Gender of the first child and labour market outcomes in the UK

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Introduction

Recent evidence on child penalties in earnings



(a) Scandinavian countries (b)

(b) German-speaking countries

Child Penalties across Countries: Evidence and Explanations, by Henrik Kleven, Camille Landais, Johanna Posch, Andreas Steinhauer, Josef Zweimüller, in AER-P&P, 2019

Recent evidence on child penalties in earnings: United Kingdom



The Importance of Child Penalties

- Childbirth has large and persistent effects on the labour market outcomes of women, but not on those of men (Kleven et al, 2009-2011)
- On average, men and women share the costs of parenthood unequally
- Mothers spend more time on non-market activities:
 - Child care (Guryan et al, 2008)
 - Other home production (Borra et al, 2021)
- Why is it important for the UK?
 - 48% of 18-24 old Brits say they definitely do want to have children one day (ONS, 2021)
 - Yet emotional & physical child care seldom shared evenly between men and women
 - UK gender pay gap progress slowed down as motherhood penalty persists (ONS, 2021)

- We analyze whether the child penalty varies with the gender of the firstborn child
- This can illuminate the role of parental preferences in shaping the child penalty
- Revisit Kleven et al. event study research design for treatment effect estimation

How does a firstborn child's gender affect parental behaviour, and does child gender feed into parental gender gaps?

- Exploit high-quality panel structure of the UK Household Longitudinal Study (UKHLS)
- Focus on the recent time period 2009-2019

- Significant motherhood penalty across countries
 - Kleven et al, 2018, Artmann et al, 2021, Berniell et al, 2022
- The most talented women leave the labor market/uptake part-time jobs
- Motherhood serves as an information shock to women's beliefs
 - Fernandez-Kranz et al, 2013, Kuziemko et al, 2018, Berniell et al, 2021
- Preferences, gender norms, and discrimination are key
- High quality publicly provided child care is more effective than paternity leave
 - Andersen & Nix, 2020, Andersen & Nix, 2021

Main findings

- Labour market penalties for mothers of daughters relative to the fathers of daughters are twice as large as penalties for mothers of sons relative to the fathers of sons
 - Probability to be employed: firstborn son \rightarrow 7% vs daughter \rightarrow 21%
 - Real wage: firstborn son 10% vs daughter 27%
- Mothers of daughters are more involved with their kids
- Fathers of daughters are **less** involved in shared childcare responsibilities in the household
- Fathers of daughters are less likely to get married after the birth of a child
- Mothers of daughters tend to become more conservative, while fathers of daughters - more progressive
- Mothers of daughters spend 50% more time on household chores

UK: Data and Descriptives

Data

- UK Household Longitudinal Study 2009-2019
- Sample \rightarrow individuals aged 16-64 who have never had a biological child before the sample period begins
- Child gender is random conditional on the decision of having a child \rightarrow firstborn
- Strongly balanced panel: 7048 parent x year excluding ethnic minorities
- We focus on labour market outcomes:
 - Employment
 - Total Working hours
 - Real wage
- Explore the following mechanisms:
 - Marriage and fertility
 - Parental involvement
 - Time spent on household chores and childcare responsibility
 - Progressiveness

Our sample is predominantly urban, UK born

| | Fathers | Mothers | Total |
|--------------|----------|----------|----------|
| Arro | 31.544 | 30.422 | 31.008 |
| Age | (13.016) | (13.004) | (13.022) |
| Married | 0.1881 | 0.209 | 0.198 |
| Warneu | (0.3907) | (0.406) | (0.399) |
| Weekly Hours | 23.40 | 22.311 | 22.868 |
| | (21.66) | (20.061) | (20.904) |
| Employed | 0.6427 | 0.651 | 0.647 |
| | (0.4792) | (0.476) | (0.478) |
| Religious | 0.445 | 0.492 | 0.467 |
| | (0.497) | (0.499) | (0.499) |
| L bele e e | 0.834 | 0.831 | 0.833 |
| Orban | (0.372) | (0.375) | (0.373) |
| Foreign Born | 0.209 | 0.184 | 0.197 |
| | (0.407) | (0.387) | (0.398) |
| Dermer | 0.281 | 0.313 | 0.296 |
| Degree | (0.449) | (0.464) | (0.457) |
| Observations | 3,069 | 4,411 | 7,480 |

Notes. Fathers(mothers) are men(women) who became a parent during the observed period of 10 years.

