

A distributional theory of household sentiment[‡]

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Abstract

A substantial body of psychological literature suggests that decision-makers' forecasts about their future situations tend to be excessively swayed by new information. We posit that households' beliefs about their future income are distorted by recent income news, positive shocks inducing overoptimism. We validate this idea empirically using the Italian Survey of Household Income and Wealth. We examine the theoretical implications of this behavioral bias on households' consumption-saving decisions by embedding the theory of diagnostic expectations into a heterogeneous agents incomplete market model. The interaction between diagnostic expectations and incomplete markets generates a poverty trap by making it harder for households to escape the borrowing limit: a positive income shock leads to overoptimism about the future and hence over-consumption, making it harder for constrained households to accumulate assets. This can explain the stickiness of the hand-to-mouth status observed in the data: it is significantly less likely to escape this state under diagnostic expectations than under rationality. In addition, this simple behavioral deviation allows us to match the (untargeted) share of hand-to-mouth households. Finally, we find that the welfare costs of diagnostic expectations are heterogeneous across households, representing on average 3.3% of lifetime consumption.

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1 Introduction

A substantial number of households consistently face the challenge of escaping financial constraints. Prevalent one-asset heterogeneous agent models tend to underrepresent this reality as they produce an insufficient number of financially constrained households. Several extensions have been proposed to address this shortcoming, as documented by [Kaplan and Violante \(2022\)](#). However behavioral explanations remain scant. This is surprising given that [Keynes \(1936\)](#) already suggested that fluctuations in households' irrational optimism may shape saving decisions. Could sudden outbursts of sentiment indeed precipitate poor financial planning, subsequently ensnaring households in financial insecurity? We put forth that a simple behavioral friction, specifically that agents over-extrapolate recent income news when predicting their future income path, can result in persistently constrained households when markets are incomplete. Constrained agents, when receiving positive income shocks, become excessively optimistic and overconsume, thereby sustaining their proximity to the borrowing limit.

To develop this idea, we make four contributions. The first contribution is methodological. We adapt a well documented theory of overreaction to news, diagnostic expectations, to idiosyncratic shocks in an incomplete market setting. To do so, we develop a new methodology to handle general deviations from rational expectations in heterogeneous agents models: the rationality wedge. The rationality wedge is akin to a transition matrix capturing all of the agent's misperceptions concerning the evolution of the states. This new tool allows us to generalize the intuition of diagnostic expectations to a wide range of stochastic processes and depart from the traditional AR(1) restriction previously used in the literature.

The second contribution is theoretical. Building on [Maxted \(forthcoming\)](#), we posit that households have 'sentiment', a bias that distorts their perception about the future evolution of their income. Positive sentiment makes agent overoptimistic and vice-versa. As our model features jumps in productivity rather than a standard diffusion, income shocks act as psychological 'traumas' to agents and make sentiment jump instantaneously. Without income news sentiment mean-reverts exponentially as the 'traumatic' memory of the shock decays. Our theory provides very intuitive predictions : positive sentiment inflates current consumption as agents believe they will be richer from a life-time prospective. Thus positive income shocks lead to consumption overreaction in the short run, and underreaction in the long-run as agents deplete their assets in the sentimental phase. Sentiment is distributed and the joint distribution of wealth, income and sentiment defines the state of the economy.

The third contribution is empirical. Existing literature has presented robust empirical evidence supporting diagnostic expectations. However, the focus is typically applied to aggregate variables and agents such as firms or investors. In contrast, we consider *households'* expect-

tations concerning *idiosyncratic* variables. To substantiate the use of diagnostic expectations in this context, we offer new empirical evidence pertaining to agents' extrapolative expectations in relation to income shocks. Using data from the Survey of Household Income and Wealth (SHIW), a comprehensive survey of Italian households executed by the Bank of Italy, we show that households having experienced income growth are inclined to forecast future income above their ex-post realized value. Additionally, we confirm a pattern previously documented in the Michigan survey that high-income households display a consistent tendency towards over-optimism, and vice-versa for low-income households. We tie these two empirical observations to our sentiment variable and its co-movement with income and income growth, and use these data to calibrate our sentiment parameters accordingly.

Lastly, our data enable three quantitative contributions. First, our calibrated model reveals that diagnostic expectations amplifies the prevalence of hand-to-mouth households, a finding aligning with our data. We underscore that such households do not necessarily belong to the low-income category. Rather, they often represent middle-income, high-sentiment households – a group we term as the 'hopeful hand-to-mouth'. This insight helps to elucidate the significant proportion of above-average income hand-to-mouth households observed in our data. Diagnostic expectations amplify the consumption response to income shocks and elevate the average marginal propensity to consume in our economy. The fact that sentiment is dispersed in our economy could serve as an additional explanatory factor for the recently debated latent heterogeneity in consumption response to shocks. Secondly, diagnostic expectations makes it harder to escape financial constraints. A positive income shock increases sentiment, which correspondingly depresses agents' savings in comparison to a rational framework. When the agent is not financially constrained, there is a offsetting effect : a negative income shock induces pessimism, thereby enhancing the household's savings relative to the rational benchmark. However, these two effects are asymmetric at the borrowing limit. In the rational and diagnostic models alike, negative shocks have no impact on savings which remain zero at the borrowing limit – hence sentiment does not distort behavior when agents are constrained and shocks are negative. This asymmetry intensifies the stickiness of the hand-to-mouth state relative to the rational benchmark : the likelihood of an agent escaping the hand-to-mouth state is consistently and significantly lower for the diagnostic agent compared to the rational one. Finally, we demonstrate that this behavioral friction can represent a substantial cost, on average 3.3 percent of consumption. This underscores the argument that income fluctuations are costly, as increased income volatility leads to more significant intertemporal errors.

Related Literature First, we build upon the recent literature embedding behavioral frictions to incomplete market models.¹ A growing strand of literature has been incorporating present bias of the [Harris and Laibson \(2013\)](#) type to heterogeneous agents models, see [Laibson and Maxted \(2023\)](#) and [Maxted \(2023\)](#). [Laibson et al. \(2021\)](#) show that present bias amplifies the household balance-sheet channels of macroeconomic policy when agents can conduct cash-out refinancing. We differ from this this strand by introducing behavioral frictions on beliefs rather than preferences, in this sense we are closer to [Rozsypal and Schlafmann \(Forthcoming\)](#) who introduce over-persistence bias to an Ayiagari-Bewley-Hugett economy. We show that introducing sentiment can potentially explain the latent heterogeneity in consumption response to shocks discussed in [Lewis et al. \(2019\)](#) and [Arellano et al. \(2023\)](#).

We extend the literature on diagnostic expectations pioneered by ([Bordalo et al., 2018](#)). We rely and extend the continuous-time representation of diagnostic expectations proposed in [Maxted \(forthcoming\)](#). Unlike this paper, we apply diagnostic expectations to idiosyncratic shocks to income, instead of aggregate shocks, and develop a methodology to apply them to jump-drift processes instead of Brownian motions. We provide novel evidence and quantitative estimates on the degree of diagnosticity in households' expectations about their own future income, complementing the results in [Gennaioli et al. \(2016\)](#), [Bordalo et al. \(2020\)](#), and [Ma et al. \(2020\)](#). Our approach will differ from that of [Bianchi et al. \(2021\)](#) and [L'Huillier et al. \(2021\)](#), who integrate diagnostic expectations into standard real business cycle and new Keynesian models respectively, but assume a representative agent. We focus instead on an incomplete market framework.

Our research also contributes to the rich body of theoretical work on psychology-driven poverty traps. The studies by [Banerjee and Mullainathan \(2010\)](#) and [Bernheim et al. \(2015\)](#) underscore the influence of present bias and temptation in driving poverty. Studies that relate to our work include that by [Thakral and Tô \(2021\)](#), which posits an alternate psychology-oriented theory of poverty traps, proposing that consumers might be disinclined to save, knowing their future selves could squander savings, or [Sergeyev et al. \(2023\)](#), in which poverty traps result from financial stress in the vicinity of the borrowing limit. Similarly, the study by [Dalton et al. \(2016\)](#) adds a valuable dimension to this discussion by examining the impact of reference dependence and aspirations on poverty traps, through a different psychological lens. We contribute to this strand of literature by showing that a well accepted form of deviation from rational expectations can generate poverty traps.

¹See also [Pappa et al. \(2023\)](#), who study the interaction between incomplete markets and expectational shocks in a macroeconomic model with labor market frictions.

Outline Section section 2 lays out our model of diagnostic expectations in an incomplete market setting. Section section 3 characterizes the effect of sentiment on household consumption-saving behavior. Section section 4 provides suggestive evidence of diagnostic expectations with respect to income and discusses calibration. Section section 5 present steady state implications of household’s sentiment. Section section 2 concludes.

2 A partial equilibrium model of sentiment

This section introduces a simple incomplete market model where agents feature diagnostic expectations for their idiosyncratic income. We then derive analytically and numerically the main implications for household’s consumption-saving behavior in partial equilibrium.

2.1 Model set up

Earnings dynamics Time t is continuous. Households are endowed with an idiosyncratic flow of productivity e^{y_t} that they supply to firms inelastically against a wage w , earning a total labor income we^{y_t} . As is standard in the quantitative heterogeneous agents literature, individual productivity follows a jump-drift process in logs. Jumps arrive at a Poisson rate λ . Conditional on a jump, a new log-productivity state y' is drawn from a Normal distribution with mean zero and variance σ^2 , i.e., $y' \sim \mathcal{N}(0, \sigma^2)$. Between jumps, the process mean-reverts exponentially at rate μy_t . Formally, the process for y_t can be represented as²

$$dy_t = -\mu y_t dt + dN_t \tag{1}$$

where the shock dN_t captures the income change from y to y' . The reason why we don’t rely on a more simple two-state Markov chain is that diagnostic expectations, described momentarily, is more amenable to a setting in which the support for productivity is continuous.

Diagnostic expectations and Sentiment We extend the definition of sentiment presented in [Maxted \(forthcoming\)](#) to the jump-drift process in (1). Households’ perception of the drift of log productivity is biased by a “sentiment” variable S_t which captures recent income shocks. Intuitively, household who received a negative sequence of income shocks will feature negative sentiment and will thus underestimate their future income (and vice-versa). Formally,

²More formally, the infinitesimal generator of this process is given by $\mathcal{I}v(y) = -\mu y \partial_y v(y) + \lambda \int v(y') - v(y) d\Phi(y'/\sigma)$, where $\Phi(\cdot)$ is the CDF of the normal distribution.

households' perceived log-productivity process is given by

$$\widetilde{dy}_t = \left(-\mu y_t + \theta \mathcal{S}_t \right) dt + dN_t, \quad \mathcal{S}_t \equiv \int_{-\infty}^t e^{-\kappa(t-s)} dN_s. \quad (2)$$

where \widetilde{dy}_t denotes the *perceived* law of motion for y_t , as opposed to the real one. Sentiment is defined as a weighted sum of past log-productivity shocks, with an exponential discounting parameter κ .³ In this sense, κ can be understood as a “traumatic memory” parameter which captures the extent to which past shocks persistently affect households' psychology. The parameter θ captures the degree of diagnosticity, with $\theta = 0$ nesting rational expectations.

A convenient property of this formulation is that the law of motion for sentiment can be expressed recursively and hence we can get the joint law of motion for log productivity and sentiment:

$$dy_t = -\mu y_t dt + dN_t, \quad d\mathcal{S}_t = -\kappa \mathcal{S}_t dt + dN_t.$$

From the equations above, it is clear that the processes for income and sentiment are correlated as they are hit with the same shocks dN_t . In particular, the infinitesimal generator for the joint process (y_t, \mathcal{S}_t) is given by

$$\mathcal{B}v(y, \mathcal{S}) = -\mu y \partial_y v(y, \mathcal{S}) - \kappa \mathcal{S} \partial_{\mathcal{S}} v(y, \mathcal{S}) + \lambda \int v(y', y' - y + \mathcal{S}) - v(y, \mathcal{S}) d\Phi(y'/\sigma).$$

Where the argument $y' - y + \mathcal{S}$ captures the fact that \mathcal{S} is shocked by the change in income $y' - y$.

