

Non-College Occupations, Workplace Automation, and the College Gender Gap

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Women overtook men in college enrollment

Introduction

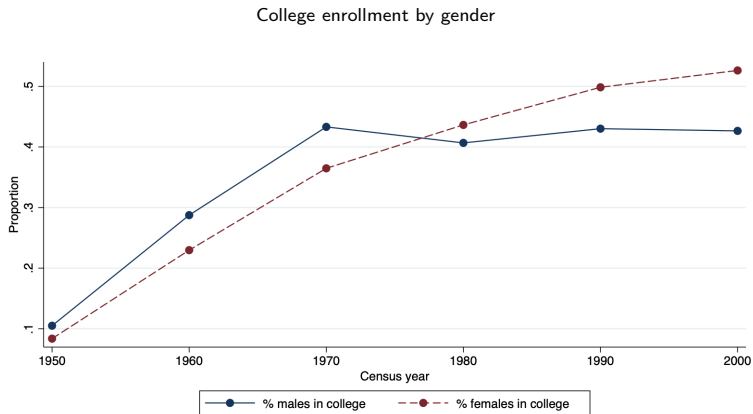
Data

Descriptive
Evidence on
Routinization

Two Stage Least
Squares Approach

Structural Model

Conclusion

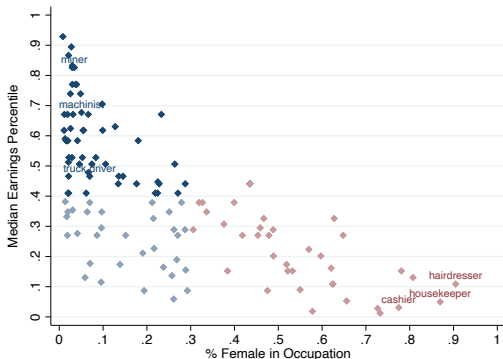


Source: IPUMS decadal census data for 1950-2000; American Community Survey yearly data for 2001-2010. 18-30 year olds.

- Open question: Why do women exceed men in college-going?
 - When men earn and work more than women
- Marriage market premium (Chiappori, Iyigun, Weiss, 2009; Chiappori, Salanie, Weiss, 2017; Low, 2019; Zhang, 2021; Ge, Isaac, Miller, 2022)
- Most of the literature focuses on gender differences in the ability to prepare for college (supply side)
 - Academic performance (Goldin, Katz, and Kuziemko, 2006; Becker, Hubbard, and Murphy, 2010); non-cognitive abilities (Bertrand and Pan, 2013)
- In contrast, this paper explores gender differences in non-college job prospects (demand side)
 - Among high school graduates, women face worse job options than men

“Missing quadrant” of high-paying female non-college occs

Non-College Occupations, 2000



- Traditionally “male” occupations can pay highly
- Occupations which employed women have significantly lower median earnings

Decline in women's non-college occupations over time

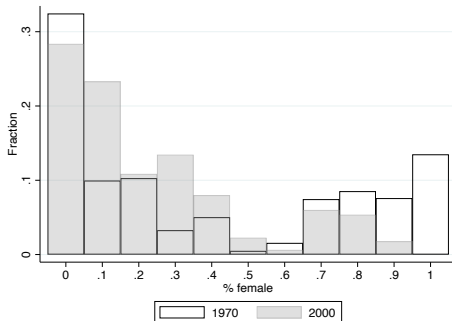
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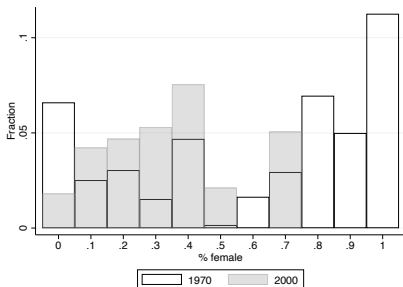


Source: IPUMS decadal census microdata (1970) and ACS data (2010). 18-30 year olds.