90% of parents are living together with their partner.

| | Fathers | | | Mothers | | |
|-----------------|----------|----------|----------------------|----------|----------|----------------------|
| | Daughter | Son | p-value (H0: Diff=0) | Daughter | Son | p-value (H0: Diff=0) |
| Age | 29.42 | 29.44 | 0.959 | 27.04 | 27.01 | 0.938 |
| Married | 0.72 | 0.69 | 0.503 | 0.60 | 0.64 | 0.380 |
| Weekly Hours | 38.05 | 36.20 | 0.349 | 33.42 | 34.69 | 0.413 |
| Employed | 0.90 | 0.88 | 0.459 | 0.90 | 0.92 | 0.373 |
| Religious | 0.80 | 0.73 | 0.110 | 0.80 | 0.76 | 0.338 |
| Urban | 0.83 | 0.85 | 0.491 | 0.82 | 0.78 | 0.357 |
| UK Born | 0.83 | 0.87 | 0.241 | 0.80 | 0.85 | 0.110 |
| Ethnic Minority | 0.04 | 0.03 | 0.393 | 0.06 | 0.08 | 0.706 |
| Degree | 0.47 | 0.55 | 0.127 | 0.51 | 0.58 | 0.081* |
| Real Wages | 2,621.79 | 2,717.78 | 0.603 | 2,008.51 | 1,995.80 | 0.917 |
| Parent × Year | 1490 | 1580 | 3070 | 2350 | 2060 | 4410 |

Notes. Characteristics are considered in the first wave of the dataset.

Empirical Strategy

Empirical strategy

• First we estimate Kleven et al event study style regression of:

$$Y_{ist}^{g} = \sum_{j \neq 1} \alpha_{j}^{g} \cdot \mathbb{1}[j = t] + \sum_{k} \beta_{k}^{g} \cdot \mathbb{1}[k = age_{is}] + \sum_{j} \gamma_{y}^{g} \cdot \mathbb{1}[y = s] + \delta_{is}^{g} \cdot \mathbb{1}[Degree_{is}] + \varepsilon_{ist}^{g}$$

- for individual *i* in year *s* and event time (time to birth) *t*, stratified by gender *g*.
 Coefficients α capture time to birth effects, β capture age fixed effects, and γ
 capture year fixed effects. All coefficients are separately identified given variation in age at birth in each year. Also control for education before birth.
- Individual *i* has his/her first child at time t = 0
- Balanced sample of parents observed every year relative to t = -1, it runs from t = -3 to t = +5
- Set of age dummies \rightarrow underlying life cycle trends
- Set of year dummies \rightarrow time trends like wage inflation and business cycle

Empirical strategy

 Second we calculate approximate proportional changes in parental outcomes, e.g. real wage, owing to child birth as:

$$P_t^g \equiv rac{\widehat{lpha}_t^g}{E[ilde{Y}_{ist}^g|t]}$$

where

$$ilde{Y}^{g}_{ist} = \sum_{k} \widehat{eta}^{g}_{k} \cdot \mathbbm{1}[k = \mathsf{age}_{is}] + \sum_{y} \widehat{\gamma}^{g}_{y} \cdot \mathbbm{1}[y = s].$$

- \tilde{Y}_{ist}^{g} is the predicted salary for a woman of age k and calendar year y, omitting the contribution of time to birth dummies, and hence P_{t}^{g} is the event year t effect of children as a percentage of the counterfactual outcome absent children. The counterfactual is estimated off all other mothers from the data who at one point were aged k, and whose salary is also observed in calendar year y.
- We also estimate child penalty for parents of boys and parents of girls separately

