

Demographic Trends and the Transmission of Monetary Policy

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- There might be long-term effects for monetary policy as well (e.g., steady state levels of inflation and interest rates)
- What about **short-term** implications? This paper studies the impact of population aging on the effectiveness of monetary policy.

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 - ⇒ Age groups are heterogeneous in their consumption bundles
 - ⇒ Older people purchase more from product categories with higher levels of price rigidity (prices are adjusted less often)
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 - ⇒ Output responds more to MP shocks

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- Provide empirical evidence that output in U.S. states with higher old-age dependency ratio responds more to MP shocks (not shown today) ▶ Macro evidence
- Develop a **two-sector OLG-NK** model to:
 - ▶ Estimate the impact of demographic trends on MP propagation
 - ▶ Quantify the size of the new channel

- **Time-varying effects of monetary policy:** Boivin et al. (2010), Imam (2014), Galesi and Rachedi (2018), Kronick and Ambler (2019)
- **Monetary policy and demographic trends:** Carvalho et al. (2016), Aksoy et al. (2019), Eggertsson et al. (2019), Papetti (2019), Lis et al. (2020), Bielecki et al. (2020), Lisack et al. (2021), Fujiwara and Teranishi (2008), Kantur (2013), Yoshino and Miyamoto (2017), Leahy and Thapar (2020), Kimberly et al. (2021), ...
- **Flattening of the Phillips curve:** Bernanke (2010), McLeay and Tenreyro (2019), Jorda et al. (2019), Rubbo (2020)

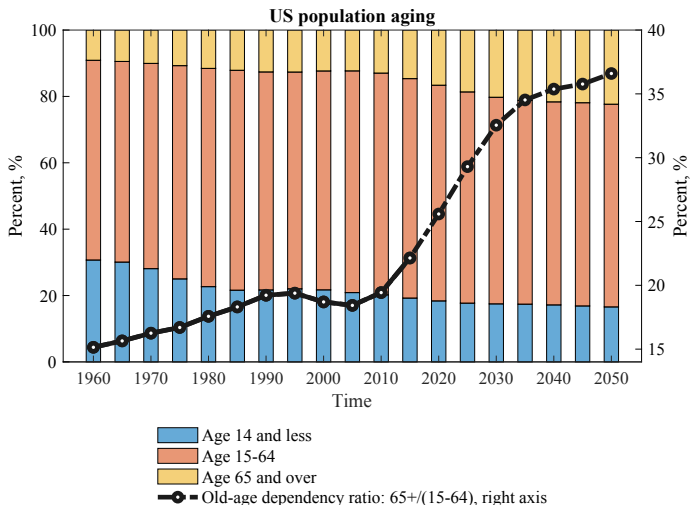
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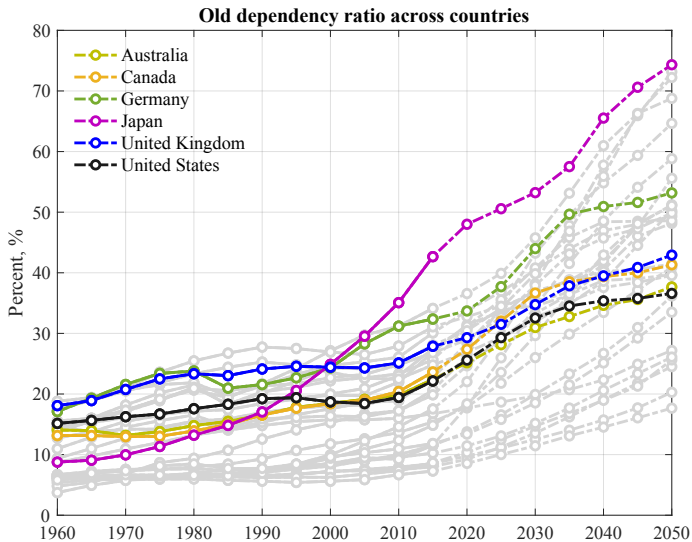
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Demographic trends



Source: UN (2017) World Population Prospects

Demographic trends



Source: World Bank Population Estimate and Projection

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Age-group level price stickiness

We combine data from:

- The **Consumer Expenditure Survey** (CEX) reports households' expenditures for around 600 Universal Classification Code (UCC) categories (e.g. *white bread*)
- The **frequency of price adjustment** estimated by [Nakamura and Steinsson \(2008\)](#) for 272 Entry Level Items (ELI) categories (e.g. *bread*)

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We define **age-group price stickiness** level as:

$$\theta_t^a = \sum_{j \in J} \omega_{t,j}^a \theta_j$$

with $\omega_{t,j}^a = \frac{C_{t,j}^a}{\sum_j C_{t,j}^a}$ the **expenditure weight** on category j for age group a and θ_j the **frequency of price adjustment**.

Frequency of price adjustment across age groups

▶ Over time

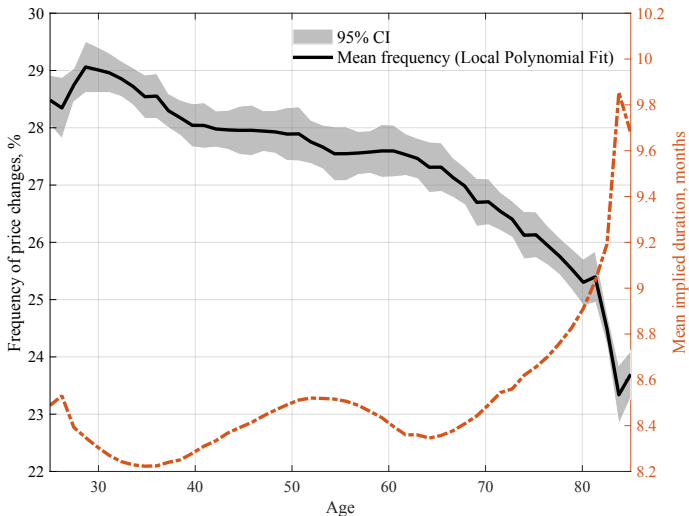
▶ Expenditure differences

▶ Expenditure weights

▶ Excluding sales

▶ Consumption quantiles

▶ Education



Alternative aggregation

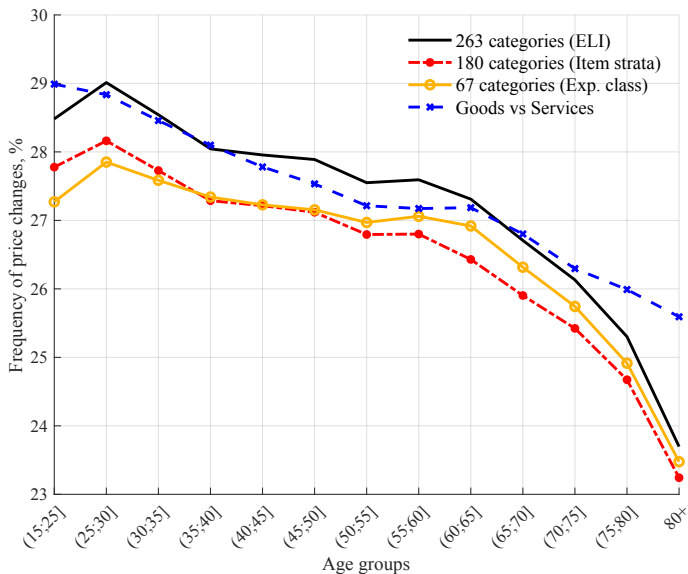


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OLG-NK model

Two-sector OLG-NK model ([Heer et al., 2017](#), [Bielecki et al., 2020](#) and [Papetti, 2021](#))

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Households of different age

- Choose consumption
- **Workers** supply labor, **retirees** receive pension
- Capital, bonds and firms' shares are transferred to a perfectly competitive and risk-neutral investment funds [▶ Equations](#)

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Firms of two sectors ▶ Equations

- Final-goods firms: Services and goods
- Monopolistically competitive intermediate-good producers
- **Calvo** price adjustment mechanism: $\theta^S > \theta^G$

Government

- Taxes labor income of workers ▶ Equations
- Provides pension benefit for retirees

Monetary authority

- Sets interest rate based on a **Taylor rule**

Demographic structure

- Households is born at age $j = 1$ (real life age of 15) and live for a maximum of $J = 85$ years (real life age of 99)
- Survive with an **age-specific probability** s_j from age j to age $j + 1$
- Work until $j_w = 50$ (real life age of 64) and then retire
- N_j denotes the **size of cohort** j relative to the overall population, $\sum_{j=1}^J N_j = 1$.