- In 1970, 70% of non-college working women held “routine” occupations
 - secretary, stenographer, clerical worker, telephone operator, typist
- 66% decline in share of secretarial jobs, 95% decline in share of typist jobs from 1970 to 2000

Decline in women's non-college occupations – driven by routine occs

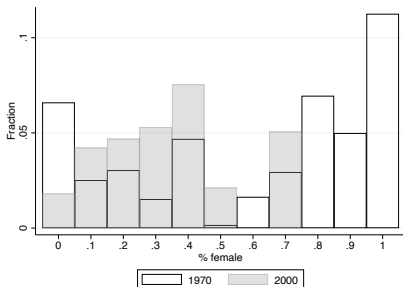
(a) Routine Occupations



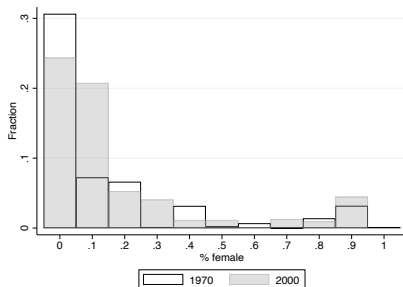
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- By 2000: no more high routine female-dominated occupations

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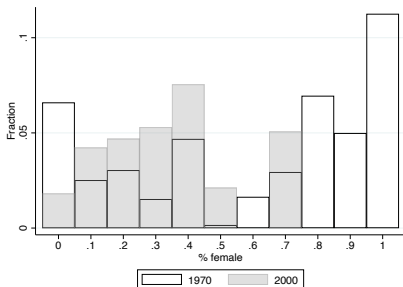
(b) Non-routine Occupations



- 1970: among high routine occupations, most were female-dominated
- By 2000: no more high routine female-dominated occupations
- Non-routine occupations: little change over time

Decline in women's non-college occupations – driven by routine occs

(a) Routine Occupations



- **Routinization: displacement of routine occupations by automation**

Research Question

- Did non-college jobs contribute to college gender gap?

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 - ① Cross-sectionally: Missing quadrant of high-paying occupations
 - ② Over time: non-college job prospects deteriorated

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- Challenge: occupation prospects are endogenous

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 - ② Over time: non-college job prospects deteriorated
- Challenge: occupation prospects are endogenous
- This paper uses automation as a shifter of non-college job opportunities
 - Premise: automation disproportionately displaced female non-college occupations
 - Approach: shift-share instruments to isolate labor demand for routine work

- 1 Instrument: Administrative share based on job posting data (1950-2000)
 - Labor markets with administratively-intense industries should undergo more routinization
 - Capture shifts in *labor demand* due to routinization

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 - Labor markets with administratively-intense industries should undergo more routinization
 - Capture shifts in *labor demand* due to routinization
 - Routinization \uparrow female enrollment, but not nec. male enrollment

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② Structural approach complements IV approach:

- Explicitly model mechanisms for individual decisions
- Quantify the changes in college enrollment due to routinization:
 - 4 pp (44%) growth in female enrollment, 1.6 pp growth in male enrollment

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Data and Measurement

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 - Sample size permits examining occupation/industry-level changes by commuting zone-year

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 - Task content of occupations: routine, manual, abstract
 - Measure of automation susceptibility: Routine Task Intensity (RTI)

$$RTI = \ln(\text{routine}) - \ln(\text{manual}) - \ln(\text{abstract})$$

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- Structural Model: National Longitudinal Survey of Youth (1979)

