Diversity and Discrimination in the Classroom

Dan Anderberg (Royal Holloway) Christina Felfe (U Konstanz) Gordon Dahl (UC San Diego) Helmut Rainer (LMU & ifo) Thomas Siedler (U Potsdam) Public education is one of the few social institutions that have the potential to bring children and youth together across ethnic, racial, and social lines

Could this prepare youth to succeed in an increasingly diverse society?

A priori unclear

- Polarization: In settings dominated by a few large groups, the drive to establish cultural domination could grow strong, causing own-group attachment to increase and integration to decrease
- Fractionalization: With many small groups, benefits could be gained by unifying under a shared identity, which in turn would ultimately foster social cohesion

Key Questions

The opposing forces of polarization and fractionalization have been explored in the context of nation and community-building in the developing world (Esteban & Ray 1994; Montalvo & Reynal-Querol 2005; Bazzi et al. 2019)

We apply the fractionalization-polarization paradigm to diverse schools and classrooms in Germany

Research questions:

- Does the type of diversity that prevails in schools matter for social cohesion?
- Is in-group bias more prevalent in polarized classrooms than in homogeneous or fractionalized classrooms?
- Does the cultural distance between majority and minority groups matter?
- What role do statistical and taste discrimination play, and what about stereotypes?

Challenges

Challenges:

Data: Measuring in-group bias among adolescents and studying polarization and fractionalization cannot be done with existing datasets

Selection bias: Interaction with ethnically diverse peers is endogeneous

Solutions:

Data: Conduct a large, incentivized lab-in-the field experiment in 220 classes spread across 57 German secondary schools

- Investment ("trust" game) to measure how native German students cooperate with in-group (other natives) versus out-group (immigrant) partners
- Survey to characterize the ethnic composition of a classroom

Identification: Exploit variation in peer group diversity arising from students' quasi-random assignment to classes within schools

We contribute to the literature on ...

- societal challenges related to ethnic diversity and the debate between polarization versus fractionalization:
 - at the macro-level (e.g., Alesina and La Ferrara, 2005; Montalvo and Reynal-Querol, 2005)
 - ▶ at the micro-level (e.g., Algan et al., 2016; Bazzi et al., 2019)
- immigration and xenophobia (e.g., Barone et al., 2016; Halla et al., 2017; Dustmann et al., 2018; Edo et al., 2019; Steinmayr, 2021)

ethnic peers in the educational context

- academic achievement (Ohinata and van Ours, 2013 & 2016)
- social cohesion (Boisjoly et al., 2006; Alan et al., 2021; Boucher et al., 2022; Corno et al., 2022;)

Data

Our data collection

- collaborated with educational authorities in 2 German states (NRW & SH),
- collected data for 57 schools (in 8 cities),
- ▶ targeted all 9th/ 10th graders (220 classes),

- 4,634 students out of which 4,094 provided complete information (1% of the parents and 3.5% of the students opted out, 7% did not complete the survey)

– $\mathbf{2,257}$ native students and $\mathbf{1,837}$ immigrant students (at least one parent born abroad)

combined

(i) a **survey** to determine students' ethnic background plus other variables

(ii) an incentivized experiment to elicit students' in-group out-group bias

The survey contained questions on parents' country of birth and student's religion allowing us to create measures for the ethnic composition of the classroom:

1. Intergroup diversity:

= Share of immigrant peers (at least one parent born abroad) per class (leave-one-out share: mean= 0.38; std. dev.= 0.21)
Distribution

2. Intra-group diversity: Distribution

- = Share and cultural distance of immigrant subgroups distinguishing between
 - Religious affiliation (individual-level data): Muslims: mean= 0.18; std dev.= 0.18 Non-muslims: mean= 0.20; std dev.= 0.11
 - ii. Linguistic distance (country-level data from the Ethnologue database): Distant: mean= 0.21; std dev.= 0.17 Close: mean= 0.17; std dev.= 0.11

Setting



Sample size: 220 classes

Sample size: 2,202 native German students

Setting



Investment Game (Berg et al, 1995)



Experimental Protocol

To elicit students' in-group bias, we conducted a **investment game** and asked students to fill out decision sheets being aware that ...