The maximization problem of the representative household of age j at time t is:

$$\max_{c_{t+i,j+1}, l_{t+i,j+1}, a_{t+i+1,j+i+1}} \mathbb{E}_t \sum_{i=0}^{J-j} \beta^i s_{j+i} \left(\frac{c_{t+i,j+i}^{1-\sigma}}{1-\sigma} - \phi \frac{l_{t+i,j+i}^{1+\eta}}{1+\eta} \right)$$

with

$$c_{t,j} = \left[\alpha_j \frac{1}{\eta} (c_{t,j}^S)^{\frac{\eta-1}{\eta}} + (1 - \alpha_j) \frac{1}{\eta} (c_{t,j}^G)^{\frac{\eta-1}{\eta}} \right]^{\frac{\eta}{\eta-1}}$$

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subject to:

$$P_{t,j} c_{t,j} + a_{t+1,j+1} = a_{t,j} R_t^a + y_{t,j}$$

$$y_{t,j} = (1 - \tau_t) w_t n_{t,j} h_j \mathbf{1}_{j \leq j_w} + pen_t \mathbf{1}_{j > j_w} + beq_t$$

$$a_{t,1} = 0 \quad a_{t+J+1,J+1} = 0$$

- We use the model to compare the transmission of monetary policy shocks around **three steady-states**:
 - ▶ 1980 (baseline), when CEX data becomes available
 - ▶ 2010
 - ▶ 2050, using population projection from World Bank
- The three steady states differ only in terms of:
 - ▶ **population distribution** N_j ▶ Population distribution
 - ▶ **mortality rate** $(1 - s_j)$ ▶ Mortality rate
 - ▶ **service preferences** α_j ▶ Service share ▶ Labor efficiency

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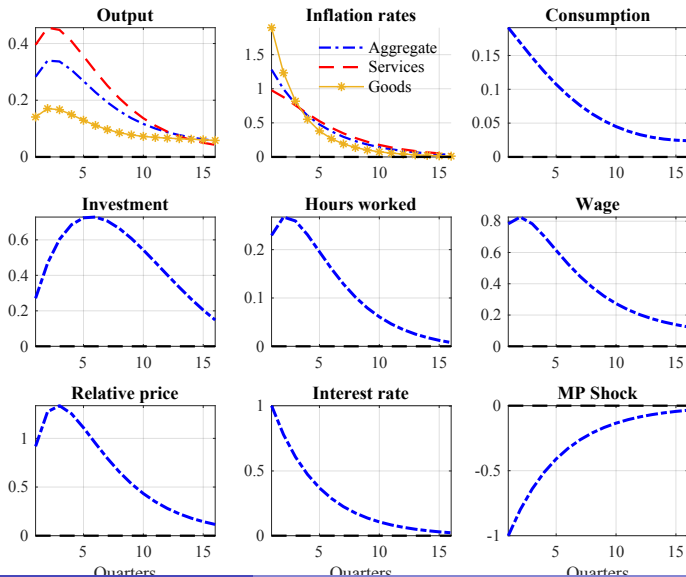
We answer the following questions:

- Do demographic trends change the way MP propagates in the U.S.?
- To what extent consumption heterogeneity across age groups contributes?

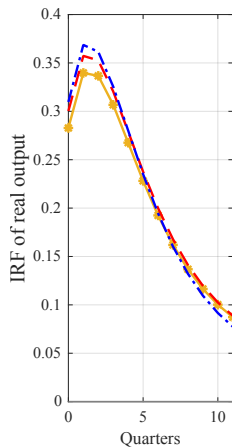
Parameter	Value	Description	Target
β	0.999	Discount factor	Annual interest rate between 4 and 5 %
δ	0.02	Depreciation rate	Capital-output ratio between 2 and 2.7
σ	1	Intertemporal elasticity of substitution	Standard value
ϕ	2	Frisch elasticity of labor supply	Standard value
N_j	• Population distribution	Population shares. Source: World Bank	
s_j	• Mortality rate	Survival probability. Source: Social Security Administration	
α_j	• Service share	Share of consumption devoted to services. Source: CEX	
h_j	• Labor efficiency	Individual life-cycle labor supply in efficiency units from Fullerton (1999)	Wage profile
ϵ	6	Elasticity of demand for each intermediate good	Steady-state markup of 20 %
θ^S	0.75	Calvo Frequency Services. Source: Nakamura and Steinsson (2008)	Price adjustment every 13 months
θ^G	0.25	Calvo Frequency Goods. Source: Nakamura and Steinsson (2008)	Price adjustment every 3 months
α	0.33	Capital share	Standard value

Model impulse response function to MP shock

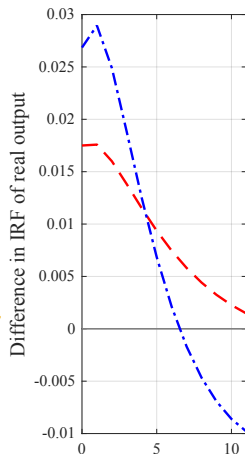
Age responses



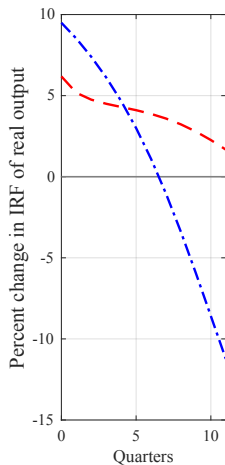
IRFs difference wrt baseline (changing only dem.), output



—●— Y_{1980}
- - - Y_{2010}
- · - · Y_{2050}

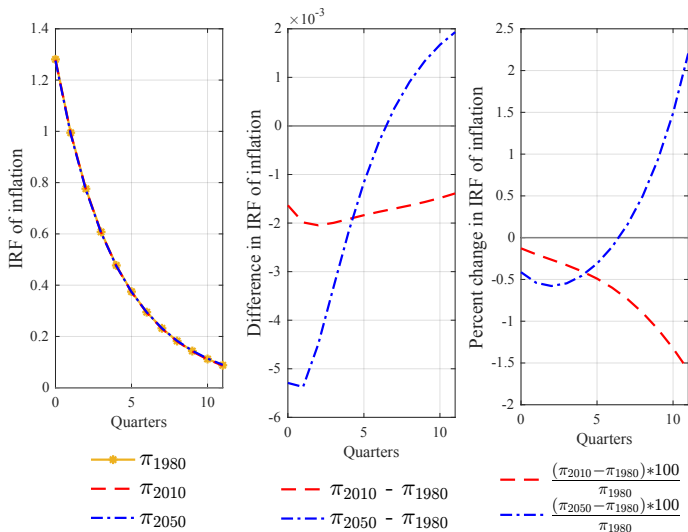


- - - $Y_{2010} - Y_{1980}$
- · - · $Y_{2050} - Y_{1980}$

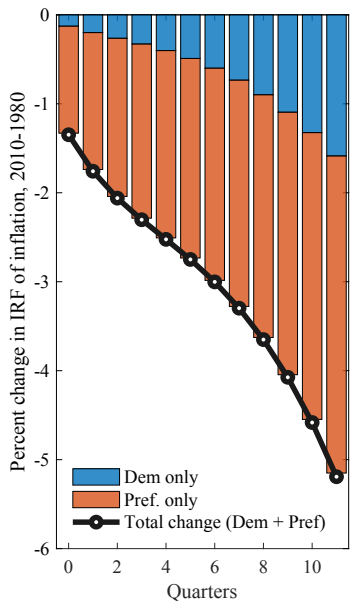
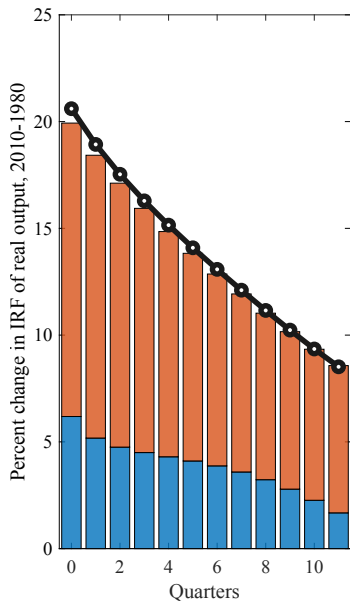