Household's problem To study the implications of sentiment on the consumption-saving behavior of households, we embed the joint process for log productivity and sentiment into an otherwise standard incomplete markets framework. Throughout the rest of the paper, we focus on partial equilibrium behavior. Households solve the income fluctuation problem by choosing how much to consume and save in a single risk-free asset a yielding a return of r . Agents face a standard borrowing limit \underline{a} . The household's problem is:

$$\max_{\{c_t\}_{t \geq 0}} \widetilde{\mathbb{E}}_0 \int_0^{\infty} e^{-\rho t} u(c_t) dt \quad \text{s.t.} \quad \dot{a}_t = r a_t + w e^{y_t} - c_t, \quad a \geq \underline{a}$$

³Since dN_s is a jump process, the integral boils down to a simple sum of shocks. In particular let $\{\dots, t_2, t_1\}$ be the timing of past realized shocks and $\{\dots, dN_2, dN_1\}$ the corresponding income changes. Sentiment then simply reads $\mathcal{S}_t = \sum_{s=1}^{\infty} e^{-\kappa t_s} dN_s$.

where the only deviation from the standard formulation is in the expectation operator $\tilde{\mathbb{E}}_0$, which captures households' deviations from rational expectations.

2.2 Recursive representation: the rationality wedge

We now present a general recursive representation of non-rational expectations, captured by a linear operator which we call the “rationality wedge”. This wedge is able to handle a wide range of deviations from rational expectations in heterogeneous agents models with an arbitrary number of states in a tractable way. In particular, by leveraging the rationality wedge, we can extend diagnostic expectations to general classes of stochastic processes, thus going beyond the AR(1) case which the literature has typically focused on.

Hamilton-Jacobi-Bellman (HJB) equation In addition to assets a and log-productivity y , we include sentiment \mathcal{S} as a third state variable. We then derive a representation of deviations from rational expectations as a wedge in the agent's HJB equation. We refer to this wedge as the “rationality wedge” and denote it by $\Psi(\mathcal{S})$. This formalization of the deviations makes computations more transparent and can be generalized to a wide range of deviations from rational expectations.

As is standard in the behavioral economics literature, we can distinguish between two types of agents: sophisticates and naïve. Sophisticates understand that they have sentiment and internalize its true law of motion. Naïve agents, on the other hand, are not aware that they misperceive the true law of motion for their income process, and as a result also ignore the law of motion for sentiment. The value function of the sophisticated household solves the following HJB equation:⁴

$$\rho V(a, y, \mathcal{S}) = \max_c u(c) + \underbrace{(ra + we^y - c)\partial_a V(a, y, \mathcal{S})}_{\text{propagation in } a} + \underbrace{BV(a, y, \mathcal{S})}_{\text{propagation in } (y, \mathcal{S})} + \underbrace{\theta \mathcal{S} \partial_y V(a, y, \mathcal{S})}_{\text{rationality wedge}} \quad (3)$$

The rationality wedge $\Psi(\mathcal{S})V(a, y, \mathcal{S}) = \theta \mathcal{S} \partial_y V(a, y, \mathcal{S})$ captures the fact that households misperceive the drift of their income by a factor $\theta \mathcal{S}$. Apart from this misperception of the income drift, and because of sophistication, they understand the true joint law of motion of income and sentiment, captured by \mathcal{B} .

Naïve agents do not internalize the movement of \mathcal{S} when solving their consumption-saving

⁴The HJB (4) is subject to the standard boundary condition $\partial_a V(\underline{a}, y, \mathcal{S}) \geq u'(\underline{a} + we^y)$.

problem. Their HJB equation is then given by:

$$\begin{aligned} \rho V(a, y, \mathcal{S}) = \max_c & u(c) + \underbrace{(ra + we^y - c)\partial_a V(a, y, \mathcal{S})}_{\text{propagation in } a} + \underbrace{\mathcal{B}V(a, y, \mathcal{S})}_{\text{propagation in } (y, \mathcal{S})} \\ & + \underbrace{\mathcal{I}V(a, y, \mathcal{S}) - \mathcal{B}V(a, y, \mathcal{S}) + \theta \mathcal{S} \partial_y V(a, y, \mathcal{S})}_{\text{rationality wedge}} \end{aligned} \quad (4)$$

The rationality wedge of the naïve agent additionally captures the fact that they don't perceive the true joint motion of y and \mathcal{S} captured by \mathcal{B} , which has to be removed, and instead only perceive the evolution of y , captured by its generator \mathcal{I} .

In this paper, we will focus the analysis on the naïve case. There are two reasons for this. First, naïveté is more intuitive: agents simply misunderstand the evolution of their income. Second, the naïve policy functions have the nice feature that when sentiment is zero they exactly boil down to the rational expectations case. This is instead not true for the sophisticated agent who realizes that they have sentiment, which distorts their behavior even when sentiment is zero. We performed all the analysis under sophistication and the results are very similar in our calibration.

Distributional dynamics Just like income and wealth are distributed in standard incomplete market models, sentiment is unequally distributed between agents and the state of the economy will be captured by a joint measure $G(da, dy, d\mathcal{S})$ over these three states. To simplify notation, let $x \equiv (a, y, \mathcal{S})$ denote the vector of state variables indexing a household. Now denoting by $g(x)$ the joint density over the three states, we can represent the economy as a stationary mean-field-game in the following way:

$$\rho V(x) = \max_c u(c) + (ra + we^y - c)\partial_a V(x) + \mathcal{B}V(x) + \underbrace{\Psi(\mathcal{S})V(x)}_{\text{rationality wedge}} \quad (5)$$

$$0 = -\partial_a[s(x)g(x)] + \mathcal{B}^* g(x) \quad (6)$$

where $s(x) \equiv ra + we^y - c(x)$ denotes optimal savings solving equation (5) and \mathcal{B}^* is the adjoint of the operator \mathcal{B} .⁵ This system differs from standard rational expectations mean-field-games insofar as the differential operator in the HJB equation is usually the infinitesimal generator of the true stochastic process and hence the differential operator in the Kolmogorov Forward equation is the adjoint of the differential operator in the HJB equation. The rationality wedge violates this duality. Note also that deviations from rational expectations can impact the true

⁵The adjoint \mathcal{B}^* is the continuous-state continuous-time equivalent of a discrete-state transition matrix, hence the adjoint of an operator is akin to the transpose of a matrix, see [Achdou et al. \(2022\)](#).

state dynamics only via the optimal decisions made by households $s^*(x)$.

Belief distortions with sentiment shocks The rationality wedge can also be used to compute the *perceived* statistical evolution of the agent's states, which allows us to understand the way individuals' belief distortions evolve over different forecast horizons. Suppose the agent starts with some states x_0 . Their perceived evolution of future states, which we capture with the evolution of the perceived density of $x_{t+\tau}$ denoted by $\tilde{f}(\cdot|x_0)$, is obtained from a Kolmogorov Forward equation featuring the adjoint of the rationality wedge :

$$\partial_t \tilde{f}_t(x|x_0) = -\partial_a[s^*(x)\tilde{f}_t(x|x_0)] + \mathcal{B}^* \tilde{f}_t(x|x_0) + \Psi(\mathcal{S})^* \tilde{f}_t(x|x_0) \quad (7)$$

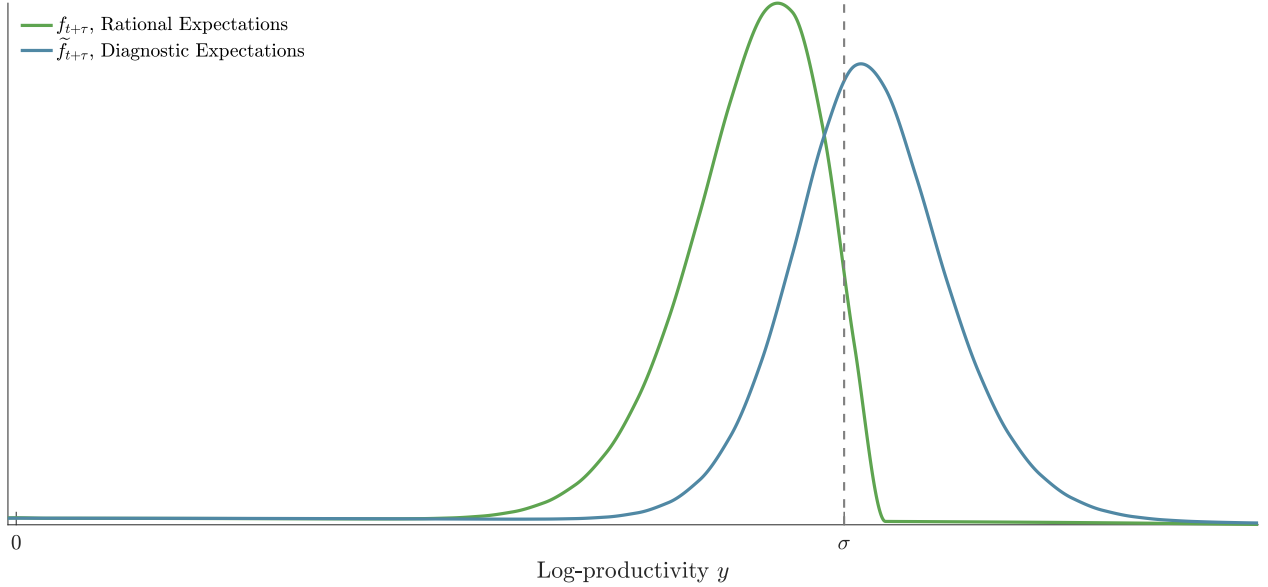
subject to the boundary condition $\tilde{f}_0(x|x_0) = \delta\{x_0\}$, where $\delta\{\cdot\}$ is the Dirac delta function. We can use equation (7) in two ways. First, for any initial state x_t , we can compute the true statistical evolution of $x_{t+\tau}$, denoted by $f(x_{t+\tau}|x_t)$, by solving the equation forward when $\Psi^*(\mathcal{S}) = 0$, which corresponds to the rational expectations case. Second, we can compute household's expected income in period $t + \tau$ under diagnostic expectations $\tilde{\mathbb{E}}_t(y_{t+\tau})$ given its states at the forecast date t using

$$\tilde{\mathbb{E}}_t(y_{t+\tau}|x_t) = \int y \tilde{f}_{t+\tau}(x|x_t) dx. \quad (8)$$

Using equation (7), we can visualise the way diagnostic agents' expectations depart from rational agents'. Figure 1 shows how diagnostic expectations distorts the agent's beliefs following an income shock. We consider a one standard deviation positive shock to log-productivity, starting with the median log-productivity ($y = 0$) and no sentiment. Instantaneously after the news, the agent forecasts the future evolution of its log-productivity for each future period $t + \tau$. The green density corresponds to the rational (and therefore accurate) forecast for log-productivity two years after the shock.⁶ The green line shows how diagnostic expectations distorts beliefs. Like in standard diagnostic expectations, the agent over-weights the probability of states that were made more likely after the shock, and in particular here perceives a probability distribution that is over-optimistic. While the diagnostic expectations framework has so far typically been applied to AR(1) processes for tractability, our rationality wedge approach allows to generalize the intuition of diagnostic expectations to very general stochastic processes such as the jump drift process we're using.

⁶We use two years in this example to align with our data, described in Section 3.

Figure 1: Belief distortion under diagnostic expectations



Note: This plot shows the subjective expected distribution of log-productivity $y_{t+\tau}$, where τ corresponds to two years. $\tilde{f}_{t+\tau}$ is based on diagnostic expectations (blue line) and $f_{t+\tau}$ is based on rational expectations (green line). The beliefs are formed right after a positive shock of magnitude σ at t , with initial conditions $y = S = 0$.

3 The effect of household sentiment on consumption behaviour

We can now describe the effect of sentiment on households' consumption-saving behaviour. We first show how sentiment, as a new state variable, distorts the agent's policy functions, and then focus on the dynamics of sentiment to study the consumption response to income shocks.

3.1 Sentiment as a state

In this section we characterize how sentiment distorts the consumption-saving behavior of households in partial equilibrium. When agents are naïve, all the results in this section are independent from the law of motion of sentiment itself.