Descriptive Evidence on Routinization

Decline in labor share of high-RTI occupations for women

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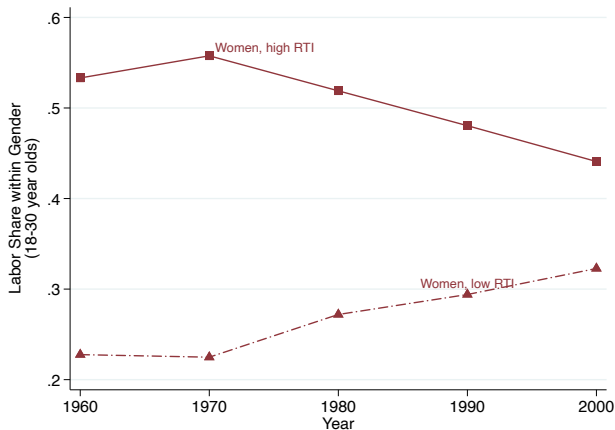
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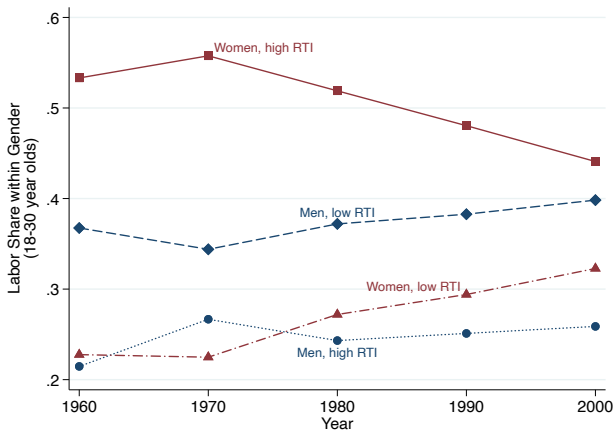
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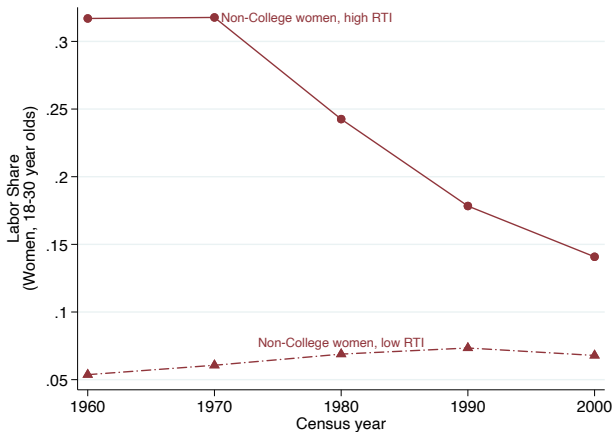
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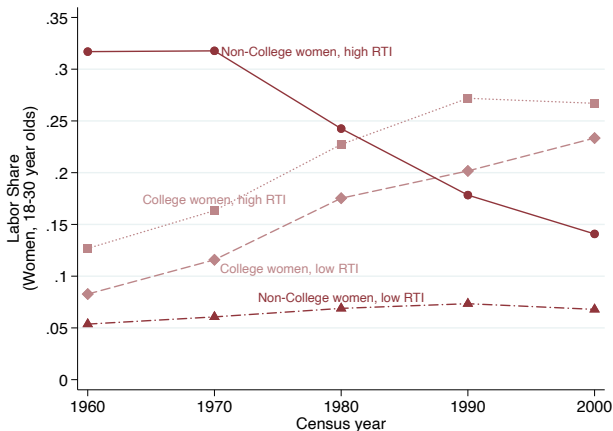
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Decline in RTI share affects college-going for women

Employment changes over time, by high and low RTI



Two Stage Least Squares Approach

Newspaper job postings from Atalay et al. (2020):

$$\text{admin share}_{ct} = \sum_{i=1}^I E_{i,c,1950} \frac{\sum_k L_{ikt} \mathbf{1}(\text{admin}_{kt} > \text{admin}_{1950}^{P66})}{\sum_k L_{ikt}} \quad (1)$$

- admin share_{ct} : predicted administrative share in commuting zone c
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- Intuition: commuting zones with high initial administrative shares experience more routinization over time
 - Wilmington, DE (legal industry)
 - Detroit, MI (manufacturing industry)

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- admin_{kt} : administrative activity in occupation k
- Identifying assumption: administrative share in national industry level can only affect college-going in commuting zone through routinization
 - Shift-share IV framework: industry-level shocks as good as random (Adao et al. 2019; Borusyak et al., 2018)
 - Standard error correction to account for correlated shocks across industries (Adao et al., 2019)

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- admin share_{ct} : predicted administrative share in commuting zone c
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- **Controls:**
 - male and female labor force participation (25-65 year olds)
 - predicted share of manual-intensive occupations
 - 10-year lagged services, manufacturing, retail, mining
 - race, gender, age, commuting zone, year, census controls
 - some specifications: median cognitive earnings, lagged routine share