- - - $\frac{(Y_{2010} - Y_{1980}) * 100}{Y_{1980}}$
- · - · $\frac{(Y_{2050} - Y_{1980}) * 100}{Y_{1980}}$

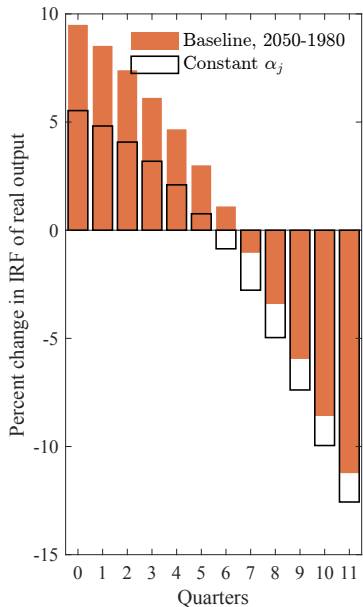
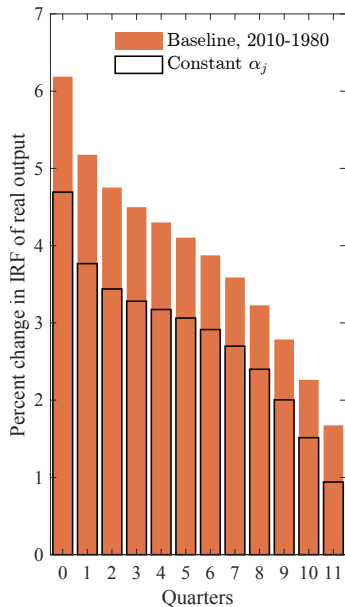
IRFs difference wrt baseline (changing only dem.), inflation



Contribution from demographics change



Size of the new channel, Output



Size of the new channel, Inflation

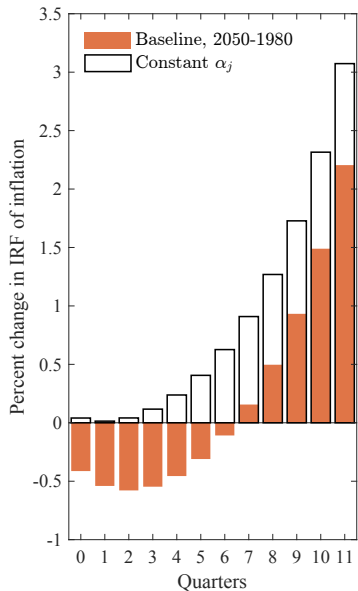
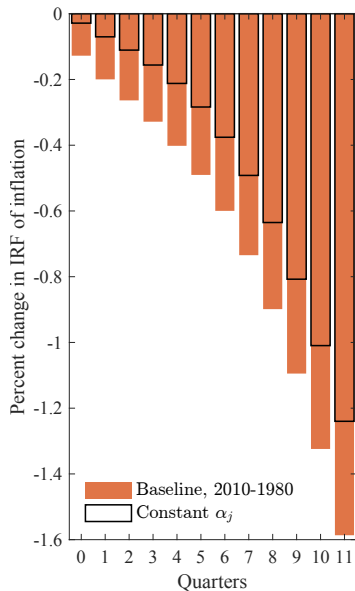


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Main results:

- Using micro data, document significant price stickiness heterogeneity across consumption bundles of different age groups

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- Using micro data, document significant price stickiness heterogeneity across consumption bundles of different age groups
- Develop a two-sector OLG-NK model to evaluate the effects of population aging on MP shocks propagation in the U.S.:
 - ▶ Output has become more responsive over time
 - ▶ Consumption differences across age groups significantly contribute to this effect

Thank you for your attention!

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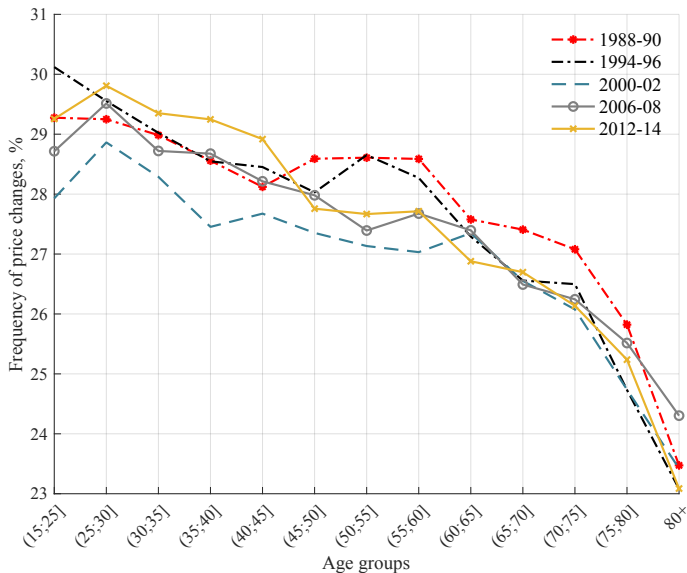
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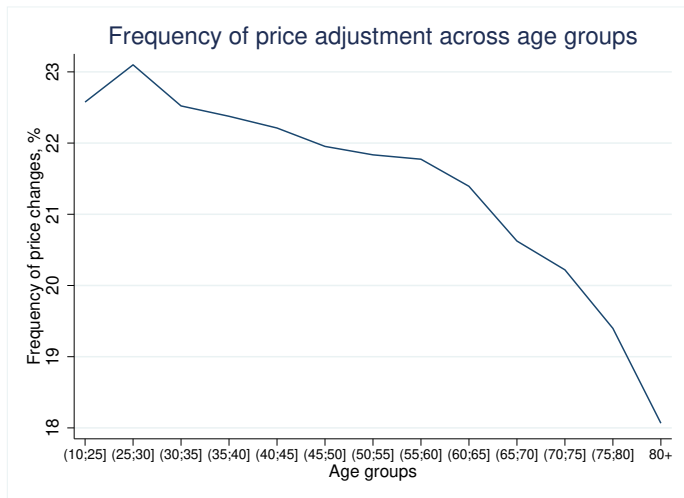
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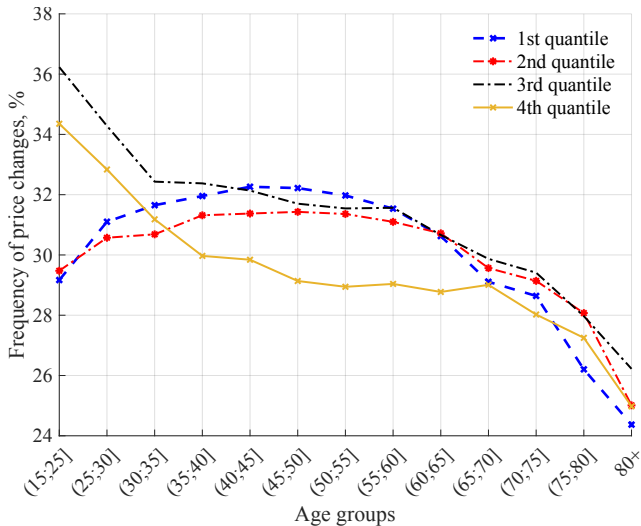
Across age groups and time



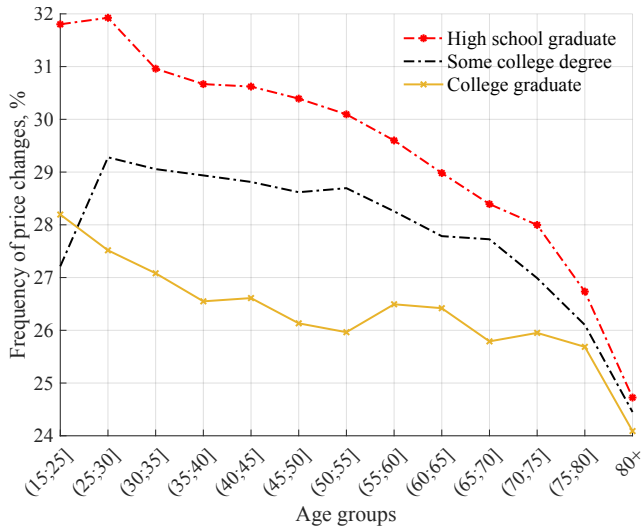
Frequency of price adjustment excluding sales