• **Third** we wish to examine whether labour market responses to girls and boys are different. Estimate:

$$Y_{ist}^{g} = \sum_{j \neq 1} \alpha_{j}^{g} \cdot \mathbb{1}[j=t] \cdot \mathbb{1}[\text{Girl}_{is}^{g}] + \sum_{k} \beta_{k}^{g} \cdot \mathbb{1}[k=\text{age}_{is}] \cdot \mathbb{1}[\text{Girl}_{is}^{g}] +$$

$$+\sum_{j}\gamma_{y}^{g} \cdot \mathbb{1}[y=s] \cdot \mathbb{1}[Girl_{is}^{g}] + \delta_{is}^{g} \cdot \mathbb{1}[Degree_{is}] \cdot \mathbb{1}[Girl_{is}^{g}] + \varepsilon_{ist}^{g}$$

Results

Child penalty

Employment



• Average Penalty: Employment \rightarrow 14.5%

Employment by child gender



(a) Boy

(b) Girl

• Average Penalty: Boy vs Girl \rightarrow 7.4% vs 20.8%

Employment by child being a girl: full interaction



Mothers of daughters face lower probability to be employed than if they'd had a son

Total working hours



• Average Penalty: Total working hours ightarrow 34.9%

Total working hours by child gender



(a) Boy

(b) Girl

• Average Penalty: Boy vs Girl \rightarrow 33% vs 36%

Total working hours by child being a girl: full interaction



Mothers of daughters work significantly less from 2 years post-birth and beyond

Real wage



• Average Penalty: Real wage ightarrow 17%

Real wage by child gender



(a) Boy

(b) Girl

- Average Penalty: Boy vs Girl ightarrow 10% vs 27%

Real wage by child being a girl: full interaction



• The wage penalty is larger for mothers of daughters than mothers of sons

Mechanisms

- Families with first-born daughters have more children
- Fathers of boys are less likely to get divorced
- Fathers of boys are convicted of significantly fewer crimes
- First-born daughter instead of son increases fathers' extraversion
- Having daughters makes people more likely to vote for left-wing political parties (Dahl & Moretti, 2007, Oswald & Powdthavee, 2008, Dustmann & Landerso, 2018, van Lent, 2022)

Mechanisms

- Parental involvement indexparent
- Fertility fertility
- Progressiveness
- Gender attitudes Indexgen
- Time spent on household chores timechores
- Care responsibility heare
- Marriage married

Parental involvement: higher index = more involved in different activities with a kid

Parental involvement by child being a girl: full interaction



 Mothers of daughters are more involved in parenting than mothers of sons (40%), while fathers are 20% less involved mechanisms Time spent on household chores

Time chores by child being a girl: full interaction



- Women increase time spent on household chores when they have a daughter
- There is around 50% increase during the first 2 years and almost 3 times 5 years later

Childcare Responsibilities: higher index = higher self responsibility
Childcare responsibility by having a girl: full interaction



(a) Fathers

(b) Mothers

- Men tend to have less care responsibilities when they have a daughter
- The burden of the childcare is 50% more likely to be on their partners

Marriage: getting married after childbirth

Getting married by having a girl: full interaction



• Men are less likely to get married after having a daughter

Fertility:

Fertility by child being a girl: full interaction



Inconclusive, but points in the direction of greater likelihood of birth following



Political Progressiveness: Voting and Newspapers' Choice

Progressive Political Views by child being a girl: full interaction



Fathers of daughters tend to become *relatively* more progressive mechanisms

Conservative News by child being a girl: full interaction



 Mothers of daughters tend to become more conservative, while fathers tend to become more progressive (in relative terms) mechanisms Gender attitudes: higher gender index = more progressive/higher gender equality

Gender Index by child being a girl: full interaction



Women tend to become more conservative while their firstborn daughter is growing up

Different Samples

- Unbalanced UKHLS 18,065 parent x year unbalUKHLS
- BHPS 24,873 parent x year unbalBHPS

next

Discussion and Next Steps

- Childbirth has much larger negative effects on mothers' than on fathers' labour market outcomes
- Child penalty is larger when having a firstborn daughter:
 - probability to be employed, total working hours, real wage
- Fathers of daughters are less likely to get married after childbirth
- Mothers tend to spend more time with daughters and on household chores
- Fathers of daughters are less involved in child-rearing
- Mothers of daughters tend to become more conservative in their views, while fathers become more progressive

Thank you very much for your attention! Gràcies! Gracias!