First Stage Regression of Routinization on Instruments

	Routinization			
	(1)	(2)	(3)	(4)
Routinizability IV	0.387 (0.026) ^{***}	0.383 (0.027) ^{***}	0.388 (0.027) ^{***}	0.383 (0.027) ^{***}
F-statistic	214.572	204.993	204.654	201.452
Observations	3610	3610	3610	3610
Commuting zone FE	Yes	Yes	Yes	Yes
Year FE	Yes	Yes	Yes	Yes
Median cognitive earnings		Yes		Yes
Lagged RTI share			Yes	Yes

Notes: First stage regression of routinization on instruments. Standard errors are clustered at the two-digit industry level and adjusted using the correction procedure of AKM (2019). Oleva-Pflueger F-statistics reported using AKM (2019) standard errors. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

IV Regression: Routinization \Rightarrow growth in female enrollment

Second Stage Regressions

	College Enrollment			
	(1)	(2)	(3)	(4)
	<i>A: Second Stage Regression, Women</i>			
Routinization	0.578 (0.163) ^{***} [0.258,0.898]	0.606 (0.166) ^{***} [0.281,0.931]	0.578 (0.160) ^{***} [0.265,0.891]	0.606 (0.161) ^{***} [0.291,0.922]
F-statistic	214.572	204.993	204.654	201.452
Observations	3610	3610	3610	3610
	<i>B: Second Stage Regression, Men</i>			
Routinization	0.436 (0.236) [*] [-0.026,0.898]	0.444 (0.238) [*] [-0.022,0.910]	0.436 (0.232) [*] [-0.019,0.891]	0.444 (0.234) [*] [-0.015,0.904]
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Structural Model

Structural approach: mechanisms

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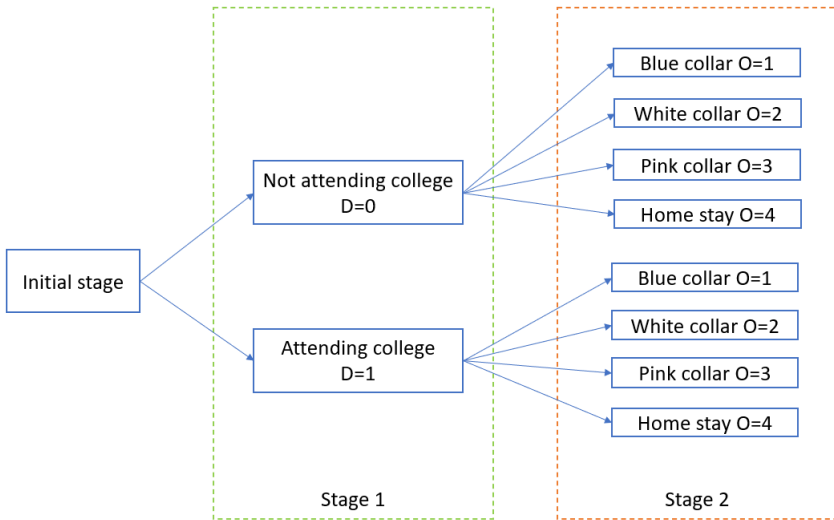
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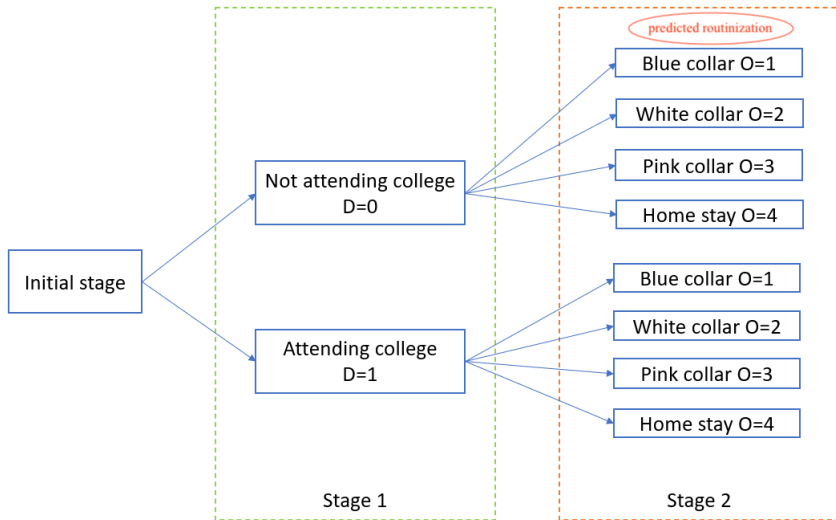
Conclusion

