1}_{K})$ $\Lambda = \begin{bmatrix} \Lambda_{K} \\ \Lambda_{N-K} \end{bmatrix}$ $\Lambda_{K} \sim \text{Chol. fact.}$ $\Gamma \sim N(0, 10)$

Setup

NUTS sampler with mostly Gaussian priors

5000 total iterations, 50% burn-in inclusive of adaptation phase

Number of common factors



Cumulative share of variance explained by principal components.

Uniform loading: Clustered PCA loadings



Estimated loadings from PCA. Row-wise K-means clustering with 5 clusters.

Result 1: Separating pure inflation from relative price changes

$$\pi_t = \nu_t + \beta \rho_t + u_t$$



Estimated ν_t (blue), ρ_t (red), and actual HICP (black)

Results 2: response to sectoral shocks: Global oil supply shocks

Figure: Cumulative impulse response function of ν_t and ρ_t to an oil supply shock



CIRFs after a **positive oil supply shock**. Shocks from Baumeister and Hamilton 2019, until December 2023. Shades mark 90% confidence intervals. Estimates from local projections.

Results 2: response to aggregate shocks: traded/non-traded goods

Figure: Cumulative impulse response of ν_t and ρ_t to a global demand shock



CIRFs after a **positive global demand shock** (economic activity shock from Baumeister and Hamilton 2019), until December 2023. Shades mark 90% confidence intervals. Estimates from local projections.



Results 2: response to aggregate shocks: ECB monetary policy

Figure: Cumulative impulse response function of ν_t and ρ_t to an ECB mon. pol. shock



CIRFs after a **"poor man's" MP shock** from Jarociński and Karadi 2020, updated until June 2023. Shades mark 90% confidence intervals. Estimates from local projections.



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- What drives headline inflation in the euro area and the United States?
 - \blacktriangleright Relative prices are the bulk of inflation over the past \sim 25 years
 - Even more so since 2021

The plan ahead

Current working areas:

- Documenting goodness of fit of alternative specifications
- Robustness to different numbers of factors and lag structures
- Incorporating model-based restrictions on IRF of pure and relative price inflation into estimation - how?
- US case study covering the '80s inflation
- Model with permanent shifts of expectations and money growth/monetary regime

Still work in progress, comments and suggestion are welcome!

Thank you!

A less granular view: HICP and its main components



Headline HICP inflation and sectoral inflation rates, year on year monthly rates. Energy rescaled on the right axis.

Long run restriction

Intuition: the long-run effect of MP shocks is zero on pure inflation.

- ► Add an appropriate observed series, *S*, in the state equation $F_t = \Gamma F_{t-1} + \eta_t$
- for two factors, the first capturing pure inflation

$$\begin{bmatrix} S_t \\ F_{1,t} \\ F_{2,t} \end{bmatrix} = \Gamma \begin{bmatrix} S_{t-1} \\ F_{1,t-1} \\ F_{2,t-1} \end{bmatrix} + \begin{bmatrix} \eta_{S,t} \\ \eta_{F_1,t} \\ \eta_{F_2,t} \end{bmatrix}$$

► Then in the long run $F_{t,\infty} = (I - \Gamma)^{-1} \eta_t$ with the following restrictions

$$Q\eta_{t} = (I - \Gamma)^{-1} \eta_{t} = \begin{bmatrix} * & 0 & 0 \\ 0 & * & 0 \\ * & * & * \end{bmatrix} \begin{bmatrix} \eta_{S,t} \\ \eta_{F_{1},t} \\ \eta_{F_{2},t} \end{bmatrix}$$

• How to sample Q and ensure Γ makes the VAR stable?

Model-based intuition

Consider a slightly modified version of Ghassibe 2021, which has

- Sectorally heterogeneous price rigidity
- Input/output intermediates supply network
- Consumption shares
- ► *M* = *PC*

Set of transitory/permanent shocks

Money, aggregate TFP



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The model helps to

- Understand how not only sectoral, but also aggregate shocks show up in relative prices
- Test our empirical approach

Temporary sectoral TFP shock



IRFs of a 1% SD sectoral TFP shock on the flexible and upstream sector. The price gap the log-difference between the sticky and flexible price.