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Appendix

| | Fathers | | | Mothers | | |
|--------------------------------|---------|----------|-----------------------|---------|----------|-----------------------|
| | Son | Daughter | p-value (Ha: Diff!=0) | Son | Daughter | p-value (Ha: Diff!=0) |
| Parental involvement | 0 | 0 | NA | 0 | 0 | NA |
| Fertility | 0 | 0 | NA | 0 | 0 | NA |
| Mental health | 49.83 | 50.03 | 0.553 | 47.82 | 47.80 | 0.942 |
| Progressive Political Views | 0.64 | 0.55 | 0.173 | 0.64 | 0.67 | 0.599 |
| Conservative News | 0.18 | 0.22 | 0.496 | 0.27 | 0.18 | 115 |
| Index Gender Attitudes | 18.11 | 18.12 | 0.989 | 18.93 | 18.74 | 0.245 |
| Time spent on household chores | 5.85 | 5.39 | 0.503 | 8.26 | 8.33 | 0.918 |
| Poor Mental health | 0.28 | 0.28 | 0.986 | 0.36 | 0.35 | 0.731 |
| | | | | | | |

Explaining the gap

| | (1) | (2) | (3) |
|--------------------------------|----------------|---------------|------------|
| | Being Employed | Working hours | Real wage |
| Parental Involvement | -0.01 | -1.12*** | -67.65*** |
| Having another child | -0.09*** | -5.67*** | -0.75 |
| Progressive Political Views | -0.05*** | -3.43*** | -272.20*** |
| Conservative News | -0.02 | -2.27*** | -421.93*** |
| Index Gender Attitudes | 0.014*** | 0.78*** | 32.36*** |
| Time spent on household chores | -0.006*** | -0.58*** | -50.36*** |
| Poor Mental Health | -0.06*** | -3.47*** | -96.99** |
| Observations | 3,780 | 3,404 | 3,104 |

 $\it Notes.$ Robust standard errors (in parentheses). Ethnic minorities are excluded. Balanced sample

* $\rho < 0.1$, ** $\rho < 0.05$, *** $\rho < 0.01$

Total working hours



- 1st Year Penalty: Total working hours \rightarrow 35%
- 1st Year Penalty (Balanced): Total working hours \rightarrow 35%

Total working hours by child gender



(a) Boy

(b) Girl

- 1st Year Penalty: Boy vs Girl \rightarrow 36% vs 40%
- 1st Year Penalty (Balanced): Boy vs Girl \rightarrow 34% vs 37%

Total working hours by child being a girl: full interaction



· Similar trend of mothers decreasing their working hours when they have a girl



- 1st Year Penalty: Real wage ightarrow 19%
- 1st Year Penalty (Balanced): Real wage ightarrow 19%

Real wage by child gender



(a) Boy

(b) Girl

- 1st Year Penalty: Boy vs Girl \rightarrow 19% vs 25%
- 1st Year Penalty (Balanced): Boy vs Girl \rightarrow 10% vs 23%

Real wage by child being a girl: full interaction



Similar to the balanced sample, the real wage is decreasing more for the mothers
of girls

Total working hours



- 1st Year Penalty: Total working hours ightarrow 47%
- 1st Year Penalty (Unbalanced UKHLS): Total working hours \rightarrow 35%
- 1st Year Penalty (Balanced UKHLS): Total working hours \rightarrow 35%

Total working hours by child gender



(a) Boy

(b) Girl

- 1st Year Penalty: Boy vs Girl \rightarrow 48% vs 44%
- 1st Year Penalty (Unbalanced UKHLS): Boy vs Girl \rightarrow 36% vs 40%
- 1st Year Penalty (Balanced UKHLS): Boy vs Girl \rightarrow 34% vs 37%



- 1st Year Penalty: Real wage \rightarrow 26%
- 1st Year Penalty (Unbalanced UKHLS): Real wage ightarrow 19%
- 1st Year Penalty (Balanced UKHLS): Real wage ightarrow 19%

