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# Station Heterogeneity and Asymmetric Gasoline Price Responses

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Current Situation				

• Public perception that retail fuel prices are higher than necessary; oil companies are suspected of collusion and asymmetric pricing strategies



'Exxon made more money than God last year' - Joe Biden



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#### Asymmetric Pricing in the Retail Fuel Market

- Aforementioned findings and perceptions are partly confirmed by the empirical literature, however, mixed results are found even for the same market (Asane-Otoo and Schneider, 2015; Verlinda, 2008)
- Asymmetric gasoline price responses at the pump in reaction to input cost fluctuations are well-documented, with two prevalent explanations for non-competitive pricing: (Eckert, 2013; Periguero-Garía, 2013; Cook and Fosten, 2019; Asane-Otoo and Dannemann, 2022)
  - Market power (Borenstein et al., 1997; Balke et al., 1998)
    - A higher number of competitors is thought to ensure competitive pricing behavior (Barron et al., 2004; Lewis, 2008; Bergantino et al., 2020)
  - Search costs (Tappata, 2009; Johnson, 2002; Lewis, 2011)
    - With price comparison being costly for consumers, retailers can charge higher prices

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Aggregation	SSUES			

- The issue of data aggregation receives increasing attention, with focus on disaggregated data sets (Asane-Otoo and Dannemann, 2022; Frondel et al., 2020)
- Temporal aggregation inadequately reflects station level pricing decisions and masks pricing patterns '
- Spatial aggregation ignores heterogeneity across stations or regions, and masks patterns in the spatial distribution of pricing strategies (Haucap et al., 2017; Balaguer and Ripollés, 2016)



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Research Cor	tribution			

- However, the majority of studies relies on panel regression techniques and presents parameter estimates as averages across stations
- Panel aggregation bias receives little to no attention
- This paper:
  - Estimation of asymmetric error correction model for 12,613 individual station daily time series
  - Focus on parameter heterogeneity
  - Station-specific input costs (wholesale price)
  - Analysis of drivers of pricing patterns in local markets



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Retail Price D	ata			

- By law, fuel station operators are obliged to publicly disclose all price changes in real-time (German Federal Cartel Office, 2011)
  - Georeferenced station-level data available since 2014 for 15,227 stations
  - 12,613 stations ( $\approx$  83%) part of the regression sample
  - Daily price averages 2014-2018





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Input Cost Data				

- Wholesale rather than international crude oil prices might be relevant for retailers (Delpachitra, 2002)
  - Currently 18 active refineries, with 8 at least partly owned by vertically integrated companies
  - Refined products stored in 117 fuel depots
- $\Rightarrow$  Both types supply fuels to stations
  - Refineries and depots are assigned to 8 distinct wholesale market regions identified by an external data provider
  - Supply depot  $\rightarrow$  station is typically done by road transport
  - Openstreetmap driving distance used as a proxy for transport costs





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Further Variables				

- Further demand-side fluctuations are proxied for by considering relevant factors that affect motorists' behavior
- Neighborhood retail price of adjacent fuel stations to account for competition (Haucap et al., 2017; Atkinson, 2009)
- Ambient temperature and precipitation capture transportation mode choice based on weather conditions (Koetse and Rietveld, 2009; Klein Tank et al., 2002)
- Public holidays and school holiday start proxy for phenomena such as wave of vacationers (Cools et al., 2007; Jum, 2010)



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Cointegration	Fauation			

- Test for cointegration follows Engle-Granger two-step procedure (Granger and Engle, 1987)
- Ong-run relationship between retail fuel price and input cost is estimated:

$$\mathbf{r}_t = \sigma + \theta \mathbf{w}_t + \gamma' \mathbf{H} + \delta' \mathbf{D} + \xi_t \tag{1}$$

- **③** Test for stationarity of residual  $(\xi_t)$ 
  - With cointegration,  $\theta$  is long-run cost pass-through
  - Residual corresponds to deviations from long-run equilibrium