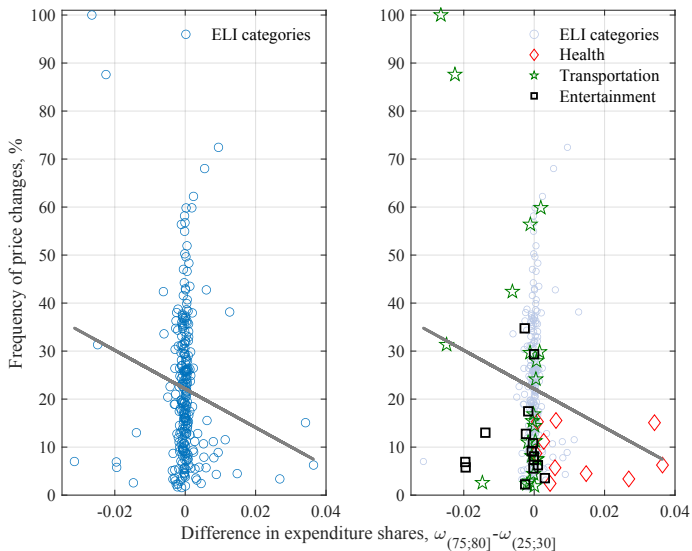
Across age groups and consumption quantiles



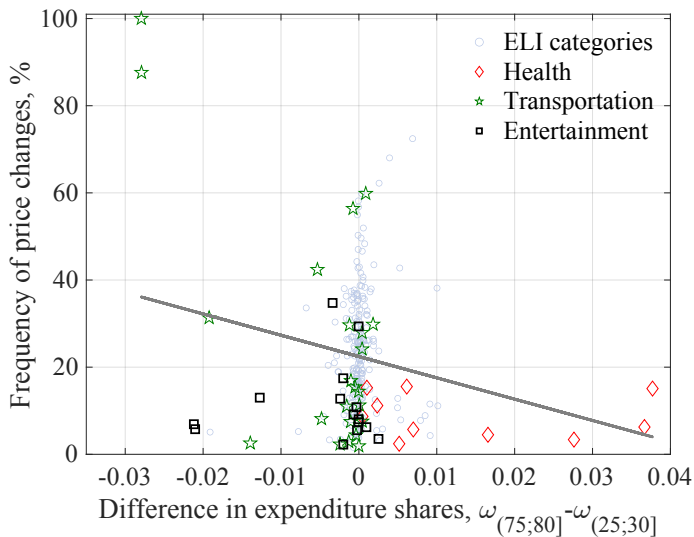
Across age groups and education levels



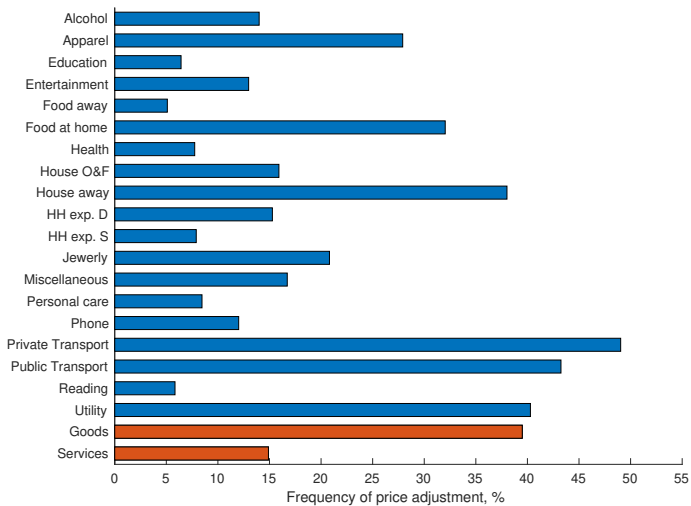
Expenditure differences across age group



Expenditure differences across age group



Price stickiness across categories



Within-between decomposition

What percentage of the increase in the share of services in total consumption is explained by changes in the age distribution?

The **share of services** in aggregate consumption can be written as:

$$\alpha_t^s = \frac{\sum_a C_t^{s,a}}{\sum_a \sum_j C_t^{j,a}} = \sum_a \alpha_t^{s,a} s_t^a$$

with $\alpha_t^{s,a} = \frac{C_t^{s,a}}{\sum_j C_t^{j,a}}$ and $s_t^a = \frac{\sum_j C_t^{j,a}}{\sum_a \sum_j C_t^{j,a}}$

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The change in services between t_1 and t_2 can then be decomposed in:

$$\Delta \alpha_t^s = \underbrace{\sum_a \Delta \alpha^{s,a} \bar{s}^a}_{\text{Within}} + \underbrace{\sum_a \bar{\alpha}^{s,a} \Delta s^{s,a}}_{\text{Between}}$$

with $\Delta x = x_{t_2} - x_{t_1}$ and $\bar{x} = \frac{x_{t_2} + x_{t_1}}{2}$ for any variable x .

Within-between decomposition

$$\Delta\alpha_t^s = \underbrace{\sum_a \Delta\alpha^{s,a} \bar{s}^a}_{\text{Within}} + \underbrace{\sum_a \bar{\alpha}^{s,a} \Delta s^{s,a}}_{\text{Between}}$$

Within-between decomposition, 1982-1990 to 2010-2018

	Service share	Contribution	Implied duration, months
Within	0.044	72.3 %	0.42 (+5.4 %)
Between	0.017	27.7 %	0.16 (+2.1 %)
Total	0.061 (46.69 % to 52.75 %)	100 %	0.58 (+7.5 %) (7.83 to 8.42)

5 Macro implications: U.S. states

U.S. states, data and empirical specification

Prediction: Economic activity in U.S. states with higher old-age dependency ratio should react more to MP shocks

▶ State variation

▶ Services

▶ Health

▶ Back

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The **average state level response** to a MP shock is estimated using Local Projection à la [Jordà \(2005\)](#):

$$y_{i,t+h} = \alpha_{i,h} + \beta_h MP_t + \theta_{i,h} X_{i,t-1} + \gamma_h X_{t-1} + \epsilon_{i,t+h}$$

for $h = 1, \dots, 16$

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- Dependent variable $y_{i,t}$:
 - ▶ Real Personal Income from the Bureau of Economic Analysis (BEA)
 - ▶ Annual inflation rate from [Hazell et al. \(2021\)](#)
 - ▶ GDP from the BEA (annual frequency)
- MP_t are the [Romer and Romer \(2004\)](#) shocks
- State controls $X_{i,t-1}$: lagged dependent variable and population size
- Aggregate controls X_{t-1} as in [Ramey \(2016\)](#): IP, CPI, FFR, unemployment rate and commodity price index
- Standard errors are clustered at state level.