- they would earn money,
- decisions would be anonymous,
- they would first play as first-mover and then as second-mover,
- they would play with different game partners,
- partners were of their age and from the same state, but not the same school,
- partners would be randomly matched after all possible decisions were made,
- role of the sender or receiver randomly assigned,
- payoff based on their own decision and the decision of their partner,
- respective amount paid within two weeks in anonymized envelopes.
- Important note: By construction, statistical discrimination is constant with respect to class type

Decision Sheet (First Mover)

"You are the sender and you have 5 EURO. Which amount would you like to send to the receiver (max. 5 EURO each time)? Please check one box in each column 1-6."

-					
The receiver is					
COLUMN 1	COLUMN 2	COLUMN 3	COLUMN 4	COLUMN 5	COLUMN 6
a boy with German parents	a girl with German parents	a boy with foreign parents	a girl with foreign parents	a boy with foreign parents who possesses German citi- zenship	a girl with foreign parents who possesses German citi- zenship
D 0 EURO	D 0 EURO	D 0 EURO	D 0 EURO	$\square 0 EURO$	D 0 EURO
\Box 0.5 EURO	□ 0.5 EURO	\Box 0.5 EURO	\square 0.5 EURO	\Box 0.5 EURO	\Box 0.5 EURO
□ 1 EURO	□ 1 EURO	□ 1 EURO	\Box 1 EURO	□ 1 EURO	□ 1 EURO
\Box 1.5 EURO	□ 1.5 EURO	\Box 1.5 EURO	\Box 1.5 EURO	□ 1.5 EURO	\Box 1.5 EURO
\Box 2 EURO	□ 2 EURO	\Box 2 EURO	\Box 2 EURO	\square 2 EURO	□ 2 EURO
□ 2.5 EURO	□ 2.5 EURO	\Box 2.5 EURO	\square 2.5 EURO	□ 2.5 EURO	\Box 2.5 EURO
□ 3 EURO	□ 3 EURO	🗆 3 EURO	□ 3 EURO	🗆 3 EURO	🗆 3 EURO
□ 3.5 EURO	□ 3.5 EURO	□ 3.5 EURO	\square 3.5 EURO	□ 3.5 EURO	□ 3.5 EURO
□ 4 EURO	□ 4 EURO	□ 4 EURO	□ 4 EURO	□ 4 EURO	□ 4 EURO

 $12 \, / \, 52$

We use the first four choices $\{S_1, ..., S_4\}$ to construct our main outcome variable, the **in-group out-group investment gap (IG)** of native students:

$$IG = S_N - S_I$$

where

•
$$S_N = \frac{1}{2}(S_1 + S_2) = \text{avg.}$$
 investment to natives
• $S_I = \frac{1}{2}(S_3 + S_4) = \text{avg.}$ investment to immigrants

Out of the 2,257 native students 2,202 completed the experiment and exhibited an in-group out-group investment gap of 0.08 Euro on average (std. dev. 0.76)

We are interested in the impact of the ethnic composition of native students' peer group on their investment gap (IG_i) :

$$\textbf{IG}_{i,c,s} = \alpha + \beta f(\textbf{Ethnic Diversity}_{c,s}) + \gamma X_i + \delta Z_{c,s} + \theta_s + \varepsilon_{i,c,s}$$

We tackle endogenous peer group formation and control for common shocks by

- conducting a within school analysis (using school fixed effects θ_s),
- relying on the quasi-random assignment of immigrants to classrooms.

Identifying Assumption

1. Sufficient within school variation in the ethnic composition of classes

Within School Variation



Identifying Assumption

- 1. Sufficient within school variation in the ethnic composition of classes
- 2. Balancing on observables (within school)

Balancing on Observables • Back

Dependent Variable:	able: Percent Immigrants in Class	
	(1)	(2)
Male	-0.004	-0.004
	(0.010)	(0.005)
Age	0.055***	-0.006
	(0.012)	(0.005)
Age missing	-0.035	-0.031
	(0.033)	(0.019)
Protestant	-0.086***	0.006
	(0.018)	(0.008)
Muslim	0.098**	0.000
	(0.039)	(0.019)
Other religion	-0.103***	-0.002
	(0.020)	(0.009)
SES: Two-parent hh; low education	0.035***	-0.006
	(0.013)	(0.007)
SES: Single-parent hh; high education	0.017	-0.001
	(0.017)	(0.009)
SES: Single-parent hh; low education	0.053***	-0.003
	(0.014)	(0.006)
SES: missing	0.075***	0.002
	(0.015)	(0.007)
Age mother	-0.002*	0.001
	(0.001)	(0.001)
Age mother: missing	0.069	-0.012
	(0.042)	(0.022)
Age father	-0.001	-0.000
	(0.001)	(0.000)
Age father: missing	0.058	0.023
	(0.042)	(0.019)
Class size	-0.004	0.000
	(0.003)	(0.003)
Observations	2,202	2,202
R-squared	0.114	0.736
F-statistic	7.06	1.21
p-value	0.000	0.263
School FE		.(