Distortion on the Euler equation Proposition 1 provides a tractable formulation of households' Euler equation in the presence of sentiment:

Proposition 1. *Consumption obeys the following Euler equation*

$$\mathbb{E}_t \frac{du'(c(x_t))/dt}{u'(c(x_t))} = \left[\rho + \frac{\theta \mathcal{S}_t \times \eta(x_t)}{IES(x_t)} \right] - r \quad (9)$$

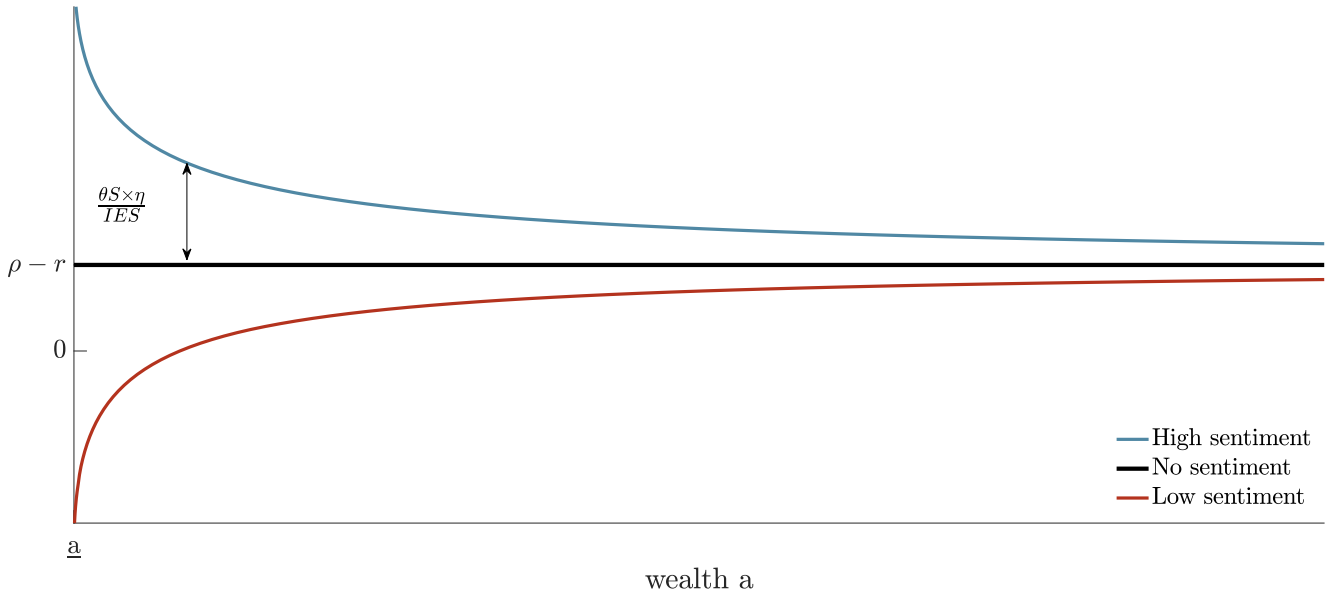
where \mathbb{E}_t is the rational expectations operator over a and y , $IES(x_t) \equiv -u'(c(x_t))/(c(x_t)u''(c(x_t)))$

and $\eta(x_t)$ is the income elasticity of consumption $\eta(x) \equiv \partial \log c(x) / \partial y$.

Proof. Appendix A.1 □

The left hand side of equation (9) is the expected growth rate in marginal utility holding sentiment fixed.⁷ The right hand side has the standard $\rho - r$ term, but with the perceived rate of return on savings being distorted by sentiment. When $\theta = 0$ we nest rational expectations for any value of S . Similarly, independently of the value of θ , when sentiment $S = 0$, the distortion vanishes and the agent behaves as if they were rational. When sentiment is positive, however, the perceived utility return on wealth is depressed and the saving motive is dampened. As a result, positive sentiment leads to more consumption, and vice-versa. This is intuitive: when agents expect their income to rise from a lifetime prospective, they consume more. This mechanism is captured by the other terms in the wedge. The extent to which sentiment distorts the Euler equation depends on the income elasticity of consumption η . When sentiment is equal to S , agents wrongly expect their labor income to go up by $\theta S dt$ percent over dt units of time. This increase in income should lead to an increase in future consumption by $\theta S_t \times \eta(x_{t+dt}) dt$ percent in the next period, which depresses the marginal utility of consumption in the future, thus reducing the saving motive today depending on the inter-temporal elasticity of substitution.

Figure 2: Euler equation distortions along the wealth distribution



Note: The calibration used is our benchmark calibration. The graph is plotted fixing a value of log-productivity y for illustration.

⁷ S is held constant in this Euler equation because the naïve agent does not realize that sentiment moves over time. The Euler equation of the sophisticated agent looks identical to (9), with the exception that the expectation operator is now taken over S as well.

An important implication of (9) is that the right hand side of the Euler equation is state dependent when markets are incomplete, as η depends on the household's states. In particular, in our framework, the distortions are decreasing in wealth, since the income elasticity of consumption is higher when households are closer to the borrowing limit. Figure 2 illustrates these distortions for a given level of productivity y . The black line corresponds to the standard right hand side of the rational Euler equation, $\rho - r$. When sentiment is positive (negative), the right hand side of the Euler Equation is inflated (depressed), and heterogeneously so depending on wealth as the blue and red lines show in Figure 2. This intuition also leads to the following lemma.

Lemma 1. *Suppose that log-productivity is constrained to be in $[\underline{y}, \bar{y}]$ and sentiment is constrained to be in $[\underline{\mathcal{S}}, \bar{\mathcal{S}}]$. Then when $r < \rho$ and with CRRA utility sentiment and income do not matter at the top of the wealth distribution and in particular as $a \rightarrow \infty$*

$$s(a, y, \mathcal{S}) = \frac{r - \rho}{\gamma} a \quad (10)$$

where γ is the inverse IES.

Intuitively, at the top of the wealth distribution, not only is the income elasticity of consumption η low, but also labor income plays a minor role in total income. As a result, agents' misperceptions about labor income are just as irrelevant as its fluctuation.⁸ This is a special case of Proposition 2 in Achdou et al. (2022). This observation motivates us to focus the rest of the analysis on agents close to the borrowing limit.

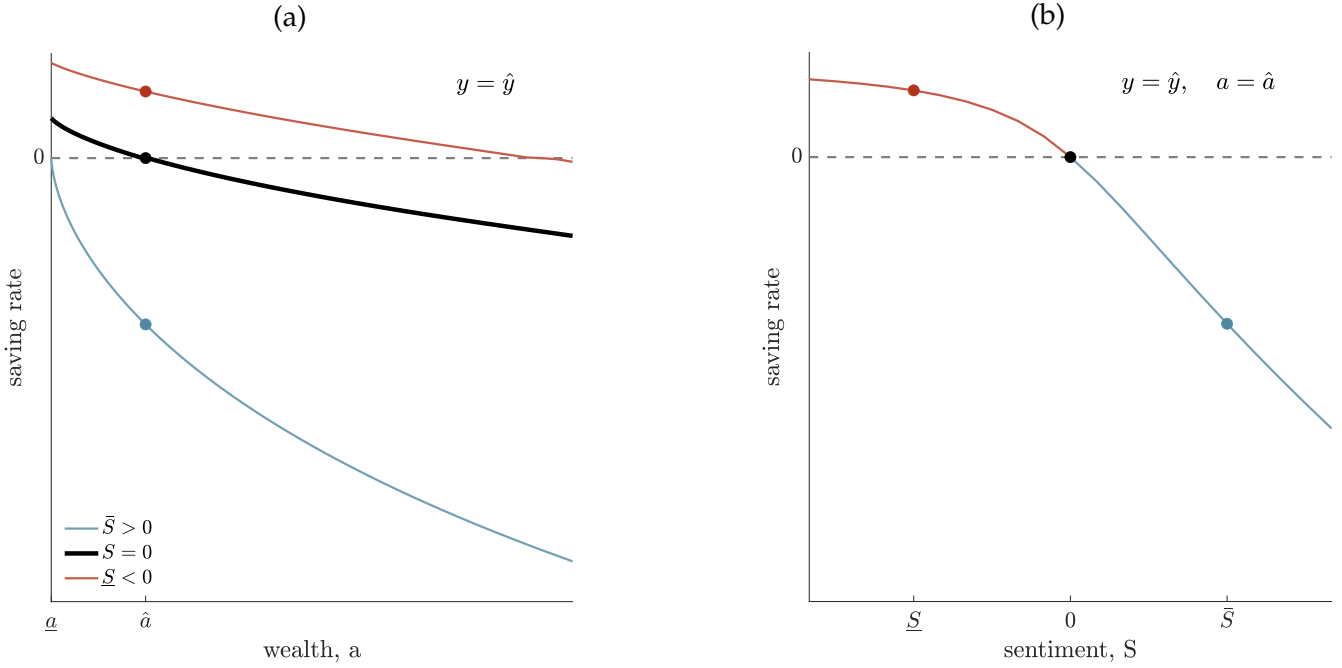
Effect on policy functions To go beyond analytical results we solve the model numerically and derive the policy functions.⁹ Figure 3 plots the saving rate policy functions, first over assets a in panel 3a and then over our additional state variable sentiment \mathcal{S} in panel 3b. We define the saving rate as the flow of savings \dot{a}_t divided by the total flow of income $ra_t + we^{yt}$. A negative saving rate implies the agent is dissaving. Each panel plots the policy functions for a specific value of log-productivity, denoted by \hat{y} .

In panel 3a, the black curve depicts the policy function for the case with zero sentiment, which, as discussed before, coincides with the policy function under rational expectations,

⁸Note that introducing a risky asset (as in, for example, Benhabib et al. (2015)) would break this result. In particular, in the case in which idiosyncratic risk comes from capital income, the distortion on the Euler Equation would be equal to $\frac{\theta \mathcal{S}_t \times a \times MPC(x_t)}{IES(x_t)}$ where MPC represents the derivative of the consumption function with respect to wealth. This extensions is outside of the scope of this paper, but we see this as a promising direction for future work.

⁹Section 4 describes our numerical approach and calibration in more detail and Section 5 derives the quantitative implications.

Figure 3: Saving rate policy function



Note: The dots represent the saving rate for a given level of wealth \hat{a} and income \hat{y} and different levels of sentiment $S \in \{\underline{S}, 0, \bar{S}\}$

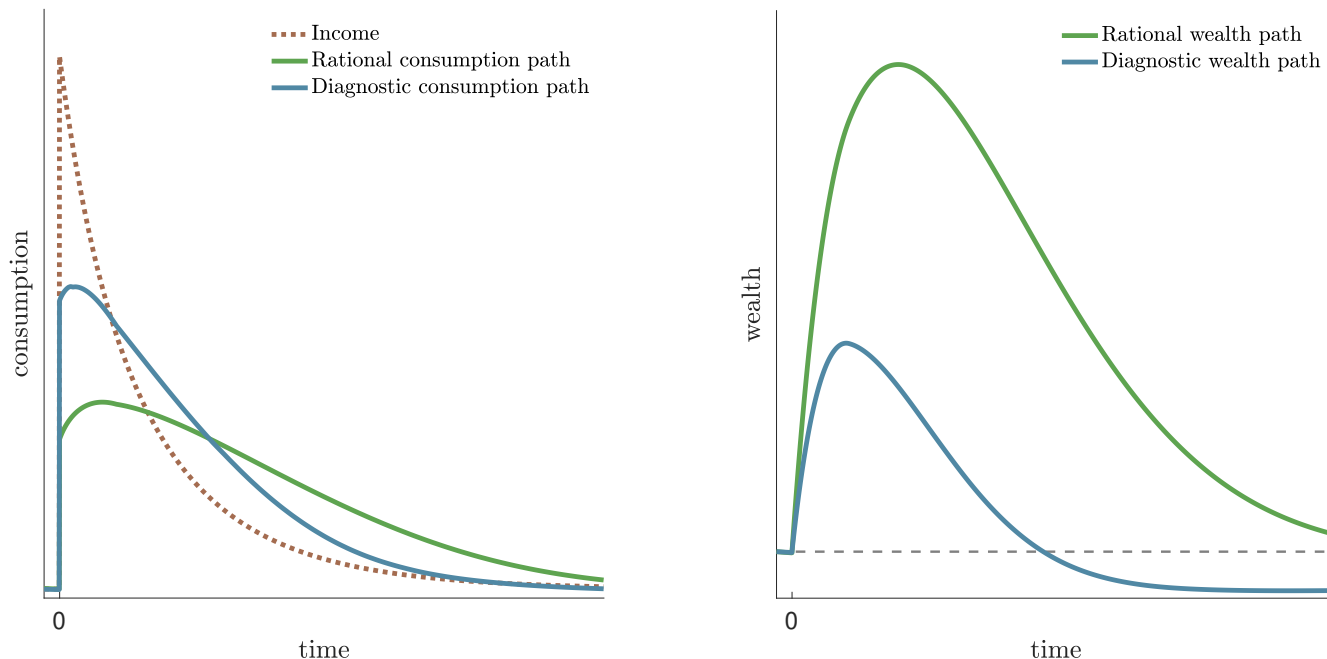
since we are focusing on the naïve case. The blue line illustrates the effects of positive sentiment on the agent’s decisions. Positive sentiment leads households to consume more than they would under full rationality, because they are now over-optimistic about their future income prospects. In the particular case we are consuming, positive sentiment pushes the agent to a hand-to-mouth state with zero savings, when they are at the borrowing limit. The red line, instead, depicts the case of negative sentiment. Quite intuitively, the effect on the agents’ behavior are opposite to the previous case. Because households are now overly pessimistic about the future evolution of their income, they save more today. Panel 3a plots the savings rate for different levels sentiment, keeping fixed the asset state at \hat{a} , which we report in panel 3a for reference. We can see how there are strong *non-linearities* in the role that sentiment plays for households’ saving decisions. In particular, while negative level of sentiment only induce mild distortions on agents’ behavior, positive sentiment has large effects on the household’s choice. These non-linearities will play an important role in the analysis in Section 6.