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Regression Ed	quation			

- Asymmetric Error Correction Model (AECM) estimated separately for 12,613 fuel stations and different input cost measures (Granger and Lee, 1989)
- Test for Symmetric pricing, that is,  $|\phi^+|=|\phi^-|$

$$\Delta r_{t} = \alpha + \phi^{+} \xi_{t-1}^{+} + \phi^{-} \xi_{t-1}^{-}$$

$$+ \sum_{l=1}^{L} \left( \beta_{l}^{+} \bigtriangleup r_{t-l}^{+} + \beta_{l}^{-} \bigtriangleup r_{t-l}^{-} \right)$$

$$+ \sum_{m=0}^{M} \left( \lambda_{m}^{+} \bigtriangleup w_{t-m}^{+} + \lambda_{m}^{-} \bigtriangleup w_{t-m}^{-} \right) + \sum_{n=0}^{N} \left( \sigma_{n}^{+} \bigtriangleup c_{t-n}^{+} + \sigma_{n}^{-} c_{t-n}^{-} \right)$$

$$+ \psi \bigtriangleup \overline{r}_{(-i)t-1} + \pi' \bigtriangleup \mathbf{W} + \gamma' \mathbf{H} + \delta' \mathbf{D} + \tau t + \varepsilon_{t}$$

$$(2)$$

with r retail, w wholesale and c crude oil price, and W weather, H holidays, D dayof-the-week and month/year FE



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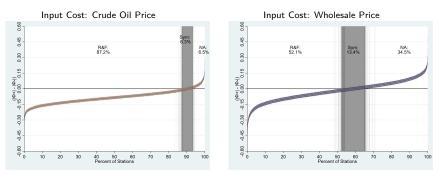
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## Part 1a: Input Cost Measure and Pricing Pattern

• Comparison of long-run adjustment parameters for individuals stations reveals drastic change in pricing patterns





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## Part 1b: Spatial Distribution of Pricing Patterns

- Large geographic differences of prevailing pricing patterns across federal states, for example, for rockets and feathers
- States with higher population densities and more refinery / fuel depot infrastructure have high share of rockets and feathers stations





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#### Part 2: Local Market Characteristics

• Magnitude of pricing asymmetry from the prior regression is employed as the dependent variable in a cross-section of stations

•	(1)	(2)	(3)	(4)	(5)
	Dependent \	/ariable: ( $ \phi^+ $ –	$ \phi^{-} )/\sqrt{SE_{\phi^{+}}}$	$N + SE_{\phi^-}/N$	
Total Number of Neighbors	-0.103*** (0.028)	-0.096** (0.029)	-0.175*** (0.039)	-0.175*** (0.039)	-0.073* (0.035)
Total Number of Neighbors <sup>2</sup>	0.003*** (0.001)	0.003*** (0.001)	0.004*** (0.001)	0.004*** (0.001)	0.002** (0.001)
No Neighbors	(****)	0.319 (0.441)	0.465 (0.444)	-0.587 (0.559)	-2.767*** (0.634)
In(Population Density)		()	0.466** (0.153)	0.504** (0.154)	0.111 (0.136)
Share of Major Brand			()	-1.322** (0.460)	-1.868*** (0.406)
Share of Same Brand				-1.276* (0.540)	-0.930 (0.476)
Share showing Rockets and Feathers				(0.0.0)	-10.035*** (0.496)
Share showing Negative Asymmetry					8.334*** (0.528)
Number of Stations	12,613	12,613	12,613	12,613	12,613
Federal State Fixed Effects?	Yes	Yes	Yes	Yes	Yes
Brand Fixed Effects? R <sup>2</sup> <sub>Adj.</sub>	Yes 0.078	Yes 0.078	Yes 0.079	Yes 0.079	Yes 0.285



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#### Part 3a: Pooled-Panel Estimation