- Develop an explicit mechanism for how routinization affects individual decisions
 - Automation affects returns to occupations by influencing *skill prices*
 - Through impacting occupational returns, automation affects college enrollment decision
- Quantify the impact of routinization on college enrollment decisions
- Rationalize the polarization of non-college occupations
- Data: National Longitudinal Survey of Youth (1979)
 - Longitudinally link individual's education, occupation, and earnings over time
 - Skills: cognitive, mechanical, and administrative
 - Armed Services Vocational Aptitude Battery (ASVAB): cognitive, mechanical, administrative test scores (see Prada and Urzua, 2017)

2-period discrete choice model



2-period discrete choice model



Model Simulations: Quantifying impact of routinization

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	Women			Men		
	Baseline: 1980 (1)	Routinization in 2000 (2)	Change (3)	Baseline: 1980 (4)	Routinization in 2000 (5)	Change (6)
Occupation choices						
White collar	0.404	0.610	0.206	0.392	0.434	0.042
Blue collar	0.038	0.049	0.011	0.485	0.452	-0.032
Pink collar	0.348	0.153	-0.195	0.048	0.050	0.002
Not working	0.209	0.188	-0.021	0.075	0.064	-0.012
Education choices						
High school	0.389	0.350	-0.040	0.481	0.465	-0.016
College	0.611	0.650	0.040	0.519	0.535	0.016

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 - Census data: 9 pp rise
- Male enrollment rises by 1.6 pp
 - Census data: 2 pp rise

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- Gender polarization among non-college occupations \Rightarrow gaps in outside options to attending college

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- 1 IV approach:
 - greater routinization \Rightarrow displacement of women's non-college jobs
 - decline in outside options \Rightarrow \uparrow female enrollment
 - no systematic effects for men

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- ② Structural approach: Routinization \downarrow women's non-college occupational returns
 - Job polarization due to gender difference in skills & skill returns
 - Decline in pink-collar returns \Rightarrow women shift from pink-collar to white-collar \Rightarrow greater female college-going

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Thank you!

For questions or comments, please contact us at
achuan@msu.edu or wz301@cam.ac.uk

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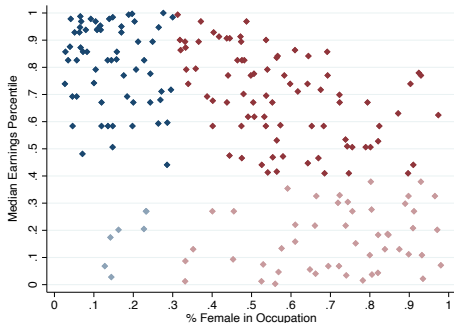
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Appendix

Opposite missing quadrant in college occs

Appendix

College Occupations, 2000

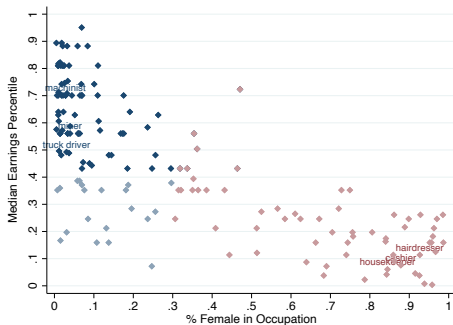


2000 Census. 18-30 year olds. College occupations are occupations with 50% of more workers who have ever enrolled in college.

◀ Non-College Occupations, 2000

- Missing “quadrant” of low-paying male college occupations
- Consistent with sorting based on outside options

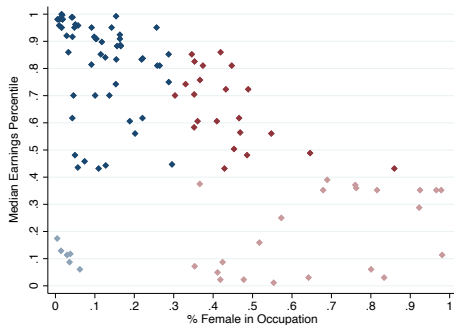
Non-College Occupations, 1970



Source: Census microdata, 1970. 18-30 year olds. Non-College occupations are occupations with 50% of more workers who have never enrolled in college.