3.2 Consumption with sentiment dynamics

After having analyzed the role of sentiment on households’ behavior in the state space, we now turn our attention to the sequence space. In particular, because sentiment reacts instantaneously

to income shocks, it turns out to have interesting implications for agents’ consumption-saving *dynamics*. Figure 4 illustrates this point, by showing the effects of a positive income shocks on the path of consumption and savings when agents have diagnostic expectations. In particular, we consider a positive shock to log-productivity –here equal to one standard deviation of the typical income shock σ – which takes place at time 0. To ease the exposition, we shut down all shocks to the income path after time 0.¹⁰

Figure 4: Consumption and wealth dynamics



Note: We start at $y_0 = 0$ and $S_0 = 0$ and wealth equal to 23% of the average wealth in the stationary distribution (this is so that the consumption path before the shock is flat, for exposition purposes). We shock log-productivity at time 0. We shut down all the other shocks over time to isolate the effects of the shock. We use our benchmark calibration described in Table 1.

The income path is represented by the dotted brown line in Figure 4. The green line then depicts the paths for consumption and savings that would be chosen under rational expectations. It can be seen how consumption jumps upon impact, then still increases slowly for some periods as the agent builds assets, before decreasing continuously as income keeps mean reverting. Savings display a similar path, as the household engages in intertemporal smoothing, but with a smoother pattern. When agents are diagnostic, the positive jump in sentiment induced by the shock leads the agent to overreact on impact, thus increasing consumption by more and savings by less than they would under rational expectations. The reason for this over-reaction

¹⁰Note that the shock we are considering is different from standard “MIT shocks”. In fact, in our case agents do not have perfect foresight about the future path of their idiosyncratic income and hence still face risk. This is the reason why in Figure 4 consumption doesn’t immediately jump to its maximum before mean-reverting.

is that the diagnostic household perceives their income to be higher in the future than what it actually will be. For this reason, they accumulate less assets to be used for future consumption smoothing, and thus consume more today. Over time, however, sentiment reverts back to zero and the consumption path of the diagnostic agent progressively reverts to the rational one. In fact, the diagnostic agent’s consumption level eventually goes below that of the rational agent. This is because during the “enthusiasm phase”, the diagnostic household over-consumes, thus depletes their assets. As sentiment fades, the household is left with less wealth than it would have in the rational counterfactual, and hence ends up consuming less. This illustrates how in our framework short run over-consumption translates into under-consumption in the long run, as inter-temporal mistakes propagate through time via wealth adjustments.

4 Evidence of sentiment in survey data

In this section, we provide evidence that households have diagnostic expectations when forecasting their own future income. We show this by relying on data from the Survey of Household Income and Wealth. We then rely on this empirical evidence to calibrate the psychology parameters of our model.

4.1 Data

Our data come from the Survey of Household Income and Wealth (SHIW), which is run biannually by the Bank of Italy.¹¹ The SHIW is a representative survey of Italian households, featuring a rotating panel component. It includes detailed and disaggregated data on households’ income, assets, and liabilities, and also provides information on consumption and saving behavior, as well as demographic characteristics. Crucially for our analysis, the 2012 and 2014 waves of the survey also asked respondents to report their income expectations for the following year.¹² Throughout the rest of our empirical analysis, the focus is going to be on these two waves. In addition, we also obtain data on realized income for the neighboring 2010 and 2016 waves. Overall, a total of 4,140 and 8,156 households reported their income expectations in the 2012 and 2014 waves respectively. However, once we restrict the sample to those respondents who also appear in the following wave of the survey –in order to be able to compare expected and realized income– the effective size of our sample shrinks to 1,288 households in 2012 and 2,038 in 2014.

¹¹This survey has been used extensively in the literature, see for example [Jappelli and Pistaferri \(2014, 2020\)](#), [Auclert \(2019\)](#).

¹²For the exact wording of the questions, see [Appendix C.1](#).

Respondents' expectations for future individual income were also elicited in the 1989 and 1991 waves of the survey. In [Appendix C.2](#) we show that all our results hold when we consider this different time period.

Income Expectations Because of the biannual nature of the survey, there is always a one year gap between our data on income expectations and realizations. For example, in the 2014 wave of the survey respondents were asked to report their expected income for the year 2015, but data on realized income are available only for the next survey wave, conducted in 2016.¹³ To address this issue, we construct our measure of time t 's expectations for income in $t + 2$ —which we denote by $y_{i,t+2}^e$ —by extrapolating expectations for income growth between t and $t + 1$. In particular, households are asked to report their expected income growth from year t to $t + 1$, which we define as $g_{i,t+1|t}^e \equiv \frac{y_{i,t+1}^e}{y_{i,t}}$, where $y_{i,t}$ denotes household i 's realized income in year t . We then construct income expectations for year $t + 2$ as:¹⁴ $y_{i,t+2}^e = y_{i,t} \times \left(g_{i,t+1|t}^e\right)^2$.

Realized Income For each household i and year t in our sample we also collect data on total net realized income, which is defined as the sum of income from labor, pension and transfers, self-employment, and capital net of income taxes. We denote this variable by $y_{i,t}$. In the rest of our analysis, we drop retired households, which we define as those households having zero labor and self-employment income, but positive income from pension and transfers. However, all of our results still hold when we consider the full sample, or when we drop self-employed households.

Because the 1989 and 1991 waves of the survey asked respondents about their expected future *individual* labor income, $y_{i,t}$ denotes individual labor income when we analyze these years in [Appendix C.2](#).

Forecast Errors Armed with data on both realized and expected income, we are now ready to construct our main variable of interest. In particular, we define household's i forecast error in period t as the percentage difference between realized and expected income in year $t + 2$, that is:

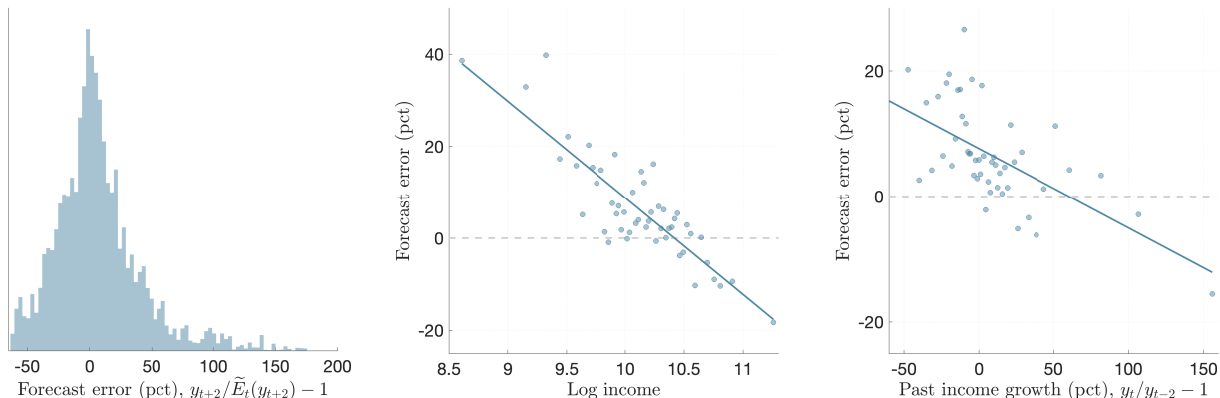
$$FE_{i,t} = \frac{y_{i,t+2}}{y_{i,t+2}^e} - 1 \quad (11)$$

Thus, according to our definition, households who make positive forecast errors turn out to be overly pessimistic when predicting their future income, and vice-versa. Note that because of

¹³Note that this feature is also shared by [Rozsypal and Schlafmann \(Forthcoming\)](#). In their setting, respondents are asked to report expected income for the next year, but data on realized income are available only for 6 months ahead income.

¹⁴Note that our results are virtually unchanged if we just assume that $y_{i,t+2}^e = y_{i,t} \times \left(g_{i,t+1|t}^e\right) \equiv y_{i,t+1}^e$.

Figure 5: Three motivating facts on households' income expectations



Note: Forecast errors are computed according to (11). The second panel displays a binned scatter plot of FE_t against $\log(y_t)$, where we define current income y_t as total household income net of taxes and capital income. Results are unchanged if we consider total household income including capital income. The third panel shows a binned scatter plot of FE_t against the past income change. We group observations in 75 bins. In both panels we residualize the x and y axis by time fixed effects, number of members of the household, area of residence, as well as age category, sex, and educational attainment of the respondent. In the third panel we also control for household's income quintile. We trim forecast errors at the 98th percentile.

the way we define it, our forecast error variable is bounded below by -1 , but is unbounded above. Throughout the rest of our analysis, we thus trim our forecast error variable at the 98th percentile.¹⁵

Other Data Finally, we also collect data on households' area of residence, number of components, and net wealth –defined as the total value of real estate, businesses, valuables, and financial wealth owned, net of debt and mortgages– as well as age, sex, educational attainment, occupation, and sector of employment of the main respondent within the household. Throughout our analysis, we always report results weighted by survey weights.

4.2 Evidence of Over-Extrapolation in Households' Income Expectations

We now document three motivating facts on households' income expectations formation process. Our focus is on the forecast error households make when predicting their future income, as defined in (11).¹⁶

The first panel of Figure 5 shows that there is large dispersion in the errors households make when forecasting future income. Note that this fact alone is not in direct contradiction with rational expectations. In fact, even in the rational expectations benchmark, the cross-section of

¹⁵Note that the particular cut-off for trimming data is immaterial for our results, as we show in Appendix C.2.

¹⁶In Appendix C.2 we show that all our facts also hold when we consider the 1989 and 1991 waves of the survey, in which questions about expected future income were also asked.

forecast errors follows a distribution with some non-zero dispersion. In particular, under rational expectations, the distribution of forecast errors would simply mimic that of idiosyncratic income shocks.

Second, households' forecast errors are negatively correlated with households' income. This is shown in the middle panel of [Figure 5](#), where we plot a binned scatter plot of the logarithm of household income against their forecast error. We residualize both axis by time fixed effects, number of members of the household, area of residence, as well as age category, sex, and educational attainment of the respondent.¹⁷ We find that high income households tend to be overly optimistic when forecasting future income, while low income ones tend to be excessively pessimistic. This fact is at odds with the predictions of the rational expectations hypothesis. In fact, under rational expectations forecast errors should be unpredictable. This result is in line with previous evidence for the US context based on the Michigan Survey of Consumers ([Rozsypal and Schlafmann, Forthcoming](#)).

Third, and finally, we document that households' errors in predicting future income not only correlate with income *levels*, but also with past income *changes*. More precisely, even after controlling for the level of income, households that experienced an income increase in the past tend to be excessively optimistic about the path of their future income, and vice-versa. We show this in the third panel of [Figure 5](#), by means of a binned scatter plot of income growth from year $t - 2$ to t against the forecast error $FE_{i,t}$ defined in (11). We include the same controls as for the second panel, but in this case we also control for the household's income quintile. To the best of our knowledge, we are the first to document correlation between past income changes and future forecast errors. Once more, this fact cannot be rationalized by rational expectations.

Taken together, we believe these three facts provide motivating evidence that households' expectations formation process for future idiosyncratic income cannot be completely approximated by rational expectations. In particular, our second and third facts suggest that households tend to over-extrapolate past income changes when forming expectations about the future, a pattern which has already been showed for aggregate variables, see in particular [Bordalo et al. \(2020\)](#). Moreover, insofar as households' income process is at least partly idiosyncratic across households, so that different households experience different histories of past income shocks, this is going to lead to a distribution of households' optimism/pessimism about future income, exactly as documented in our first fact. We now use these last two facts to put some discipline on our theory.