Identifying Assumption

- 1. Sufficient within school variation in the ethnic composition of classes
- 2. Balancing on observables (within school)
 - \rightarrow 0/15 observables significantly predict the share of immigrant peers
 - \rightarrow Observables do not jointly predict the share of immigrant peers
 - \rightarrow Similar balancing for muslim share
- Random assignment of immigrants to classrooms (within school)
 → Class fixed effects do not significantly predict immigrant background
 conditional on school fixed effects (p=0.266)

 \rightarrow Similar result for muslim background (p=0.471)

Identifying Assumption

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- 3. Random assignment of immigrants to classrooms (within school) \rightarrow Class fixed effects do not significantly predict migrant background conditional on school fixed effects (p=0.266)
 - \rightarrow Similar result for muslim background (p=0.471)
- 4. Simulation of random assignment of immigrants to classrooms (within school) \rightarrow Actual distribution gualitatively similar to the simulated one





Impact of the share of immigrants peers in the classroom:

$$\mathsf{IG}_{\mathfrak{i},\mathfrak{c},\mathfrak{s}} = \alpha + \beta f(\mathsf{Diversity}_{\mathfrak{c},\mathfrak{s}}) + \gamma \mathsf{X}_{\mathfrak{i}} + \delta \mathsf{Z}_{\mathfrak{c},\mathfrak{s}} + \theta_{\mathfrak{s}} + \varepsilon_{\mathfrak{i},\mathfrak{c},\mathfrak{s}}$$

where we use a 2^{nd} order polynomial of the immigrant share π_I to model the function of ethnic diversity:

$$\beta f(\text{Diversity}_{c,s}) = \beta_1 \pi_I + \beta_2 \pi_I^2$$

In-Group Favoritism: An Inverted-U



Empirical Specification II

Building upon the measures of Polarization and Fractionalization ...

$$\label{eq:Polarization} \text{Polarization} = 4 \sum_{i=1}^N \pi_{Ii}^2 (1-\pi_{Ii}) \quad \text{ and } \quad \text{Fractionalization} = 2 \sum_{i=1}^N \pi_{Ii} (1-\pi_{Ii})$$

... and starting with three subgroups – two immigrant subgroups (π_{I1} , π_{I2}) and natives ($1 - \pi_{I1} - \pi_{I2}$) – ethnic diversity can be expressed as:

$$\mathcal{F}(\pi_{I1},\pi_{I2}) = \underbrace{4(\pi_{I1} + \pi_{I2} - \pi_{I1}^2 - \pi_{I2}^2 - \pi_{I1}\pi_{I2}) - 12(\pi_{I1}\pi_{I2} - \pi_{I1}^2\pi_{I2} - \pi_{I1}\pi_{I2}^2)}_{= POLAR}$$

Empirically, a model that nests both polarization and fractionalization is:

$$\begin{split} IG_{i,c,s} = & \alpha + \beta_1 \pi_{I1,c,s} + \beta_2 \pi_{I2,c,s} + \beta_3 \pi_{I1,c,s}^2 + \beta_4 \pi_{I2,c,s}^2 + \beta_5 \pi_{I1,c,s} \times \pi_{I2,c,s} + \\ & \beta_6 \pi_{I1,c,s}^2 \times \pi_{I2,c,s} + \beta_7 \pi_{I2,c,s}^2 \times \pi_{I1,c,s} + \gamma X_i + \delta Z_{i,c,s} + \sigma_s + \varepsilon_{i,c,s} \end{split}$$

In-Group Favoritism: How Type of Diversity Matters



Payoff Losses Due to In-Group Favoritism



 $^{(1 \}text{ std.dev.} = 36 \text{ Euro cents})$

Mechanism I: Taste Discrimination

To examine taste discrimination, we look at the choices of natives when playing as "receivers" $% \left({{{\mathbf{r}}_{\mathrm{s}}}^{2}}\right) = {{\mathbf{r}}_{\mathrm{s}}}^{2}$

These choices capture other-regarding preferences towards natives vs. immigrants

- Receivers choose how much to transfer back to senders, as a function of how much they received from them
- Since we implemented a strategy vector method, receivers act as dictators, with no strategic considerations (e.g., no reciprocity)