Temporary aggregate shocks – Technology



IRFs of a 1SD TFP shock. "Sticky Price" is a downstream sector whose prices adjust infrequently; "Flexible Price" is an upstream sector with very frequent price adjustment. The price gap is the log-difference between the sticky and flexible price.

Temporary aggregate shocks – Monetary policy shock



IRFs of a 1% money stock increase from steady state. The price gap is the log-difference between the sticky and flexible price.

Persistent monetary shock



IRFs of a 1% persistent money stock increase from steady state. The price gap is the log-difference between the sticky and flexible price.

A look at the first 5 common factors using PCA and DFM



Principal Components and Dynamic Factors estimates.

Clustered loadings – DFM



Estimated loadings from DFM, using Doz, Giannone, and Reichlin 2012 approach. Row-wise K-means clustering with 5 clusters.

Estimated Factors



(a) EA

(b) US

Estimated factors from the BDFM. First factor's loading restricted to be close to one.

Estimated Factors



Estimated factors from the BDFM on 2 years inflation rates. First factor's loading restricted to be close to one.

Estimated loadings – BDFM



Loadings heatmap: median values of the posterior distributions.

Oil price inflation and estimated relative price components



Oil price and estimated relative price components

Result 1: Robustness to a longer time lag

$$\pi_t = \nu_t + \beta \rho_t + u_t$$



Estimated ν_t (blue), ρ_t (red), and actual HICP (black)

Check on artificial data from the theoretical model



Estimates of ν (blue, pure inflation) and ρ (red, relative prices) on simulated data from Ghassibe 2021. Inflation data are centered and scaled.



Pre-Covid robustness: Global oil supply shocks

Figure: Cumulative impulse response function of ν_t and ρ_t to an oil supply shock



CIRFs after a **positive oil supply shock**. Shocks from Baumeister and Hamilton 2019, until December 2019. Shades mark 90% confidence intervals. Estimates from local projections.

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Results 2: shocks spillover: crude US fiscal policy

Figure: Cumulative impulse response function of ν_t and ρ_t to a US fiscal policy shock



CIRFs after a "crude" US fiscal shock, computed as percentage deviation of local and federal spending from SPF estimates. Shades mark 90% confidence intervals. Estimates from local projections.

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Robustness: Oil supply news shocks

Figure: Cumulative impulse response function of ν_t and ρ_t to an oil supply news shock



CIRFs after a **restrictive oil supply news shock**. Shocks from Kanzig 2021, until June 2023. Shades mark 90% confidence intervals. Estimates from local projections.

Robustness: actual oil supply shock

Figure: Cumulative impulse response function of ν_t and ρ_t to an oil supply shock



CIRFs after a **restrictive oil supply shock**. Shocks from Kanzig 2021, until June 2023. Shades mark 90% confidence intervals. Estimates from local projections.

Check on artificial data from the theoretical model



Simulated data from Ghassibe 2021: 161 sectors with input/output intermediate goods network, three degrees of price rigidity, differing final consumption shares. 300 periods in total at monthly frequency. Black solid line is aggregate HICP, grey shaded lines are sector-level inflation rates.



Computing ν and ρ

With the estimated factors f_k in F, the matrix Γ with elements γ_{ij} , and observed inflation π

$$\nu_{t} = f_{1,t} - \sum_{k=1}^{K} f_{k,t-1} \gamma_{1,k}$$

$$\rho_t = F_{-1}\hat{\beta} + \sum_{k=1}^{K} f_{k,t-1}\gamma_{1,k}$$

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$$F_{-1} = (f_2' \quad f_3' \quad f_4' \quad f_5')'$$

and

$$\hat{\beta} = \left(F_{-1}'F_{-1}\right)^{-1}F_{-1}'(\pi - \nu)$$