Prediction: Economic activity in U.S. states with higher old-age dependency ratio should react more to MP shocks

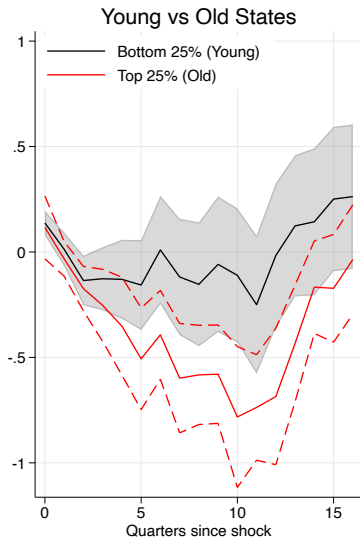
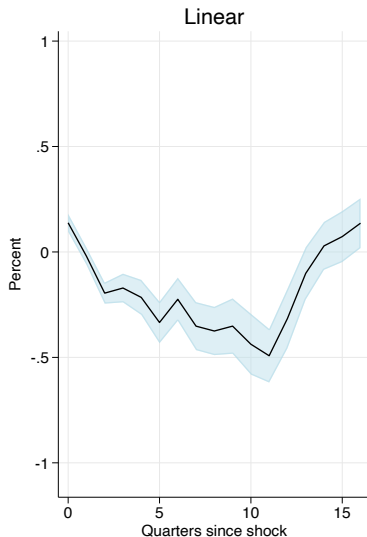
The **average state level response** to a MP shock is estimated using Local Projection *à la* [Jordà \(2005\)](#):

$$y_{i,t+h} = \alpha_{i,h} + \beta_h MP_t + \theta_{i,h} X_{i,t-1} + \gamma_h X_{t-1} + \epsilon_{i,t+h}$$

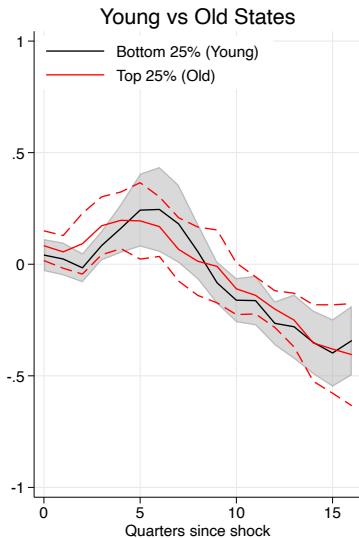
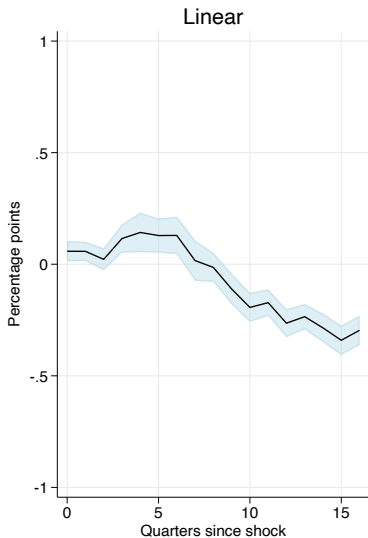
for $h = 1, \dots, 16$

As in [Cloyne et al. \(2018\)](#), we define **dummy variables** for different percentiles P of the old-age dependency ratio distribution and we interact them with our MP shock:

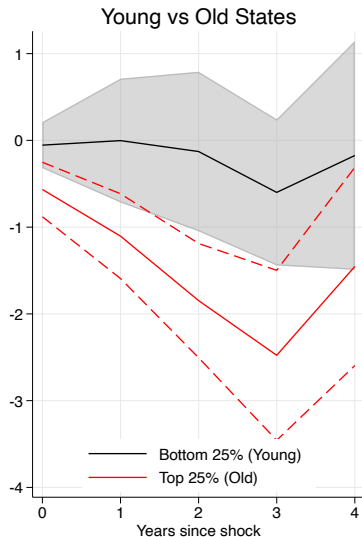
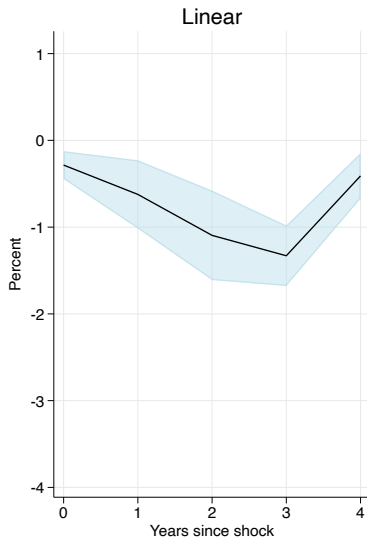
$$y_{i,t+h} = \alpha_{i,h} + \sum_{p=1}^P \gamma_h D_{i,t}^p + \sum_{p=1}^P \beta_h^p D_{i,t}^p MP_t + \theta_{i,h} X_{i,t-1} + \gamma_h X_{t-1} + \epsilon_{i,t+h}$$



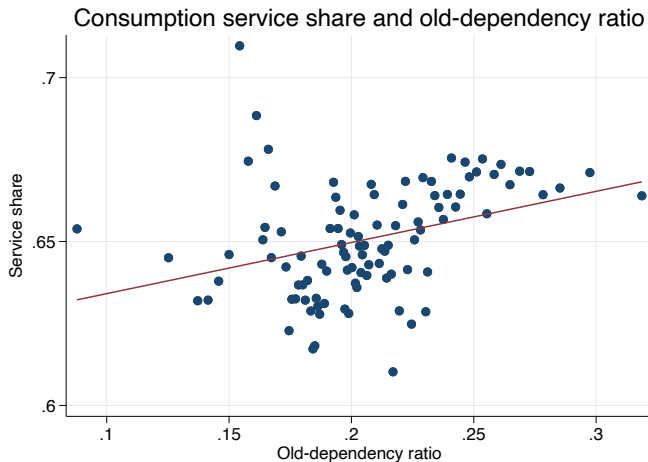
IRF Inflation rate



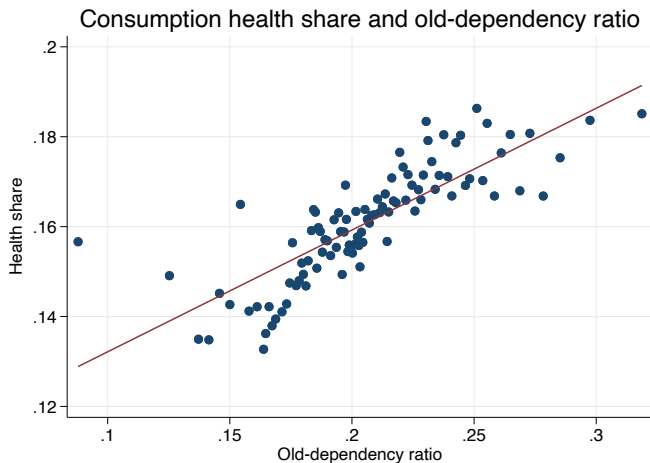
IRF Regional GDP



- Above/below median [▶ Go](#)
- Top/bottom 10 % [▶ Go](#)
- Share of working population [▶ Go](#)
- No financial crisis [▶ Go](#)
- Controlling for income [▶ Go](#)
- Excluding small states [▶ Go](#)
- High-frequency identification with IV-LP [▶ Go](#)
- Services [▶ Go](#)



Health share and old-age dependency ratio



The firms side has two sectors:

- **services** and **goods**
- Each sector has competitive final goods firm and a continuum of monopolistically competitive intermediate goods firms (standard NK model)
- Different frequency of price adjustment
- Only the output of the goods-sector can be used for capital investment

Price stickiness: Each period a fraction θ^S of intermediate producers cannot reset their price, $\theta^S > \theta^G$.

▶ Market clearing

▶ Back

Market clearing

Both aggregate labor and capital markets clear:

$$L_t = L_t^S + L_t^G = \sum_{j=1}^{jw} N_j h_j n_{t,j}, \quad K_t = K_t^S + K_t^G = \sum_{j=1}^J N_{j-1} a_{t,j}$$

$$beq_t = \sum_{j=1}^J (N_{j-1} - N_j) a_{t,j} R_t^a$$

The markets of goods and services clear:

$$Y_t^S = (K_t^S)^\alpha (L_t^S)^{1-\alpha} = C_t^S$$

$$Y_t^G = (K_t^G)^\alpha (L_t^G)^{1-\alpha} = C_t^G + I_t$$

and bonds are in zero net supply, $B_t = 0$.