¹⁷In [Appendix C.2](#) we show that this pattern is robust to the particular set of controls included, and in particular also holds unconditionally, as well as after controlling for household wealth.

4.3 Matching the psychology parameters

In this section we provide a detailed description of how we map our model to the patterns for forecast errors observed in the data.

Computing forecast errors in the model As explained in Section 2, for each point of the state space we can leverage the Kolmogorov Forward equation to characterize (i) households' beliefs about future productivity $\tilde{f}(y_{t+h}|x_t)$ for any horizon h , and (ii) the true statistical distribution $f(y_{t+h}|x_t)$. In order to match our empirical evidence, we focus on a 2 years ahead horizon, i.e., $h = 2$. Then, for each point x_t in the state space, we draw N realizations $\{y_{t+2}^n\}_{n=1}^N$ of y_{t+2} from the true distribution $f(y_{t+2}|x_t)$.¹⁸ We also compute agents' expected income $\tilde{\mathbb{E}}_t(\exp(y_{t+2}))$, as described in (8). Finally, we construct the realized forecast errors using $FE^n(x_t) = \exp(y_{t+2}^n)/\tilde{\mathbb{E}}_t(\exp(y_{t+2})) - 1$ for each $n = 1, \dots, N$. We weight each point of the state space by its stationary mass $G(da, dy, d\mathcal{S})$.

Calibrating θ and κ We calibrate θ and κ jointly by targeting the slope of the linear relationship between (i) forecast errors and current income and (ii) forecast errors and past income growth that we estimate in the data.¹⁹ Figure 8 shows the results of this calibration. Our model is able to account for the empirical correlation between forecast errors, current income, and past income growth –something that would not be possible to replicate in a rational expectations framework. Note that θ and κ act on the slope of the two curves via different channels. θ directly controls the extent to which sentiment influences the forecast error. κ , on the other hand, has no effect on the forecast when the agent is naïve, but controls the variance of the sentiment distribution. Hence κ controls the slope of these two curves via composition effects.

In the next iteration of the draft, we will use the correlation between the forecast error in year 2012 and the forecast error in year 2014 to put further discipline on κ . Since the forecast errors are very correlated in our data, this will lead to a more stable calibration of κ .

4.4 Calibration of structural parameters

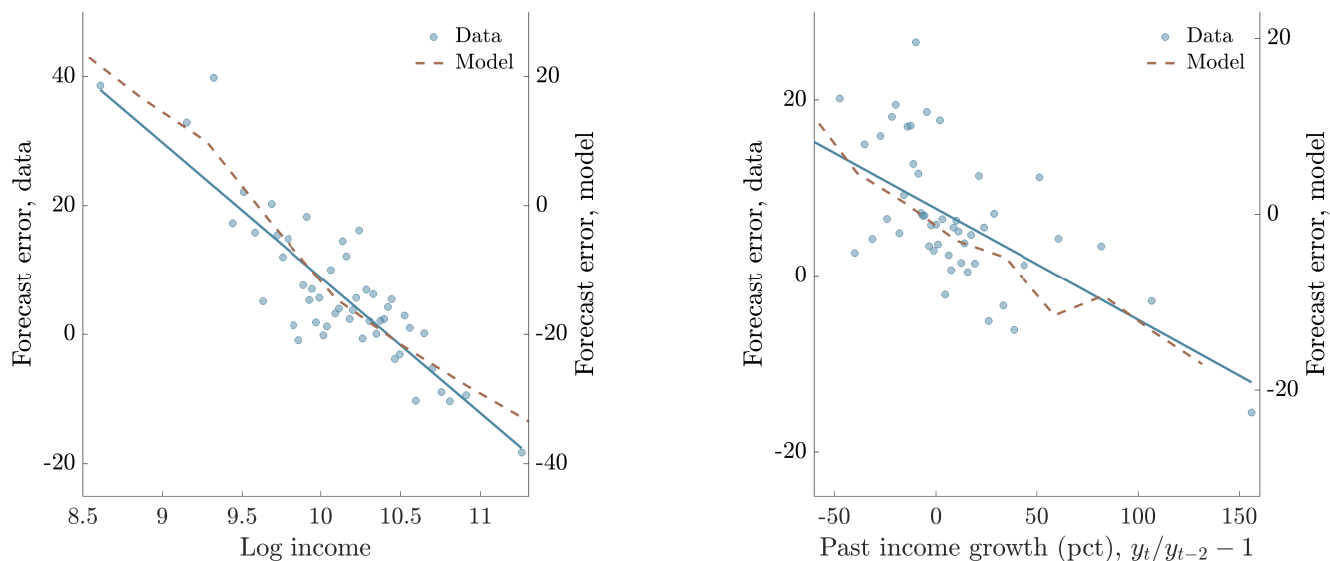
We solve the model numerically using the finite-differences scheme proposed in Achdou et al. (2022).²⁰ Table 1 provides the parameter values we use for the calibration. We calibrate the model to a quarterly frequency.

¹⁸In practice we use $N = 100$.

¹⁹Since agents in our data tend to have a positive forecast error on average we don't target the intercept.

²⁰See Appendix B for a detailed description of our numerical algorithm.

Figure 6: Matching forecast error data



Note: Forecast errors and the income change are expressed in percentage deviations: $FE_t \equiv y_{t+2}/\mathbb{E}_t y_{t+2} - 1$

Income Our income process is similar to the one described in [Kaplan et al. \(2020\)](#). We use their calibration of μ and λ , and calibrate σ to match the standard deviation of log-income in the SHIW data, which is 0.616. We then set the wage such that mean labor income is normalized to 1. In the next iteration of the draft we'll rely on the procedure described in [Kaplan et al. \(2018\)](#), thus using income changes.

Wealth Following [Kaplan and Violante \(2022\)](#), we calibrate ρ such that the average wealth to average income ratio in our model is equal to the one observed in the data, which is 9.69 in our case. We calibrate ρ separately for the rational and diagnostic model, to ensure that the wealth to income ratio is the same in both cases. However, calibrating both models with a common ρ does not change our results.

Other structural parameters As is standard in the literature, we set the intertemporal elasticity of substitution to one. Finally, we follow [Kaplan and Violante \(2022\)](#) we set the partial equilibrium interest rate to 1% per annum and the borrowing limit to 0.

Table 1: Model Calibration

Parameter	Description	Value	Justification
<i>Preferences</i>			
ρ^{DE}	Discount rate (p.a.)	3.9%	Match wealth to income ratio
ρ^{RE}	Discount rate (p.a.)	4%	Match wealth to income ratio
γ	Inverse IES	1	Standard
<i>Diagnostic Expectations Parameters</i>			
θ	Diagnosticity	4%	Calibrated
κ	Decay of New Information	0.25%	Calibrated
<i>Income process</i>			
λ	Arrival rate of income shocks	3.46%	Kaplan et al. (2020)
μ	Mean reversion rate of income	3.48%	Kaplan et al. (2020)
σ	Standard deviation of shocks	0.736	Match standard deviation of log-income
<i>Other structural parameters</i>			
r	Interest rate	1%	Kaplan and Violante (2022)
w	Wage	0.82	Normalize mean income to 1
\underline{a}	Hard borrowing limit	0	Kaplan and Violante (2022)

Note: All parameters are expressed at quarterly frequency unless indicated otherwise.

5 Steady state household behaviour

In this section, we present three key predictions for steady-state household behavior under our diagnostic expectations framework. First, we show that sentiment amplifies the consumption response to income shocks, thus providing a potential rationalization for the “excess sensitivity” observed in the data. Second, sentiment generates latent heterogeneity in the consumption response to income and wealth shocks, which has been recently emphasized in the empirical literature. Third, our diagnostic expectation framework predicts a larger persistence of the hand-to-mouth state than the rational expectations benchmark. We also show that the stationary distribution of our model features a substantial fraction of hand-to-mouth households, which more closely matches the distribution observed in the data, without the need to rely on the presence of an illiquid asset. Finally, we show that our behavioral friction has non-trivial welfare cost. Moreover, these welfare costs are heterogeneous in the cross-section of households, and are substantially larger for poorer agents.

5.1 Anatomy of diagnostic hand-to-mouth households

In this section we show that our economy features a larger mass of hand-to-mouth (HtM) households than in the rational benchmark, thus providing a much better fit of the data. Moreover, diagnosticity makes the HtM state “stickier” in the sense that, compared to the rational benchmark, once diagnostic agents enter the HtM state they have a lower probability of escaping it at any point in time.

The hopeful Hand-to-Mouths We follow [Kaplan et al. \(2014\)](#) and adopt a standard definition of hand-to-mouth households:

Definition 1 (Hand-to-Mouth (HtM)). *The region of the state space such that the household is HtM is denoted by \mathcal{H} and is the set of states such that household’s wealth is less than half their monthly income: $\mathcal{H} = \{(a, y, \mathcal{S}) | a \leq we^y / 6\}$*

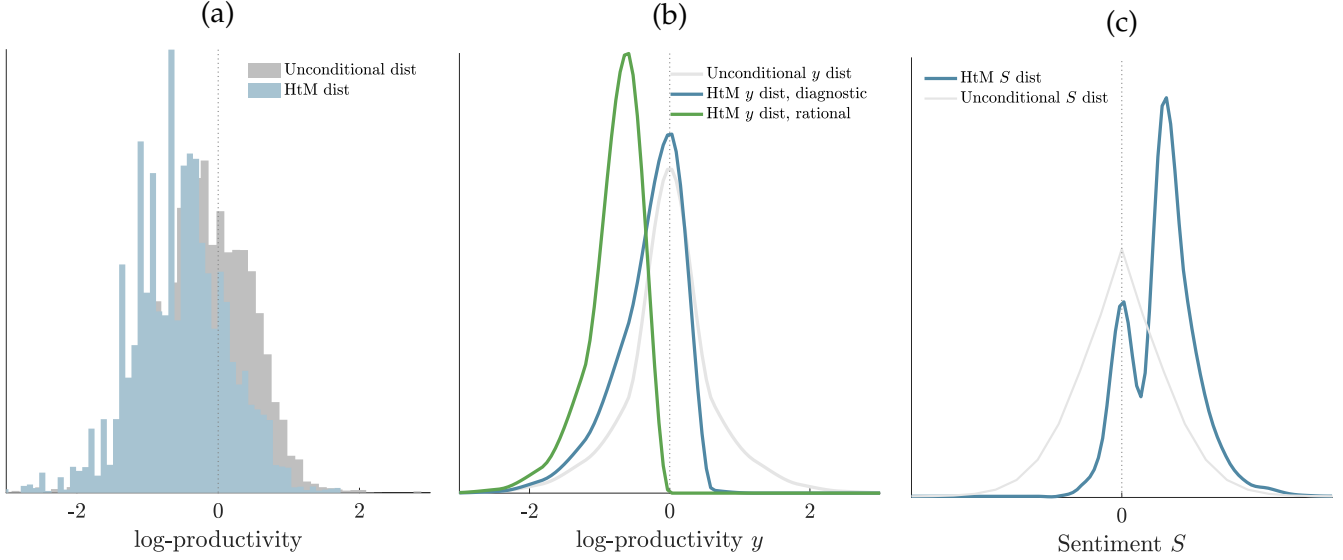
The stationary distribution of our diagnostic economy features a share of 20.1% of agents in the HtM state as defined in Section 5.1. Despite the fact that we do not target this moment when calibrating our model, this figure is very close to the empirical estimates for the Italian economy. In fact, [Kaplan et al. \(2014\)](#) find that the share of HtM households in the Italian economy is around 20%, using data from the Household Finance and Consumption Survey (HFCS) for the period 2008-2010. When we estimate the share of HtM agents in our data, we find it to be approximately 21%.²¹ Note that our rational expectations benchmark model is not able to match the share of HtM households observed in the data. In particular, only 10.2% of agents are HtM in the stationary distribution of this model. In fact, the literature usually resorts to ad-hoc modeling tools in order to match the share of HtM households observed in the data, such as the introduction of illiquid assets ([Kaplan and Violante, 2014](#)) or preference heterogeneity ([Aguiar et al., 2020](#)).