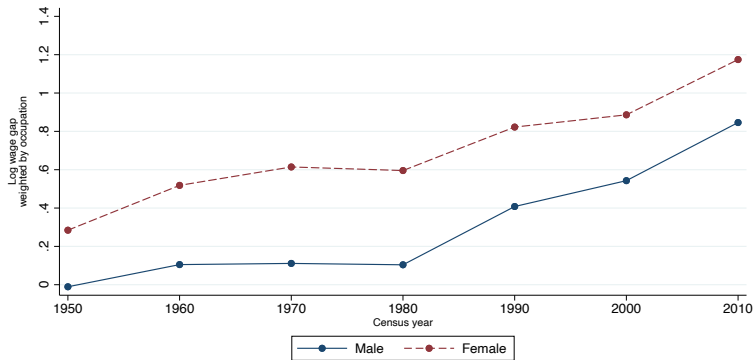
◀ Non-College Occupations, 2010

College Occupations, 1970



Source: Census microdata, 1970. 18-30 year olds. College occupations are occupations with 50% of more workers who have ever enrolled in college.

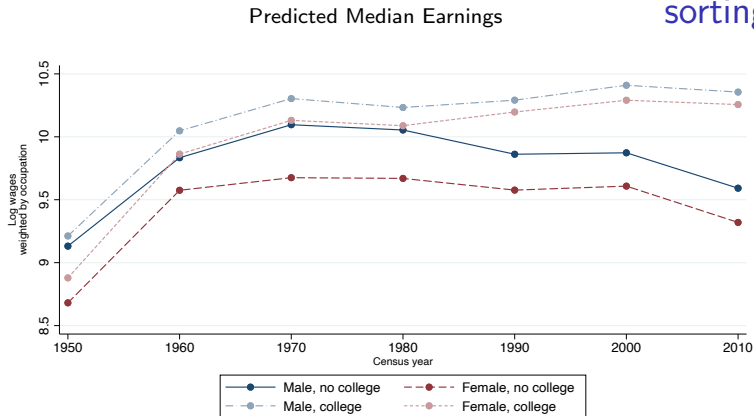
◀ Non-College Occupations, 2010

Wage gap due to occupational sorting
College-High School Median Wage Difference

Source: Census and American Community Survey. 18-30 year olds. Difference in predicted median earnings for college and non-college workers, where

$$\text{predicted median earnings} = \sum_o^O \text{median earnings}_o \times \text{proportion in occupation } o \text{ by gender}$$

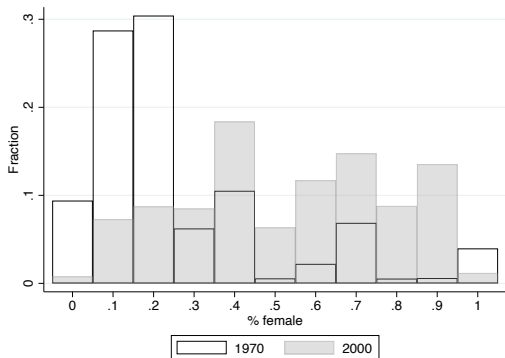
Predicted median earnings due to occupational sorting



Source: Census and American Community Survey. 18-30 year olds. Predicted median earnings for college and non-college workers computed as

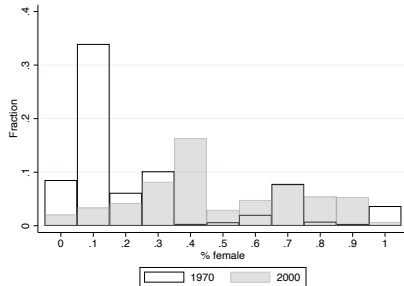
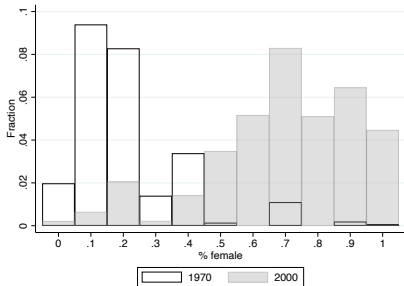
$$\text{predicted median earnings} = \sum_o^O \text{median earnings}_o \times \text{proportion in occupation } o \text{ by gender}$$

Increase in gender-equal college occupations over time



Source: IPUMS decadal census microdata (1970) and ACS data (2010). 18-30 year olds.