Real wage by child gender



(a) Boy

(b) Girl

- 1st Year Penalty: Boy vs Girl \rightarrow 29% vs 22%
- 1st Year Penalty (Unbalanced UKHLS): Boy vs Girl ightarrow 19% vs 25%
- 1st Year Penalty (Balanced UKHLS): Boy vs Girl ightarrow 10% vs 23%

Total hours child penalty by STEM



(a) STEM - boys

The second secon

(b) STEM - girls







(d) No STEM - girls

Real wage child penalty by STEM



(a) STEM - boys



(b) STEM - girls





(c) No STEM - boys

(d) No STEM - girls

Adopting Callaway & Sant'Anna estimation using working hours



 Preliminary comparison with two different control groups: never treated and not yet treated. Results are similar with stabilized inverse probability weighting and ordinary least squares.

Adopting Callaway & Sant'Anna estimation using working hours



 Preliminary comparison with two different control groups: never treated and not yet treated. Results are similar with stabilized inverse probability weighting and ordinary least squares.



Adopting Sun & Abraham estimation using working hours



• Sun and Abraham approach uses never treated as a comparison group

checks

How would the model change if we are including mother fixed effects?

• We estimate the following equation with mother fixed effects:

$$Y_{ist}^{g} = \sum_{j \neq 1} \alpha_{j}^{g} \cdot \mathbb{1}[j=t] + \sum_{k} \beta_{k}^{g} \cdot \mathbb{1}[k = age_{is}] + \sum_{j} \gamma_{y}^{g} \cdot \mathbb{1}[y=s] + \mu_{i} + \varepsilon_{ist}^{g},$$
(1)

- where the difference owes to the µ_i term requiring the estimation of a single parameter for each woman. At minimum, we must now have some multicolinearity as for each woman we cannot separately estimate a fixed effect and full set of dynamic coefficients to time to birth, as well as age, as well as her average salary.
- Resolving this \rightarrow omitting two mother fixed effect terms.

How would the model change if we are including mother fixed effects?

Then, to calculate the proportional change term, we would once again follow:

$$P_t^g \equiv \frac{\widehat{\alpha}_t^g}{E[\widetilde{Y}_{ist}^g|t]},\tag{2}$$

however here

$$\tilde{Y}_{ist}^{g} = \sum_{k} \widehat{\beta}_{k}^{g} \cdot \mathbb{1}[k = age_{is}] + \sum_{y} \widehat{\gamma}_{y}^{g} \cdot \mathbb{1}[y = s] + \widehat{\mu}_{i}.$$
(3)

- The first two terms are estimated off cohorts, while the final term is estimated off each individual.
- Our counterfactual to compare to each women is the woman's average salary over the entire period, plus the cohort level wage and calendar year fixed effects, omitting time to birth dummies.
- As μ̂_i is estimated within woman, and each woman does have a birth in the sample period, then this μ_i term is partially contaminated by the effect of child birth.
- Many wage profiles will mean that this tends to lead to downward biases to coefficients in years leading up to child birth, and upward biases to coefficients in years following childbirth, and these biases will be worse as we move further away from the omitted baseline at year -1.



Simulation exercise

- We generate a simple wage path where mothers experience a decrease in their salaries by 50% in the year when they give birth.
- Salary at mean are set to be 2500 in year -1, then in year 0 falls by 1250.
- The simulations include mother-specific heterogeneity (mother FEs are relevant).



- No FE model \rightarrow exactly -50% in salaries. The model follows exactly the change that we have in our synthetic data.
- FE model \rightarrow -47%, which is less than "true" value. We face a multicollinearity issue in event time (omitted last wave) and in calendar year (omitted 2020).

Fraction of chores made self



- 1st Year Penalty: Fraction of chores made self \rightarrow -4%
- Last Year Penalty: Fraction of chores made self ightarrow -13%

Fraction of chores made self by child gender



(a) Boy

(b) Girl

- 1st Year Penalty: Boy vs Girl \rightarrow -11% vs -2%
- Last Year Penalty: Boy vs Girl \rightarrow -24% vs 2%

Fraction of chores made self by being a girl: full interaction



(a) Fathers

(b) Mothers

Mothers of daughters start doing less chores by themselves

timechores