It turns out that our model is also able to produce a better empirical fit when it comes to the *composition* –as opposed to the *mass*– of HtM households. Notably, compared to the rational benchmark, HtM households in the diagnostic model exhibit substantially higher average income. This can be seen from [Figure 7b](#), which plots the conditional log-productivity distribution for hand-to-mouth households, comparing the rational case (green) and diagnostic case (blue).²² The reason for this discrepancy is depicted in [Figure 7c](#): a significant portion of HtM households in the diagnostic model exhibit positive sentiment. These households, which

²¹We define HtM households as those households whose liquid assets are below half of their non-capital monthly income.

²²For reference, we also report the unconditional distribution in grey. Naturally, this distribution is unaffected by the level of rationality.

Figure 7: Hand-to-mouth composition



Note: In panel 7a, we compute log-productivity as the data analogue of our model: we divide income by average income and take log. We follow section 5.1 to define the hand-to-mouth households in our model and data (we follow the literature in using liquid wealth for the data definition).

we call the “hopeful” hand-to-mouth, are middle-income households that have experienced extended spells of positive sentiment. As a result, they depleted their wealth and have become financially constrained, without necessarily being in a low income state. Thus, the presence of sentiment gives rise to a mass of middle-income, optimistic, hand-to-mouth households. This is qualitatively in line with the empirical evidence in our data. In fact, Figure 7a shows the empirical analog of Figure 7b in the SHIW data. We find that around 20% of Italian households that are classified as HtM have an income level which is above the cross-sectional average. Once more, our diagnostic model is decently able to match this untargeted moment, with around 29% of HtM agents featuring above average income. In the rational benchmark, on the other hand, virtually all HtM households have below average income.

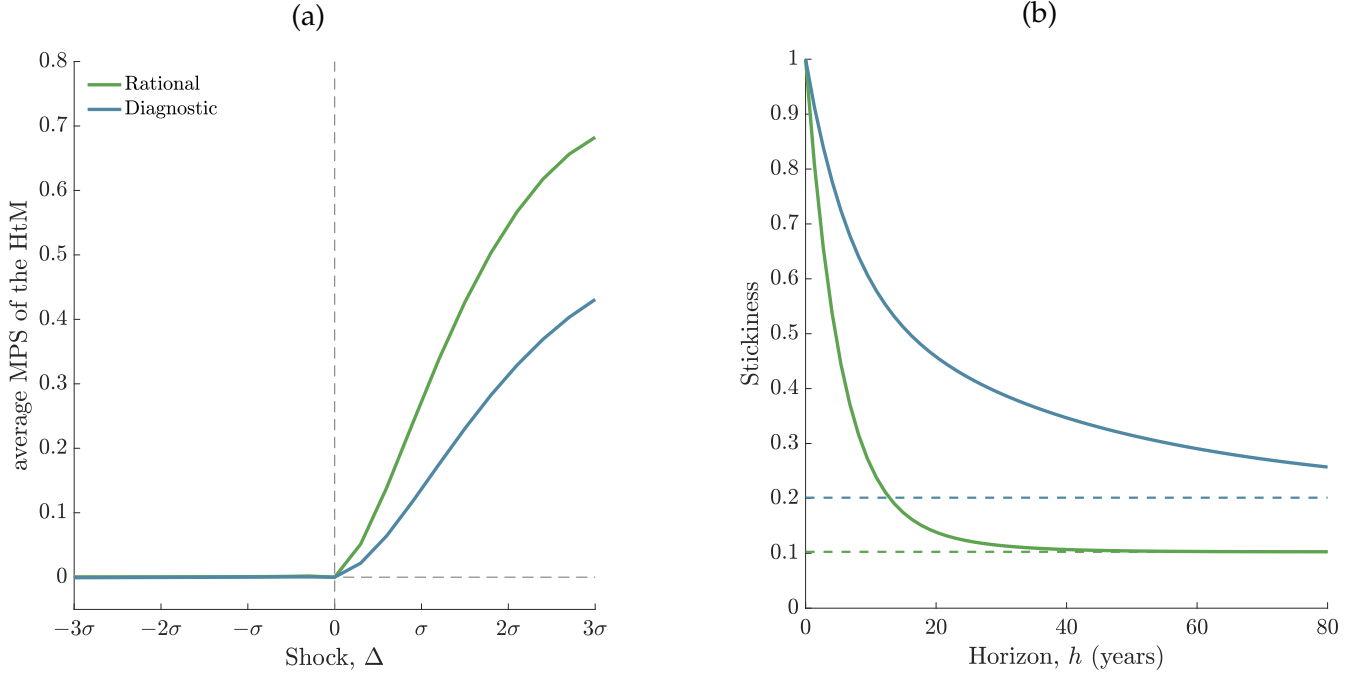
Sticky hand-to-mouth Before discussing the dynamics of the HtM state, it is useful to define the Marginal Propensity to Save (MPS) out of log-productivity shocks Δ .

Definition 2. The Marginal Propensity to Save out of productivity shocks Δ , for a household with state vector $x \equiv (a, y, \mathcal{S})$ over a period τ is given by

$$s_{\tau}(\Delta; x) = \frac{Sav_{\tau}(a, y + \Delta, \mathcal{S} + \Delta) - Sav_{\tau}(a, y, \mathcal{S})}{e^{\Delta} - 1} \quad (12)$$

$$\text{where } Sav_{\tau}(x) = \mathbb{E}_0 \left[\int_0^{\tau} s^{DE}(x_t) dt \mid x_0 = x \right] \quad (13)$$

Figure 8: Stickiness



where $s^{DE}(\cdot)$ is the diagnostic expectations savings policy function and \mathbb{E}_0 is the rational expectations operator with respect to all the state variables.

Consistently with the non-linearities generated by sentiment and emphasized before, our model also predicts asymmetric effects of negative and positive log-productivity shocks on the savings rate of HtM agents. This is displayed in Figure 8a, which plots the MPS out of income shocks of different sign and size for HtM agents under the rational and diagnostic benchmark. Because we are focusing on HtM households, negative shocks have no effect on the savings rate both under rational and diagnostic expectations, since the agent is constrained both before and after the shock. Positive shocks, on the other hand, may push the agent out of the HtM region, thus inducing them to save out of the shock. When agents are diagnostic, however, the MPC out of sentiment puts downward pressure on this saving motive. In fact, sentiment jumps up in response to the shock and, as a result, agents don't save as much as in the rational benchmark, thus remaining closer to the borrowing limit. The lower MPS out of positive shocks induced by sentiment implies that under diagnostic expectations it is more difficult for agents to escape the HtM state. To quantify this effect, we now define a concept of "stickiness" to evaluate the extent to which agents are "trapped" in the HtM state.

Definition 3 (Stickiness). We define the horizon- h stickiness of \mathcal{H} , $\mathfrak{S}_h(\mathcal{H}|x)$ as the average probability

that a HtM households is still HtM h years in the future:

$$\mathfrak{S}_h(\mathcal{H}) = \mathbb{E}_x [\mathbb{P}(x_h \in \mathcal{H} | x_0 = x) | x \in \mathcal{H}]$$

where the expectation \mathbb{E}_x is taken over the stationary distribution.

Stickiness thus captures how easy it is to “escape” the hand-to-mouth condition. To characterize $\mathfrak{S}_h(\mathcal{H})$ we once again rely on the true evolution of the statistical distribution of the states x_t :

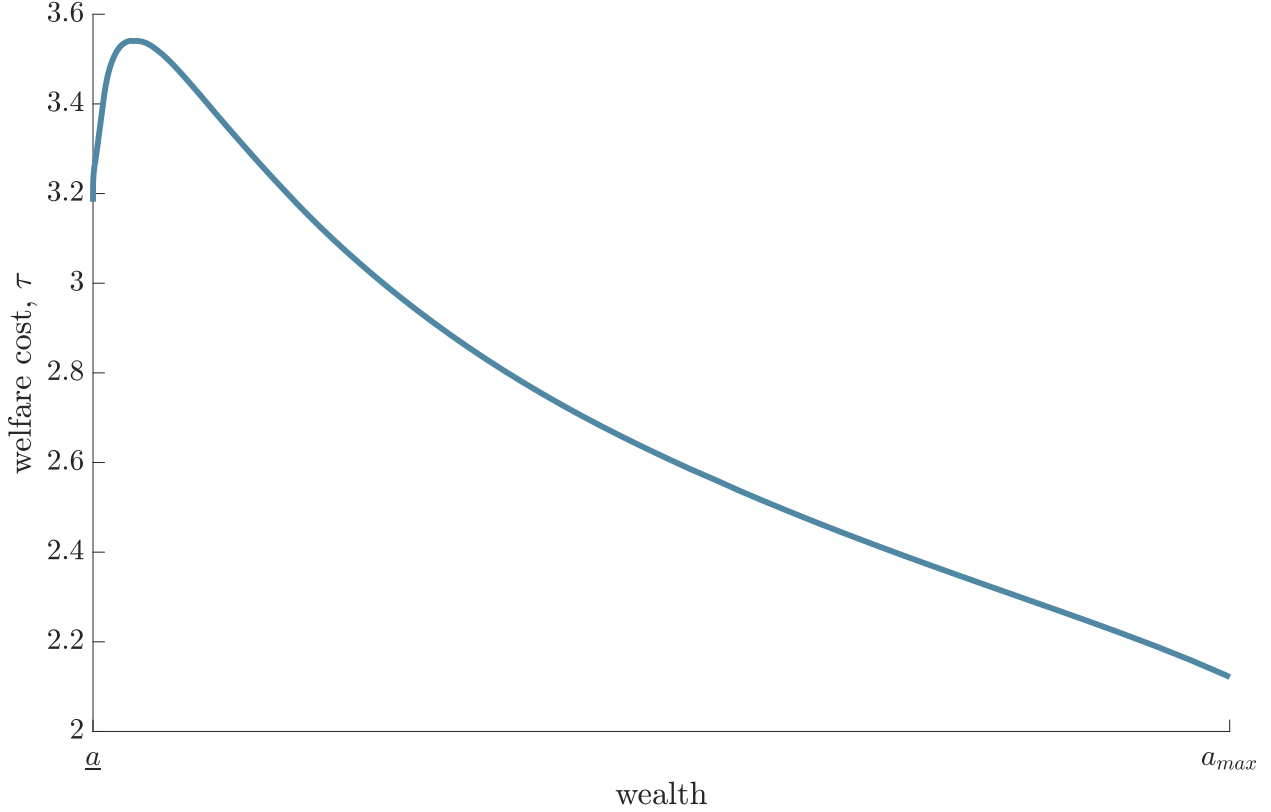
$$\mathfrak{S}_h(\mathcal{H}) = \int_{\mathcal{H}} dF_h(dx) \tag{14}$$

where $dF_h(dx)$ is obtained using equation (7) with boundary condition $F_0(dx) = G(dx)$, where $G(dx)$ is the stationary distribution.²³ Clearly, stickiness for horizon $h = 0$ is equal to 1, i.e., $\mathfrak{S}_0(\mathcal{H}) = 1$. Moreover, $\mathfrak{S}_h(\mathcal{H})$ converges to the stationary mass of HtM households as the horizon h goes to infinity. [Figure 8b](#) depicts the stickiness of the HtM state for both the rational and the diagnostic model. In the rational model, stickiness decays at a much higher rate than in the diagnostic one, taking about 40 years to reach its stationary value. In the diagnostic model instead, the average probability that an agent is still HtM in 40 years is still nearly twice its stationary level. Because of the mistakes induced in consumption-saving decisions, diagnostic expectations thus makes it more difficult for agents to escape the HtM state and in this sense it generates a poverty trap. To check that we are indeed capturing a poverty trap and not generally slower transition dynamics generated by diagnostic expectations, we run a placebo test in [Appendix A.3](#). In particular, [Figure A.1](#) analyzes the stickiness of the top 0.1% of the wealth distribution state in the rational and diagnostic benchmarks. The difference in stickiness between the two models appear to be substantially smaller than in the HtM case, thus suggesting that diagnostic expectations does not produce slower dynamics in general, but has particularly stronger effects on the stickiness of the HtM state.

In a future iteration of the draft, we plan to empirically estimate the stickiness of the HtM state in the data. We then plan to compare the diagnostic and rational model in their ability to fit the data. If the diagnostic model is able to outperform the rational one in matching this untargeted moment as well, we then plan to discuss how our behavioral framework can provide an “anatomy of hand-to-mouth households”, accounting for their mass, composition, and dynamics.