Mechanism I: Taste Discrimination



 $^{(1 \}text{ std.dev.} = 1.05 \text{ Euro cents})$

Mechanism II: Statistical Discrimination

Recall that individuals know they are playing against a randomly selected individual from another school, and not someone from their own classroom

By construction, statistical discrimination is constant with respect to class type

 \Rightarrow As such, we can rule out statistical discrimination as a mechanism

An important rationale for sending money is "trust", i.e., the sender's beliefs that the receiver will reciprocate generously

Natives may lack trust in immigrants if they are treated worse by immigrants in their classroom relative to how they are treated by natives

If natives (mistakenly) extrapolate their experience within their classroom, they would trust **all** immigrants less than natives, and rationally invest less in immigrants vs. natives given these biased beliefs

To examine the scope for stereotypes as a mechanism, we first look at the predicted gap in how much natives transfer back to natives vs. how much immigrants transfer back to natives

- ▶ a reasonable proxy for differential treatment by immigrants versus natives
- feasible due to the strategy vector method

Mechanism III: Stereotypes - Role of Peers' Behavior



(1 std.dev. = 1.3 Euro cents)

Mechanism III: Stereotypes - Trust

We just showed evidence suggesting that natives are treated worse by immigrants than natives in polarized classrooms

If they extrapolate this experience to the broader population of immigrants, we would expect them to express less trust in immigrants relative to natives

To test this, we use survey questions, recorded on a scale from 1 to 10, asking how much do you trust a native German and how much do you trust an immigrant

Mechanisms I: Stereotypes - Trust



Conclusions

Contrary to what one naively may expect, social cohesion does not necessarily increase when bringing people from different background together

- Natives' in-group bias has an inverse U-shape in the share of out-group members (immigrants),
- Digging deeper, in-group bias peaks in polarized classrooms where only natives and culturally distant immigrants are present and are roughly equally represented
- At the peak, in-group bias comes with substantial payoff losses
- Mechanisms:
 - No statistical discrimination due to our design
 - Minor role of taste discrimination
 - Stereotypes due to differential classroom experiences as the main driver

Thank you!

Appendix

Share of Immigrant Peers



Sample size: 220 classes

Sample size: 2,202 native German students

Back / Identification

38 / 52

Share of Muslim and Non-Muslim Immigrant Peers



Share of Linguistically Distant and Close Immigrant Peers



In-group Favoritism: Share of immigrant peers

Dependent Variable:	In-group/out-group			
	investment gap (std.)			
	(1)	(2)	(3)	
		1 600444		
% of immigrant peers	1.568***	1.602***	1.605***	
	(0.398)	(0.399)	(0.398)	
% of immigrant peers squared	-1.757***	-1.809***	-1.797***	
	(0.468)	(0.470)	(0.467)	
p-value: Both coeff. are jointly equal to zero	0.0005	0.0004	0.0004	
p-value: Sum of both coeff. is equal to zero	0.280	0.241	0.267	
Observations	2,202	2,202	2,202	
R-squared	0.054	0.063	0.068	
Basic controls	\checkmark	\checkmark	\checkmark	
Religious background		\checkmark	\checkmark	
Family background			\checkmark	

Additional Specification - First Order Polynomial

Dependent Variable:	In-Group/Out-Group Investment Gap		Group Gap
	(1)	(2)	(3)
% of immigrant peers	0.080 (0.294)	0.052 (0.295)	0.034 (0.298)
Observations R-squared	2,145 0.047	2,145 0.053	2,145 0.061
Basic controls Religious background Family background	\checkmark	\checkmark	\checkmark \checkmark \checkmark

Back

Additional Specification - Third Order Polynomial

Dependent Variable:	In-group/out-group investment gap (std.)		
	(1)	(2)	(3)
% of immigrant peers	2.187***	2.124**	2.045**
	(0.805)	(0.820)	(0.809)
% of immigrant peers squared	-3.482*	-3.264	-3.024
	(2.076)	(2.113)	(2.092)
% of immigrant peers cubed	1.292	1.089	0.919
	(1.571)	(1.589)	(1.578)
Observations	2.202	2.202	2.202
R-squared	0.054	0.063	0.068
Basic controls	\checkmark	\checkmark	\checkmark
Religious background		\checkmark	\checkmark
Family background			\checkmark

In-Group Favoritism: An Inverted-U (Cubic)