College Occupations, by RTI

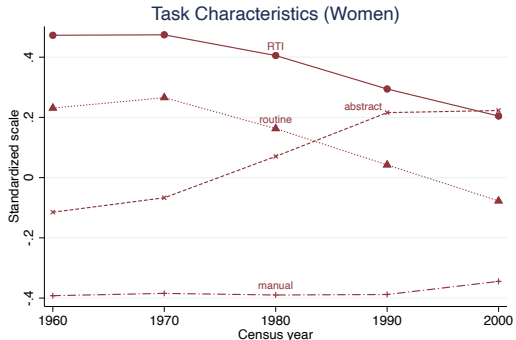


Source: IPUMS decadal census microdata (1970) and ACS data (2010). 18-30 year olds.

◀ Non-college occupations

Women's RTI over time

Routine Task Intensity (RTI) in Labor Market

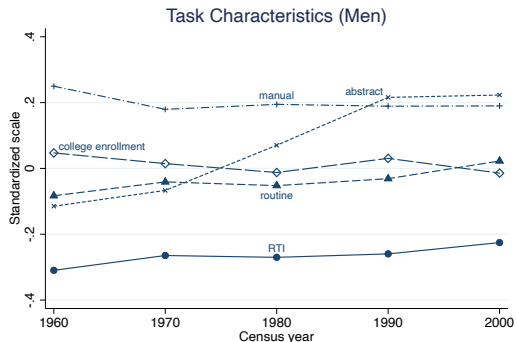


Source: IPUMS decadal census data for 1950-2000; American Community Survey yearly data for 2001-2005. 18-30 year olds.

$$RTI = \ln(\text{routine}) - \ln(\text{manual}) - \ln(\text{abstract})$$

Men's RTI over time

Routine Task Intensity (RTI) in Labor Market



Source: IPUMS decadal census data for 1950-2000; American Community Survey yearly data for 2001-2005. 18-30 year olds.

$$RTI = \ln(\text{routine}) - \ln(\text{manual}) - \ln(\text{abstract})$$

Regression equations

First stage regression:

$$\text{routinization}_{ct} = \alpha_0 + \alpha_1 \text{admin share}_{ct} + \alpha_2 X_{ct} + \phi_t + \theta_c + e_{ct} \quad (2)$$

- $\text{routinization}_{ct} = \text{RTI share}_{c,1950} - \text{RTI share}_{ct}$
- $\text{RTI share}_{c,1950}$: share of high RTI occs

Second Stage Regressions, Additional Specifications

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	<i>A: Female Enrollment</i>							
Routinization	0.495 (0.167)*** [0.167,0.822]	0.607 (0.158)*** [0.297,0.917]	0.628 (0.161)*** [0.312,0.944]	0.736 (0.242)*** [0.263,1.210]	0.504 (0.126)*** [0.256,0.752]	0.365 (0.145)** [0.080,0.650]	0.548 (0.145)*** [0.213,0.884]	0.784 (0.145)*** [0.350,1.219]
F-statistic	137.280	205.558	182.001	51.040	103.421	203.540	161.233	111.715
Observations	3610	3600	3610	3610	3610	3610	3610	3610
	<i>B: Male Enrollment</i>							
Routinization	0.315 (0.266) [-0.207,0.838]	0.503 (0.234)** [0.044,0.961]	0.441 (0.238)* [-0.025,0.907]	0.540 (0.315)* [-0.077,1.157]	0.218 (0.174) [-0.123,0.558]	0.254 (0.196) [-0.129,0.638]	0.432 (0.196) [-0.101,0.964]	0.616 (0.196)* [-0.048,1.280]
F-statistic	137.280	205.558	182.001	51.040	103.421	203.540	161.233	111.715
Observations	3610	3600	3610	3610	3610	3610	3610	3610
Minimum controls	✓							
Excluding Boston and NYC		✓						
Control for abstract occupation share			✓					
RTI share: non-college workers	✓	✓	✓		✓	✓	✓	✓
RTI share: college and non-college workers				✓				
IV: Administrative Share (top third)	✓	✓	✓	✓				
IV: Routine Share					✓			
IV: Administrative Share (top half)						✓		
IV: Administrative Activities							✓	
IV: Clerical Requirements								✓