²³Since there is a Dirac mass point at the borrowing limit we define \mathfrak{S} in terms of an integral over a measure rather than over a density.

Figure 9: Welfare cost along the wealth distribution



Note: The range for wealth used in this graph goes from 0 to the maximum net worth observed in our SHIW data, which is around 15 million euros. The welfare cost is expressed in percentage terms.

5.2 Welfare evaluation

Welfare metric Diagnostic expectations generate non-trivial welfare losses. To get a better understanding on how quantitatively important these losses may be, we develop a welfare evaluation in the spirit of Lucas (1987). First we evaluate welfare function from a paternalistic view point based on the true law of motion of the states. In particular we define $W^{DE}(a, y, S)$ to be the welfare of an agent using the diagnostic expectations policy function $c^{DE}(a, y, S)$ and starting with initial conditions $a_0 = a, y_0 = y, S_0 = S$:

$$W^{DE}(a_0, y_0, S_0) = \mathbb{E}_0 \int_0^{\infty} e^{-\rho t} u(c^{DE}(a_t, y_t, S_t)) dt$$

where \mathbb{E}_0 is the rational expectations operator capturing the true evolution of the states. Now suppose the diagnostic agent with no sentiment has access to a technology making their beliefs rational, with a cost expressed as a flow consumption tax τ . This consumption tax will depend

on the household’s initial conditions and will serve as our welfare metric. Formally:

$$W^{DE}(a_0, y_0, 0) = \mathbb{E}_0 \int_0^\infty e^{-\rho t} u \left[(1 - \tau(a_0, y_0)) c^{RE}(a_t, y_t) \right] dt$$

Where c^{RE} is the rational expectations policy functions. Assuming log utility we immediately get the welfare cost schedule as a function of the initial states $a_0 = a$ and $y_0 = y$:

$$\tau(a, y) = 1 - \frac{e^{\rho W^{DE}(a, y, 0)}}{e^{\rho W^{RE}(a, y)}} \quad (15)$$

Distribution of welfare cost Given our calibration, we find the welfare cost of diagnosticity, averaged across the wealth and income distribution, to be 3.3 percent. Compared to the rule of thumb in [Lucas \(1987\)](#) that a cost of 0.5 percent of lifetime consumption is “large”, our estimate indeed reveals a substantial welfare toll of diagnostic expectations. In particular, the interaction of income volatility and non-rational expectations leads households to commit intertemporal errors that prove to be very expensive in our model. Furthermore, these welfare costs are not uniformly distributed across the wealth distribution. As depicted in [Figure 9](#), cost tends to be decreasing in wealth. This is connected to our prior discussion that diagnostic expectations particularly distort decisions of low wealth households. Interestingly, though, the cost turns out to be increasing in wealth in the vicinity of the borrowing constraint. This is because when agents are exactly at the borrowing limit, the scope for making mistakes is largely reduced, as intertemporal decisions are constrained and beliefs do not matter as much. Finally, as wealth goes to infinity the welfare cost converges to zero.²⁴

6 Conclusion

In this paper we introduced diagnostic expectations in an otherwise standard incomplete market model. We proposed that over-extrapolation of recent income news can maintain agents into a state of financial constraint, as positive income shocks drive overoptimism and overconsumption. To develop this idea, we adapted the theory of diagnostic expectations to idiosyncratic shocks in incomplete markets, introducing a ‘rationality wedge’ to handle deviations from rational expectations with heterogeneous agents. This enables generalizing diagnostic expectations beyond traditional AR(1) processes and introduce this behavioral friction to more standard quantitative models of income fluctuation. We suggested that households’ perceptions of future income are distorted by ‘sentiment’, a new state variable in an incomplete

²⁴This result is reminiscent of the result in [Allais et al. \(2020\)](#), who find that the welfare costs of volatile inflation vanish for agents at the top of the wealth distribution.

market framework, together with wealth and productivity. We validated empirically the way sentiment can distort households' expectations using survey data on households' expectations from the Survey of Household Income and Wealth. Households having experienced growth in their income tend to forecast future income above their actual realized value, a bias linked to our sentiment variable. Lastly, we found that diagnostic expectations increase the persistence of the hand-to-mouth state. When agents are financially constrained, diagnostic expectations make it harder to escape financial constraints as agents have a higher marginal propensity to consume out of positive income shocks. In incomplete market models, diagnostic expectations can represent a significant cost, averaging 3.3 percent of lifetime consumption.

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A Proof of propositions

A.1 Proof of proposition 1

We prove the equation for a general log-productivity process. We’ll derive the proposition for a general a generator \mathcal{B} for the joint process of y and \mathcal{S} .²⁵ The HJB of the sophisticated agent is

²⁵We’ll impose the regularity condition that $\partial_a BV = B\partial_a V$, which is satisfied for our baseline income process.

given by

$$\rho V = \max_c u(c) + V_a(ra + we^y - c) + \mathcal{B}V + \mathcal{S}V_y \quad (\text{A.1})$$

From this we can get the first order condition and the envelop condition:

$$u'(c) = V_a \implies u''(c)c_a = V_{aa}, u''(c)c_y = V_{ay} \quad (\text{A.2})$$

$$\rho V_a = V_{aa}s + rV_a + \mathcal{B}V_a + \mathcal{S}V_{ya} \quad (\text{A.3})$$

From which we get

$$\rho u'(c) = u''(c)c_a s + ru'(c) + \mathcal{B}u'(c) + \mathcal{S}u''(c)c_y \quad (\text{A.4})$$

$$\iff u''(c)c_a s + \mathcal{B}u'(c) = (\rho - r)u'(c) - \mathcal{S}u''(c)c_y \quad (\text{A.5})$$

$$\iff \mathbb{E}(du'(c)/dt) = (\rho - r)u'(c) - u''(c)c_y \mathcal{S} \quad (\text{A.6})$$

$$\iff \frac{\mathbb{E}_t[du'(c_t)/dt]}{u'(c_t)} = \rho - r - \frac{u''(c_t)c_t}{u'(c_t)} \frac{\partial \log c_t}{\partial y} \mathcal{S}_t \quad (\text{A.7})$$

Using the definition of intertemporal elasticity of substitution and income elasticity of consumption we get the desired result.

For the naïve case, the HJB is given by:

$$\rho V = \max_c u(c) + V_a(ra + we^y - c) + \mathcal{I}V + \mathcal{S}V_y \quad (\text{A.8})$$

From this we can get the first order condition and the envelop condition:

$$u'(c) = V_a \implies u''(c)c_a = V_{aa}, u''(c)c_y = V_{ay} \quad (\text{A.9})$$

$$\rho V_a = V_{aa}s + rV_a + \mathcal{I}V_a + \mathcal{S}V_{ya} \quad (\text{A.10})$$

From which we get

$$\rho u'(c) = u''(c)c_a s + ru'(c) + \mathcal{I}u'(c) + \mathcal{S}u''(c)c_y \quad (\text{A.11})$$

$$\iff u''(c)c_a s + \mathcal{I}u'(c) = (\rho - r)u'(c) - \mathcal{S}u''(c)c_y \quad (\text{A.12})$$

$$\iff \mathbb{E}(du'(c)/dt) = (\rho - r)u'(c) - u''(c)c_y \mathcal{S} \quad (\text{A.13})$$

$$\iff \frac{\mathbb{E}_t[du'(c_t)/dt]}{u'(c_t)} = \rho - r - \frac{u''(c_t)c_t}{u'(c_t)} \frac{\partial \log c_t}{\partial y} \mathcal{S}_t \quad (\text{A.14})$$

Using the definition of intertemporal elasticity of substitution and income elasticity of consumption we get the desired result.

A.2 Feynmann-Kac equation

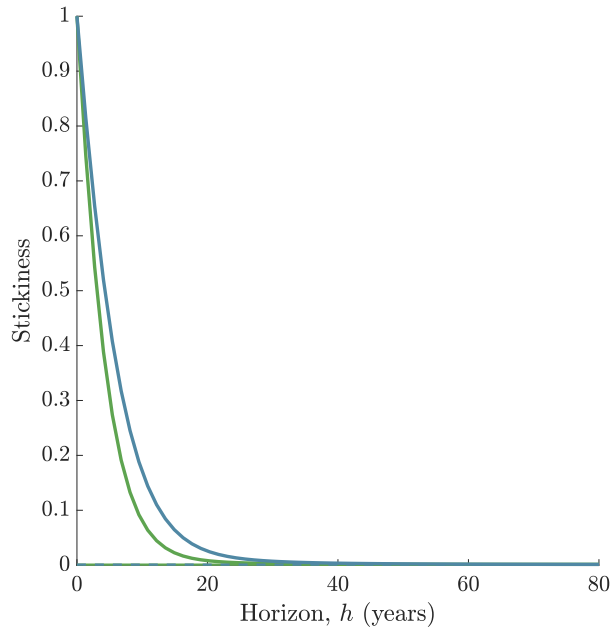
The conditional expectation $C_\tau(x)$ (13) can be computed as $C_\tau(x) = \Gamma(x, 0)$ where $\Gamma(x, t)$ satisfies the PDE

$$0 = c^{DE}(x) + \mathcal{G}^*\Gamma(x, t), \quad \text{where} \quad \mathcal{G} \equiv (ra + we^y - c^{DE}(x))\partial_a + \mathcal{B} + \partial_t \quad (\text{A.15})$$

with terminal condition $\Gamma(x, \tau) = 0$ for all x .

A.3 Additional figures

Figure A.1: Stickiness of the ultra-rich



Note: we define the ultra rich as the top 0.1% of the wealth distribution in each model.

B Numerical appendix

We can define the income process in the more standard way: at some poisson rate λ , the agent draws a new log-productivity drawn from a normal centered around zero and with variance σ^2 . In this case, upon the new draw, sentiment will move by a jump equal to the difference between the new draw and the previous log-productivity. Hence at some poisson rate sentiment is drawn from a normal with mean $-y$ and variance σ^2 . First, we discretize the exogeneous generator in

the HJB equation. In our case, this generator is given by

$$\mathcal{B}W(a, y, S) = -\beta y \partial_y V(a, y, S) - \eta S \partial_S V(a, y, S) + \lambda \int (V(a, x, S + x - y) - V(a, y, S)) \phi(x) dx$$

where $\phi(x)$ is a normal pdf with mean 0 and variance σ^2 .

Discretize We want to get the discretized operator \mathcal{B} , which captures the true transition for the joint process of income and sentiment.

$$\begin{aligned} \mathcal{B}v_{ijk} = & (v_{ij+1k} - v_{ijk}) \left(\frac{-\beta y_j}{\Delta y} \right)^+ + (v_{ijk} - v_{ij-1k}) \left(\frac{-\beta y_j}{\Delta y} \right)^- \\ & + (v_{ijk+1} - v_{ijk}) \left(\frac{-\eta \mathcal{S}_k}{\Delta s} \right)^+ + (v_{ijk} - v_{ijk-1}) \left(\frac{-\eta \mathcal{S}_k}{\Delta s} \right)^- \\ & + \lambda \sum_{j'=1}^J (v_{ij'(k+j'-j)} - v_{ijk}) \phi(y_{j'}) \Delta y \end{aligned}$$

The crucial step is to rearrange:

$$\begin{aligned} \rho v_{ijk} = & v_{ijk} Y_{ijk} & Y_{ijk} \equiv & \sum_{m=1}^4 X_{ijk}^m + \sum_{n=1}^J Z_{ijk}^n \\ & + v_{ij+1k} X_{ijk}^1 & X_{ijk}^1 \equiv & \left(\frac{-\beta y_j}{\Delta y} \right)^+ \\ & + v_{ij-1k} X_{ijk}^2 & X_{ijk}^2 \equiv & \left(\frac{-\beta y_j}{\Delta y} \right)^- \\ & + v_{ijk+1} X_{ijk}^3 & X_{ijk}^3 \equiv & \left(\frac{-\eta \mathcal{S}_k}{\Delta s} \right)^+ \\ & + v_{ijk-1} X_{ijk}^4 & X_{ijk}^4 \equiv & \left(\frac{-\eta \mathcal{S}_k}{\Delta s} \right)^- \\ & + v_{i1(k+1-j)} Z_{ijk}, & Z_{ijk} \equiv & \lambda \phi(y_1) \Delta y \\ & \vdots & & \\ & + v_{ij(k+J-j)} Z_{ijk}, & Z_{ijk} \equiv & \lambda \phi(y_J) \Delta y \end{aligned}$$

We call the X_{ijk} the coefficient meshes. We assume an equal grid for \mathcal{S} and y and hence we take $\min\{\max\{(k + j' - j), 1\}, J\}$.