▶ Back

The FOCs of the representative investment fund are:

$$K_{t+1} = (1 - \delta)K_t + \left[1 - \frac{S}{2}\left(\frac{I_t}{I_{t-1}} - 1\right)^2\right] I_t$$

$$A_{t+1} = q_t(1 - \delta)K_t + I_t + p_t^d$$

$$\frac{R_t^a}{\pi_t} A_t = \left[r_t^k + q_t(1 - \delta)\right] K_t + f_t + p_t^d$$

$$R_t q_t = \mathbb{E}_t \left[\left(r_{t+1}^k + q_{t+1}(1 - \delta) \right) \pi_{t+1} \right]$$

$$R_t p_t^d = \mathbb{E}_t \left[\left(p_{t+1}^d + f_{t+1} \right) \pi_{t+1} \right]$$

$$1 = q_t \left[1 - \frac{S}{2} \left(\frac{I_t}{I_{t-1}} - 1 \right)^2 - S \left(\frac{I_t}{I_{t-1}} - 1 \right) \frac{I_t}{I_{t-1}} \right] + \mathbb{E}_t \left[\frac{\pi_{t+1}}{R_t} q_{t+1} S \left(\frac{I_{t+1}}{I_t} - 1 \right) \left(\frac{I_{t+1}}{I_t} \right)^2 \right]$$

Government and Monetary Authority

The government funds a **pay-as-you-go social security system**. The tax rate on labor income τ_t is set such that the budget is balanced in each period.

$$pen_t = \bar{d}(1 - \tau_t)w_t \sum_{j=0}^{jw} N_j h_j$$

$$\tau_t w_t \sum_{j=0}^{jw} N_j h_j = pen_t \sum_{j=jw+1}^J N_j$$

with d_t the amount of pension benefit and \bar{d} the replacement rate.

The central bank follows the following simple **Taylor-type rule**:

$$\frac{R_t}{R} = \left(\frac{\Pi_t}{\bar{\Pi}}\right)^{\phi_\pi} \left(\frac{Y_t}{\bar{Y}}\right)^{\phi_y} e^{\nu_t^r}$$

The **sectoral Phillips Curve**:

$$\hat{\pi}_t^S = \beta \mathbb{E}_t \hat{\pi}_{t+1}^S + \kappa^S \hat{m}c_t^S$$

$$\hat{\pi}_t^G = \beta \mathbb{E}_t \hat{\pi}_{t+1}^G + \kappa^G \hat{m}c_t^G$$

with

$$\kappa^S = \frac{(1 - \theta^S)(1 - \theta^S \beta)}{\theta^S}, \quad \kappa^G = \frac{(1 - \theta^G)(1 - \theta^G \beta)}{\theta^G}$$

Since $\theta^S > \theta^G$, it follows that $\kappa^S < \kappa^G$.

The **sectoral Phillips Curve**:

$$\hat{\pi}_t^S = \beta \mathbb{E}_t \hat{\pi}_{t+1}^S + \kappa^S \hat{m}c_t^S$$

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with

$$\kappa^S = \frac{(1 - \theta^S)(1 - \theta^S \beta)}{\theta^S}, \quad \kappa^G = \frac{(1 - \theta^G)(1 - \theta^G \beta)}{\theta^G}$$

Since $\theta^S > \theta^G$, it follows that $\kappa^S < \kappa^G$.

It can be shown that:

$$\hat{\pi}_t = \omega \hat{\pi}_t^S + (1 - \omega) \hat{\pi}_t^G = \beta \mathbb{E}_t \hat{\pi}_{t+1} + \left[\omega \kappa^S + (1 - \omega) \kappa^G \right] (\hat{w}_t - \alpha(\hat{k}_t - \hat{l}_t)) - \lambda \hat{z}_t$$

where $\omega = \sum_j \alpha_j s_j \frac{P_j^{\eta-1}}{\sum_j s_j P_j^{\eta-1}}$, $s_j = \frac{N_j P_j C_j}{\sum_j N_j P_j C_j}$ and $\hat{z}_t = \log P_t^G - \log P_t^S$.

Effect of population aging on the slope of the Phillips Curve

	Baseline 1980	Dem+Pref 2010	Only Dem 2010
Service weight ω	0.4498	0.4953 (+10.11 %)	0.4542 (+0.97 %)
PC slope	1.2759	1.1773 (-7.72 %)	1.2665 (-0.74 %)

▶ Back

Response of Output and Inflation - Robustness Checks

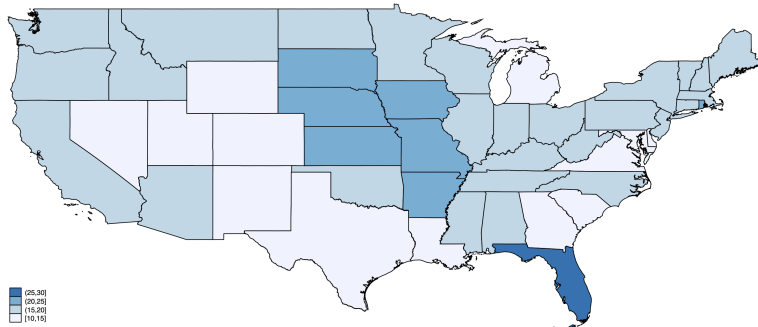
	Output response (%)			Inflation response (%)		
	Time 0	After 1 year	After 2 years	Time 0	After 1 year	After 2 years
Baseline	6.18	4.30	3.22	-0.12	-0.40	-0.89
Different ψ	5.63	4.01	2.93	-0.07	-0.26	-0.64
Different ϵ	5.07	3.72	2.83	-0.15	-0.34	-0.63
Different ϕ	6.97	4.58	2.95	-0.12	-0.36	-0.82
Constant τ	5.79	4.03	3.02	-0.09	-0.31	-0.71

*

Notes: The table reports the percent change in IRFs of output and inflation between 1980 to 2010 under alternative assumptions of the model.

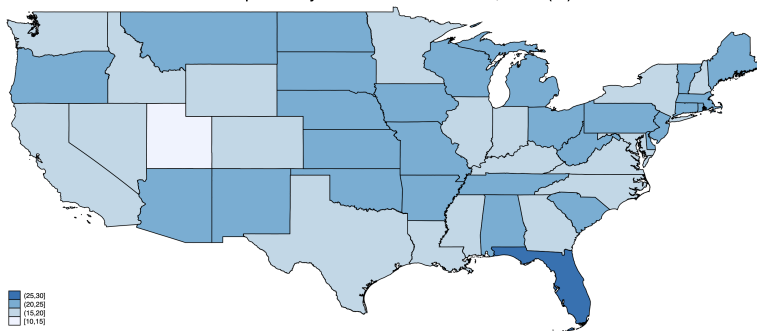
old-age dependency ratio across states

Old-dependency ratio across U.S. states, 1980 (%)



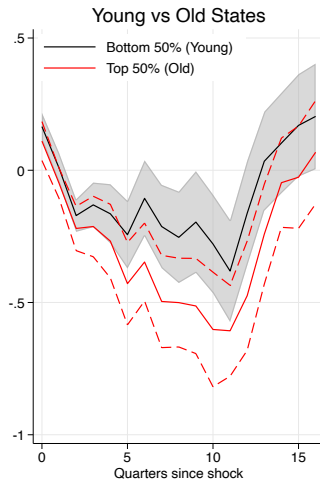
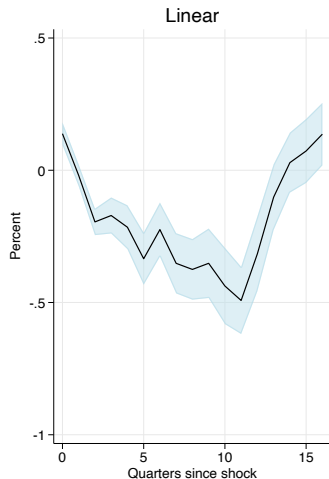
old-age dependency ratio across states

Old-dependency ratio across U.S. states, 2010 (%)

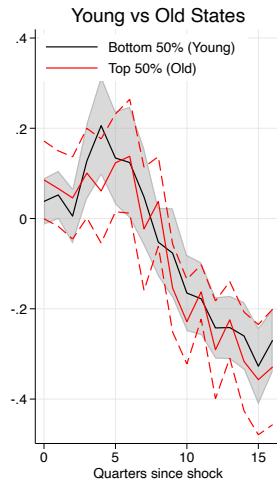
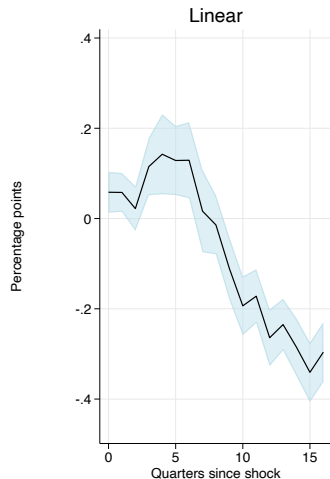