Back

In-Group Bias: How Polarization Matters •••••

Dependent Variable:	In-Group/Out-Group Investment Gap		
	(1)	(2)	(3)
% of muslim immigrant peers $(\pi_{\mathfrak{m}})$	3.012***	3.007***	2.902***
	(0.804)	(0.798)	(0.809)
$\%$ of non-muslim immigrant peers $(\pi_{ ext{n}})$	1.490***	1.486***	1.438***
_	(0.527)	(0.531)	(0.533)
$\pi_{\rm m}^2$	-3.818***	-3.790***	-3.596***
	(1.012)	(1.008)	(1.024)
π_n^2	-1.544	-1.587	-1.534
	(0.974)	(0.983)	(0.984)
$\pi_m \times \pi_n$	-12.069**	-11.923**	-11.215**
	(5.023)	(5.023)	(5.122)
$\pi_m^2 \times \pi_n$	9.493*	9.004*	8.037
	(4.994)	(5.031)	(5.101)
$\pi_{\rm m} \times \pi_{\rm n}^2$	9.412	9.594	9.172
	(6.481)	(6.476)	(6.566)
Observations	2,163	2,163	2,163
R-squared	0.058	0.059	0.064
Basic controls	\checkmark	\checkmark	\checkmark
Religious background		\checkmark	\checkmark
Family background			\checkmark

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Robustness: Linguistically Close and Distant Immigrants



Model which Can Rationalize Our Findings

We write down a model of social group formation for two types of individuals with different innate cultural identities. We combine the ideas of

 Brock and Durlauf's (2001) "discrete choice with social interactions model" for social conformity preferences,

▶ Hotelling's (1929) "linear city" model with cultural identity distances.

Consistent Model

Setting:

- Individuals of type $A \equiv (N; I)$,
- ▶ Share of type I equals $\Pi \in (0, 1)$ and of type N equals (1Π) ,
- \blacktriangleright Cultural identity θ of type N equals 0 and of type I equals 1 .

Individuals choose

- \blacktriangleright with probability μ_{α} to join the mixed group and adopt cultural identity $\theta \in (0,1),$
- \blacktriangleright with probability $(1-\mu_{\alpha})$ to stay with their own type group and retain their own culture.

▶ Individuals have preferences for group size, but against cultural distance

▶ Details

Consistent Model



Symmetric benchmark case:

- Homogenous preferences for group size and against cultural distance
- If share of immigrants $\Pi = 0.5$, then cultural identity of the mixed group $\theta = 0.5$

Consistent Model



Key to the model are two opposing forces

- direct effect of group size holding cultural identity of the mixed group constant
- indirect effect of a changing cultural identity as the share of immigrants changes

Model Details

Utility of Natives if in

- $\blacktriangleright \mbox{ mixed group: } \mathfrak{u}^{\theta}_i = \beta_N[(1-\Pi)\mu_N + \Pi\mu_I] h_N(\theta) + \varepsilon^{\theta}_i$
- \blacktriangleright own type group: $\mathfrak{u}_{\mathfrak{i}}^{0}=\beta_{N}[(1-\Pi)(1-\mu_{N})]+\varepsilon_{\mathfrak{i}}^{0}$

Utility of Immigrants if in

- $\blacktriangleright \mbox{ mixed group: } \mathfrak{u}^{\theta}_i = \beta_I[(1-\Pi)\mu_N + \Pi\mu_I] h_1(1-\theta) + \varepsilon^{\theta}_i$
- \blacktriangleright own type group: $u_i^1 = \beta_I [\Pi(1-\mu_I)] + \varepsilon_i^1$

where

 $\boldsymbol{\beta}$ is the utility gain from group size,

h() is the utility loss from cultural distance,

 ε_i^{θ} is an i.i.d. extreme value distributed utility shocks.

Back

Model Assumptions

Assumption 1:

The social preference parameters satisfy $\beta_{\alpha}\in(0,2)$ for $\alpha\in(N,I).$ The cost of cultural distance $h_{\alpha}()$ is twice continuously differentiable and satisfies $h_{\alpha}(0)=0,$ $h_{\alpha}'(\theta)>0,$ $h_{0}''(\theta)>0$ with $h_{0}''(\theta)>(h_{\alpha}'(\theta))^{2}>0$ for any $\theta\in[0,1]$ and for $\alpha\in A.$

Assumption 2: Cultural identity $\theta \in [0,1]$ of the mixed group maximizes the size of that group given Π

▶ Back