Standard errors are clustered at the two-digit industry level and adjusted using the correction procedure of AKM (2019). Montiel Olea-Pflueger F-statistics reported using AKM (2019) standard errors. Anderson-Rubin (1949) confidence intervals reported using AKM (2019) correction. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

Second Stage Regressions, Additional Specifications

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Observations	3610	3600	3610	3610	3610	3610	3610	3610
Minimum controls	✓							
Excluding Boston and NYC		✓						
Control for abstract occupation share			✓					
RTI share: non-college workers	✓	✓	✓		✓	✓	✓	✓
RTI share: college and non-college workers				✓				
IV: Administrative Share (top third)	✓	✓	✓	✓				
IV: Routine Share					✓			
IV: Administrative Share (top half)						✓		
IV: Administrative Activities							✓	
IV: Clerical Requirements								✓

Standard errors are clustered at the two-digit industry level and adjusted using the correction procedure of AKM (2019). Montiel Olea-Pflueger F-statistics reported using AKM (2019) standard errors. Anderson-Rubin (1949) confidence intervals reported using AKM (2019) correction. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

Parametric assumptions (1/2)

Period 2: Occupation

$$U(O_i|D_i) = \underbrace{Y(O_i|D_i)}_{\text{earnings}} + \underbrace{P(O_i|D_i)}_{\text{nonpecuniary returns}}$$

Parametric assumptions (1/2)

Period 2: Occupation

$$U(O_i|D_i) = \underbrace{Y(O_i|D_i)}_{\text{earnings}} + \underbrace{P(O_i|D_i)}_{\text{nonpecuniary returns}}$$

where

- $\theta_i = \{\theta_{ci}, \theta_{si}, \theta_{mi}\}$ is a vector of skills
- D_i : whether have college degree
- O_i : occupations (white, blue, pink, home)
- X_i : background characteristics

Parametric assumptions (1/2)

Period 2: Occupation

$$U(O_i|D_i) = \underbrace{Y(O_i|D_i)}_{\text{earnings}} + \underbrace{P(O_i|D_i)}_{\text{nonpecuniary returns}}$$

where

$$Y(O_i|D_i) = X_i^Y \alpha_O^g + D_i \alpha_{OD}^g + \epsilon_{O,D,i}^Y +$$

$$\theta_i (\alpha_{O\theta}^{g,0} + \alpha_{O\theta}^{g,1} \widetilde{\text{routinization}}_{c(i),t}) + \theta_i D_i (\alpha_{OD\theta}^{g,0} + \alpha_{OD\theta}^{g,1} \widetilde{\text{routinization}}_{c(i),t})$$

- $\theta_i = \{\theta_{ci}, \theta_{si}, \theta_{mi}\}$ is a vector of skills
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$$P(O_i|D_i) = X_i^Y \beta_O^g + D_i \beta_{OD}^g + \epsilon_{O,D,i}^P + \theta_i (\beta_{O\theta}^{g,0} + \beta_{O\theta}^{g,1} \widetilde{\text{routinization}}_{c(i),t}) + \theta_i D_i (\beta_{OD\theta}^{g,0} + \beta_{OD\theta}^{g,1} \widetilde{\text{routinization}}_{c(i),t})$$

- $\theta_i = \{\theta_{ci}, \theta_{si}, \theta_{mi}\}$ is a vector of skills
- D_i : whether have college degree
- O_i : occupations (white, blue, pink, home)
- X_i : background characteristics

Parametric assumptions (2/2)

Period 1: Education

$$D_i = \mathbf{1}[V_i^1 + \xi_i^D > V_i^0]$$

$$V_i^0 = E \max_{O_i} U(O_i | D_i = 0)$$

$$V_i^1 = X_i^D \lambda_X^g + \theta_i \lambda_\theta^g + \rho E \max_{O_i} U(O_i | D_i = 1)$$

[◀ Model Diagram](#)

Occupational Returns