▶ Back

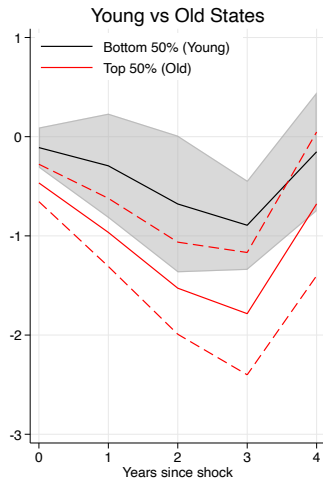
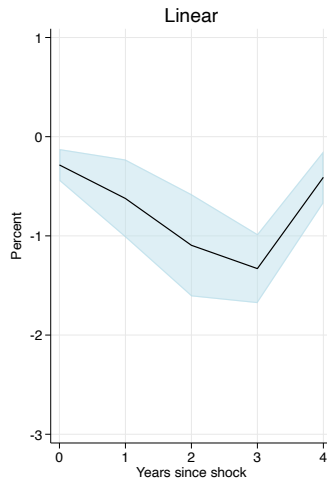
IRF Personal Income, Above/below median



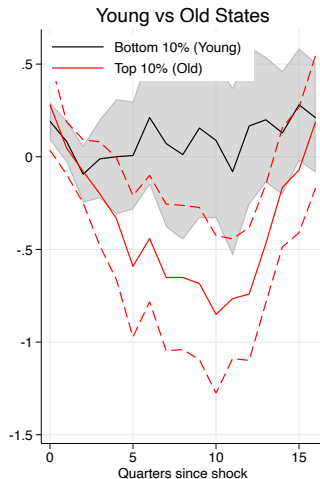
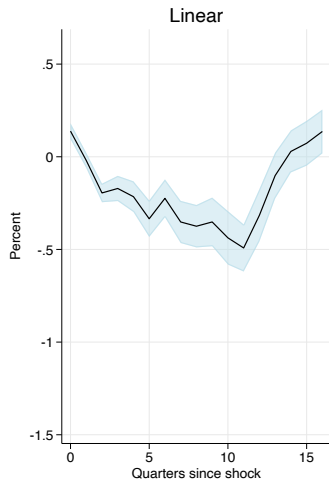
IRF Inflation rate, Above/below median



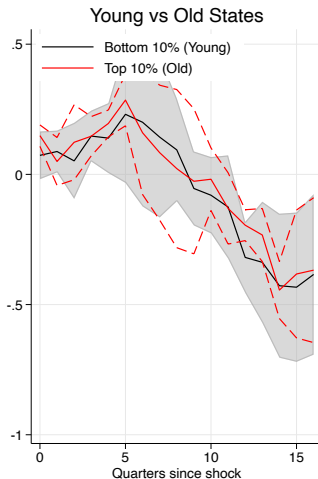
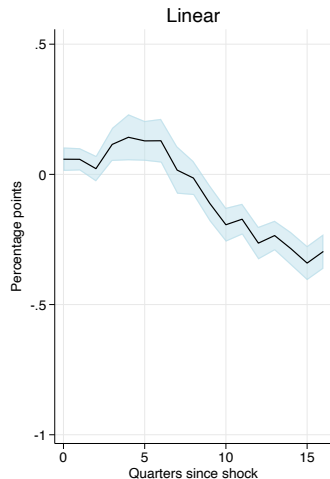
IRF Regional GDP, Above/below median



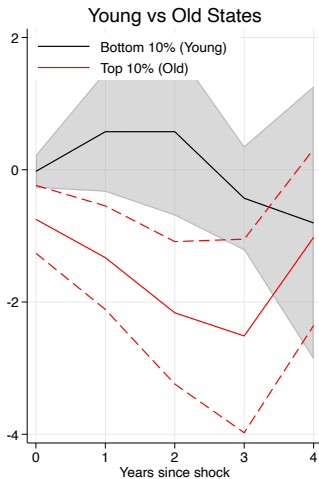
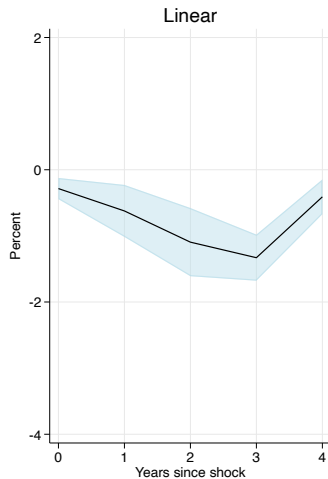
IRF Personal Income, Top/bottom 10 %



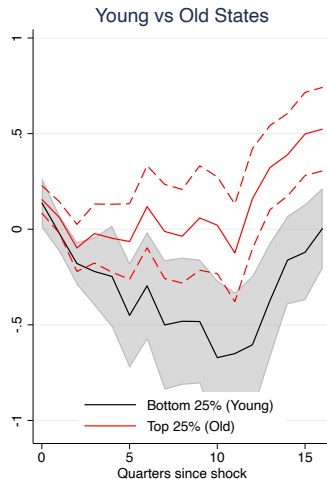
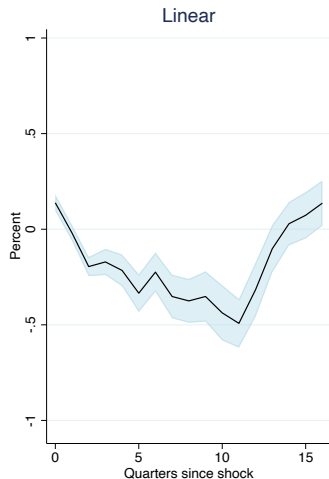
IRF Inflation rate, Top/bottom 10 %



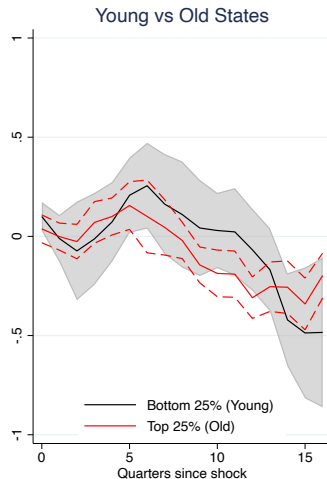
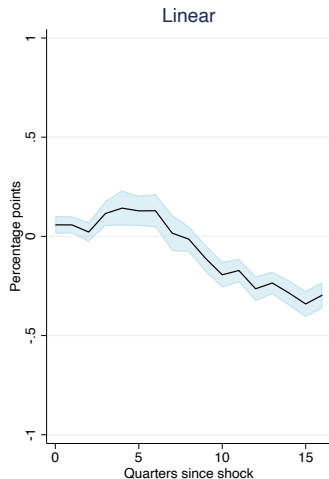
IRF Regional GDP, Top/bottom 10 %



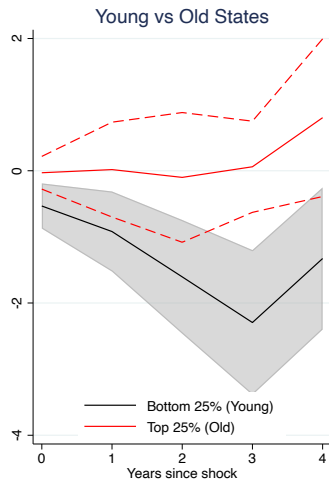
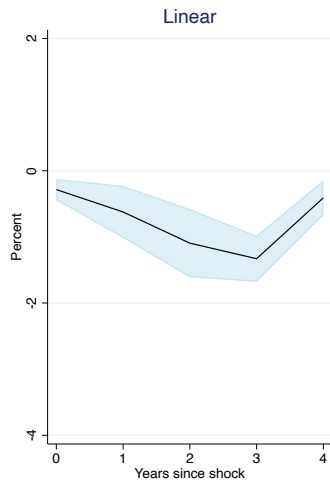
IRF Personal Income, share working population