C Empirical Appendix

C.1 Survey Questions

2012 Wave The following question was asked in the 2012 wave of the survey:

Twelve months from now, your household's income will be (please distribute 100 points):

Respondents were then asked to assign probabilities to 5 different scenarios: (i) higher than today (by 10% or more), (ii) somewhat higher than today (2 to 10%), (iii) basically the same (no more than 2% increase or decrease), (iv) somewhat lower (2 to 10%), (v) much lower than today (by 10% or more).

2014 Wave In 2014 the wording of the question changed as follows:

Consider your household's overall income in 2015. Compared with 2014, how much higher/lower do you think it will be in percentage terms?

In this case, respondents were asked to report their point forecast.

1989 Wave The following question was asked in the 1989 wave of the survey:

Consider the evolution of your total labor or pension income from now to May 1991. Please distribute 100 points among the following scenarios:

Respondents were then asked to assign probabilities to 12 different scenarios ranging from negative growth (for which they were also asked to report a point estimate of the percentage decrease in income) to above 25% income growth.

1991 Wave The following question was asked in the 1991 wave of the survey:

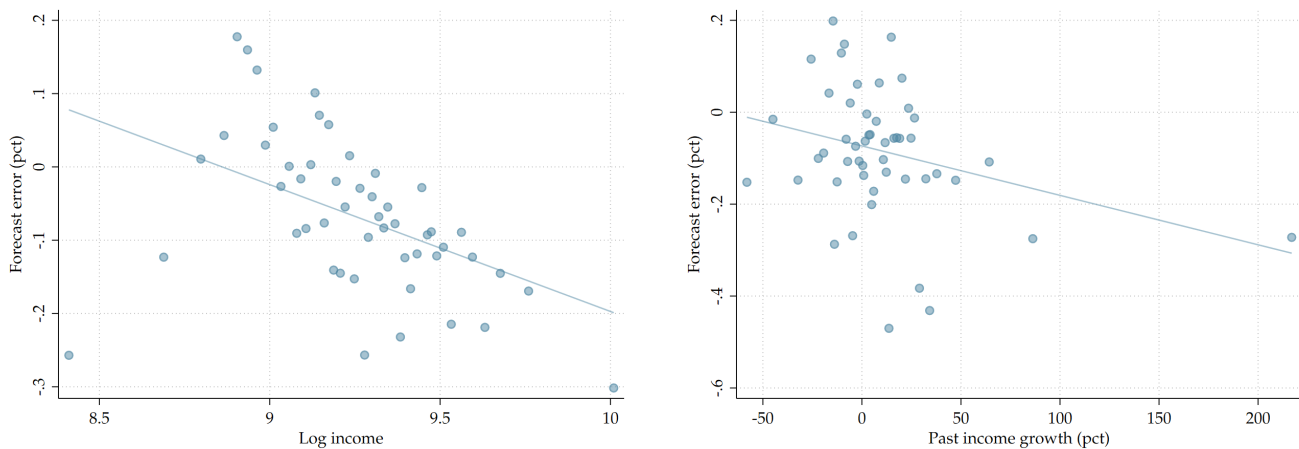
Consider your total labor or pension income one year from now. Please distribute 100 points among the following scenarios:

Respondents were then asked to assign probabilities to 12 different scenarios ranging from negative growth (for which they were also asked to report a point estimate of the percentage decrease in income) to above 25% income growth.

C.2 Robustness

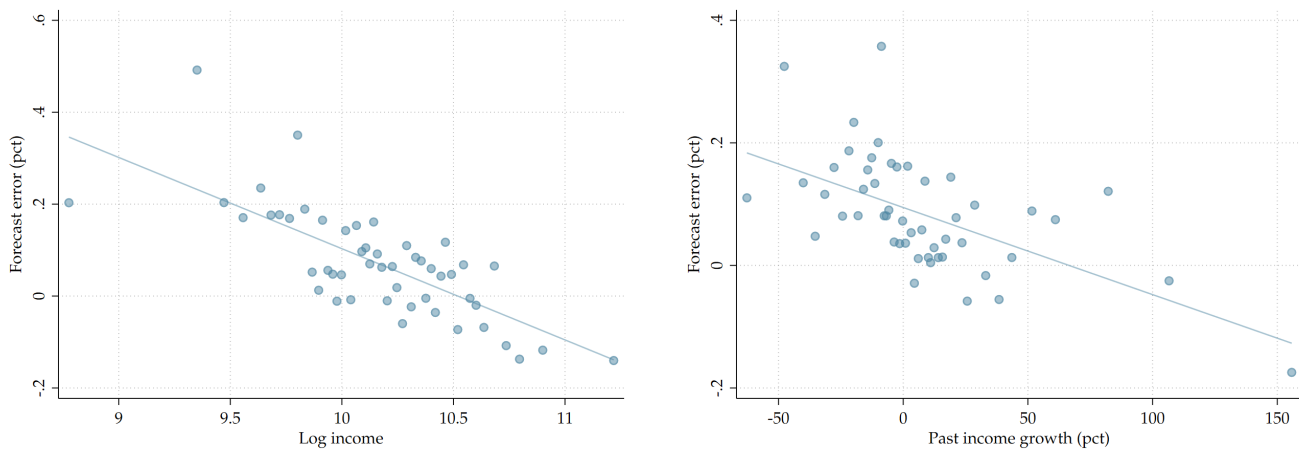
'89-'91 Waves Appendix C.2 below reproduces the last two panels of Figure 8 using data from the 1989 and 1991 waves of the SHIW. Because of the format of the question, both the y and the x axis of Appendix C.2 are expressed in terms of individual level, rather than household level, income.

Figure A.2: 1989 and 1991 Survey Waves



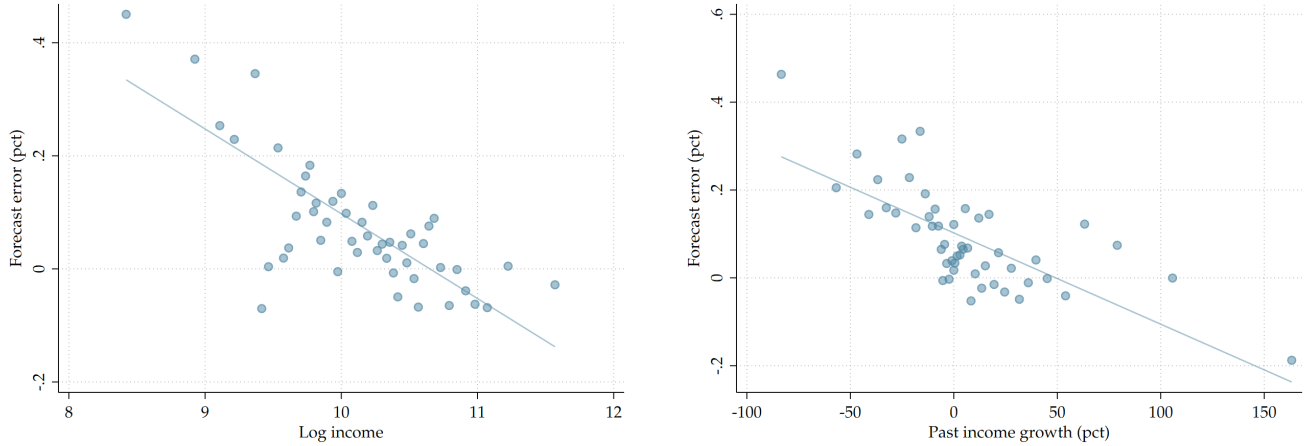
Control for Wealth Appendix C.2 below reproduces the last two panels of Figure 8 after residualizing both the x and y axis for the logarithm of household net wealth.

Figure A.3: Controlling for Wealth



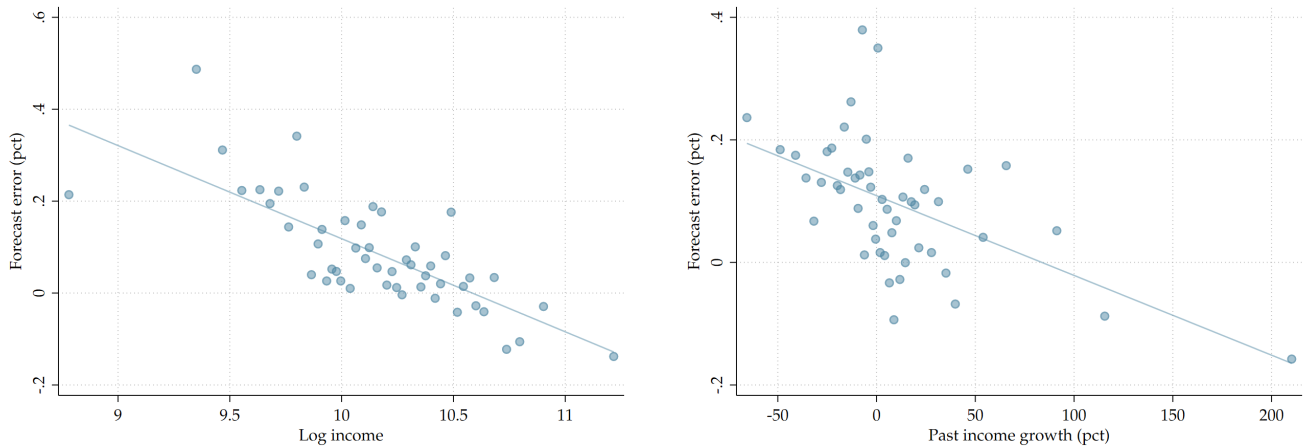
Unconditional Correlations: No Controls Appendix C.2 below reproduces the last two panels of Figure 8, without including any control before plotting the binned scatter plot.

Figure A.4: Unconditional Correlations



Trim Data at 99th Percentile Appendix C.2 below reproduces the last two panels of Figure 8 but trimming the forecast error and income change data at the 99th –rather than 98th–percentile.²⁶

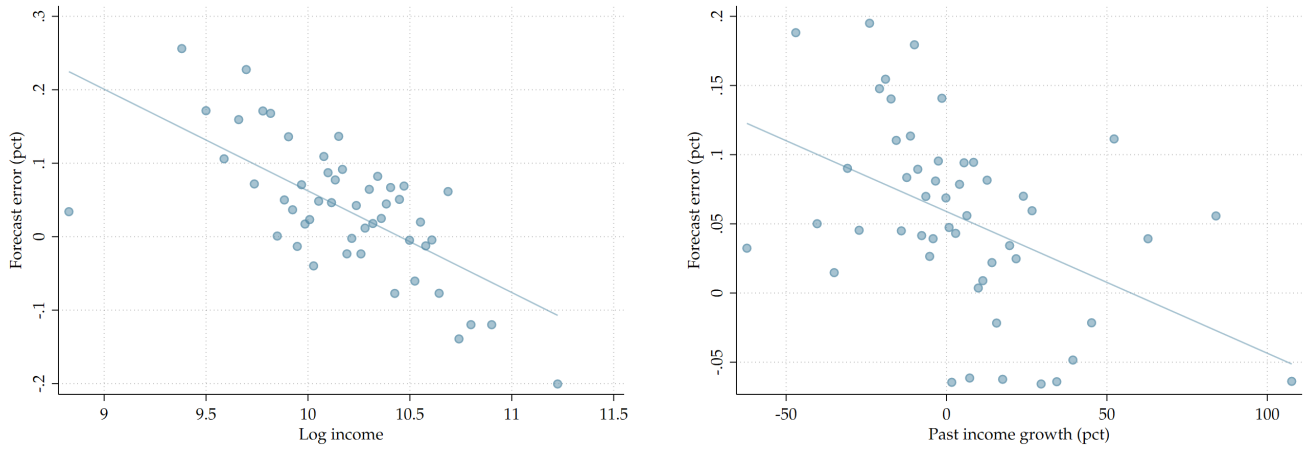
Figure A.5: Trimming Forecast Errors at the 99th Percentile



²⁶Note that for the income change variable, we perform two-sided trimming, by trimming observations based on their absolute value.

Trim Data at 95th Percentile Appendix C.2 below reproduces the last two panels of Figure 8 but trimming the forecast error and income change data at the 95th –rather than 98th–percentile.²⁷

Figure A.6: Trimming Forecast Errors at the 95th Percentile



²⁷Note that for the income change variable, we perform two-sided trimming, by trimming observations based on their absolute value.