#### • Estimation of a pooled-panel of all fuel stations

	(1)	(2)	(3)
	Crude Oil		Both
	Price	Price	
	Dependent Var	iable: $ riangle$ Retail P	rice of E5 Fuel
Input Cost Decrease $(\phi^+)$	-0.021***	-0.024***	-0.024***
	(0.000)	(0.000)	(0.000)
Input Cost Increase ( $\phi^-$ )	-0.056***	-0.042***	-0.042***
	(0.000)	(0.000)	(0.000)
	F-Tests for Syn	nmetry	
$\phi^+ = \phi^-$	20,022.02***	3,210.96***	3,146.30***
$\beta_{l}^{+} = \beta_{l}^{-},  l \in [1, 7]$	6,292.55***	5,150.56***	5,145.28***
$\lambda_m^+ = \lambda_m^-,  m \in [0,7]$		390.63***	429.67***
$\sigma_n^+ = \sigma_n^-,  n \in [0,7]$	4,522.24***		766.80***
Cointegration based on	Ct	Wt	Wt
Ν	21,621,581	21,621,581	21,621,581
R <sup>2</sup> <sub>Adi</sub>	0.286	0.332	0.332
Number of Stations	12,613	12,613	12,613
Month/Year Fixed Effects (Y)	Yes	Yes	Yes
Controls ( $W$ , $H$ , $T$ , $\triangle \overline{r}_{(-i)t-1}$ )	Yes	Yes	Yes

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#### Part 3b: Pooling by Pricing Pattern

• Fuel stations are pooled by the pricing pattern obtained from individual time series regression

6	(1)	(2)	(3)
	Rockets and	Symmetric	Negative
	Feathers	Adjustment	Asymmetry
	Dependent Vari	iable: $ riangle$ Retail Pr	ice of E5 Fuel
Input Cost Decrease ( $\phi^+$ )	-0.019***	-0.028***	-0.031***
	(0.000)	(0.001)	(0.000)
Input Cost Increase ( $\phi^-$ )	-0.049***	-0.040***	-0.032***
	(0.000)	(0.001)	(0.000)
	F-Test for Sym	metry	
$\phi^+ = \phi^-$	4,887.60***	179.18***	1.89
$\beta_{I}^{+} = \beta_{I}^{-},  I \in [1, 7]$	2,115.47***	618.19***	2,527.68***
$\lambda_m^+ = \lambda_m^-,  m \in [0,7]$	230.27***	45.84***	177.30***
$\sigma_n^+ = \sigma_n^-,  n \in [0,7]$	425.37***	87.40***	267.95***
Cointegration based on	Wt	Wt	Wt
Ν	11,480,801	2,428,496	7,712,284
$R^2_{Adi.}$	0.328	0.336	0.339
Number of Stations	6,718	1,421	4,474
Month/Year Fixed Effects (D)	Yes	Yes	Yes
Controls ( $W$ , $H$ , $T$ , $\triangle \overline{r}_{(-i)t-1}$ )	Yes	Yes	Yes

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Summary of N	lain Results			

- Individual station time series reveal wide scope of pricing patterns upon switching to wholesale prices as input costs
  - Findings masked by pooled-panel analysis
  - Controversially, individual price patterns deviate from pooled-panel results
- Brand structure and station density affect pricing pattern in local markets
- Geographical differences in prevailing pricing patterns
  - Access to refinery infrastructure as a possible explanation
  - States with many depots/refineries show higher prevalence of asymmetric pricing



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Conclusion				

- In spite of full market transparency in Germany facilitating consumer serach, asymmetric pricing patterns still prevail
- Use of crude oil rather than wholesale price seems to overstate prevalence of non-competitive pricing
- Market transparency unit makes price comparison possible, but could be a double-edged sword
  - Firms can also track competitors prices
  - Leads to tacit collusion and algorithmic pricing (Calvano et al., 2020)



## Thank you!

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