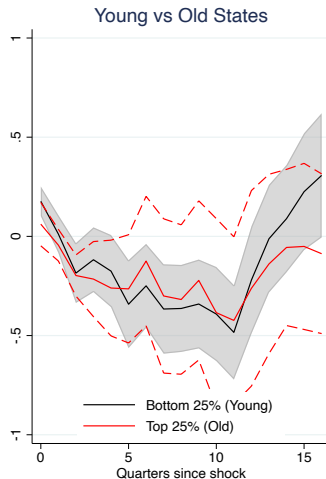
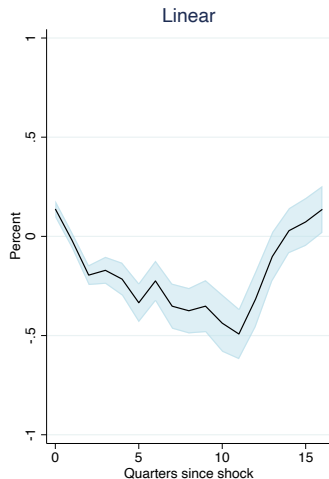
IRF Inflation rate, share working population



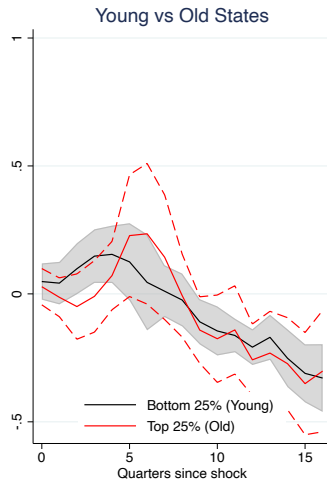
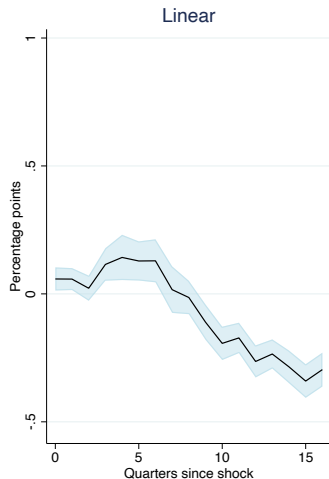
IRF Regional GDP, share working population



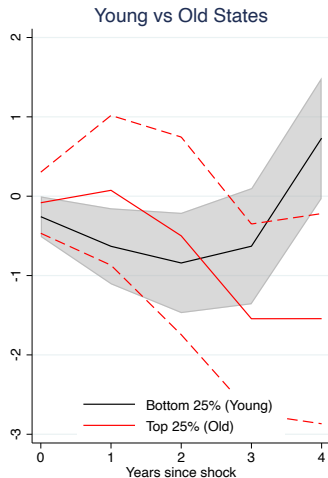
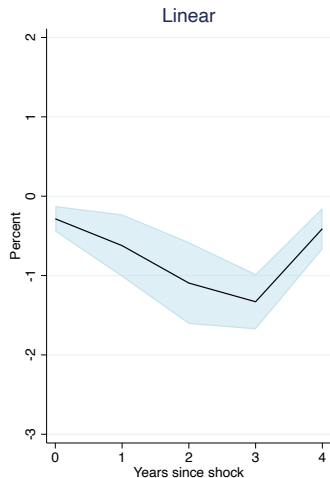
IRF Personal Income, share young



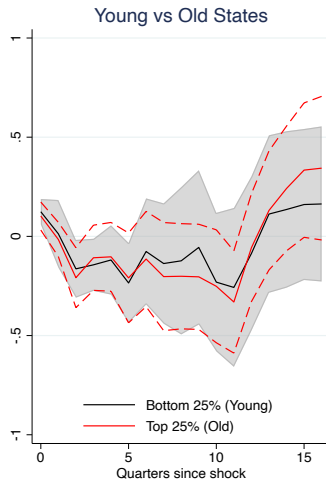
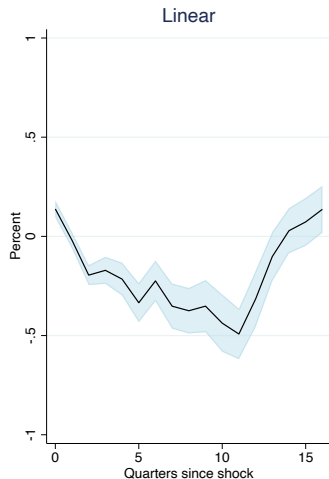
IRF Inflation rate, share young



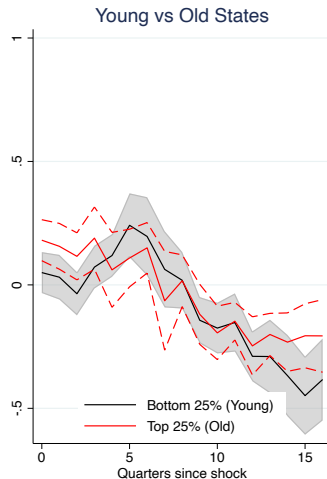
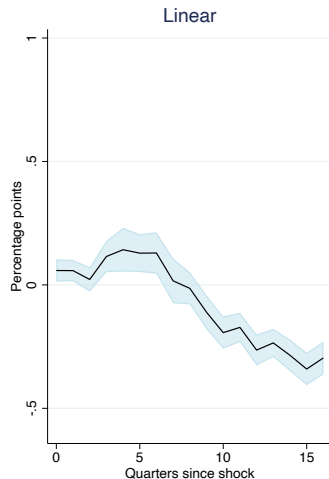
IRF Regional GDP, share young



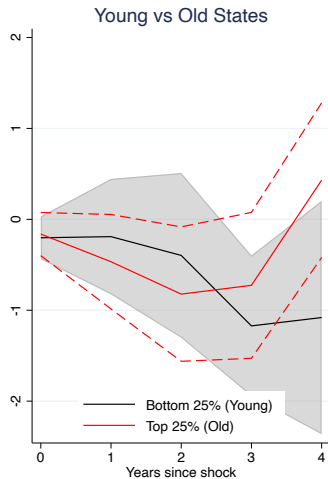
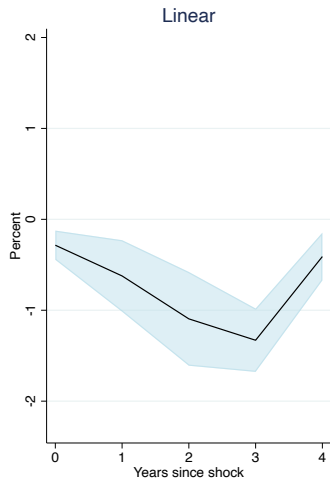
IRF Personal Income, share old

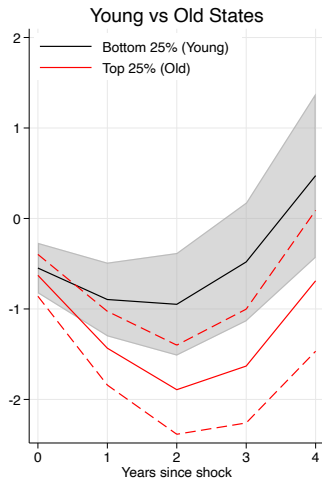
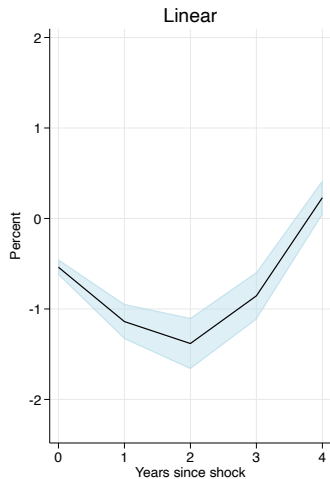


IRF Inflation rate, share old



IRF Regional GDP, share old





Lagged birth rates as IV

Baseline regression:

$$y_{i,t+8} = \alpha_i + \beta MP_t + \tau D_{i,t} + \delta MP_t \times (D_{i,t} - \bar{D}) + \theta_{i,h} X_{i,t-1} + \gamma X_{t-1} + \epsilon_{i,t}$$

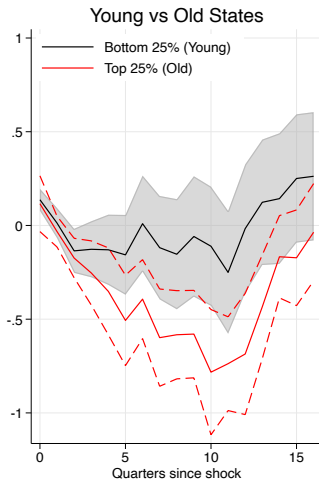
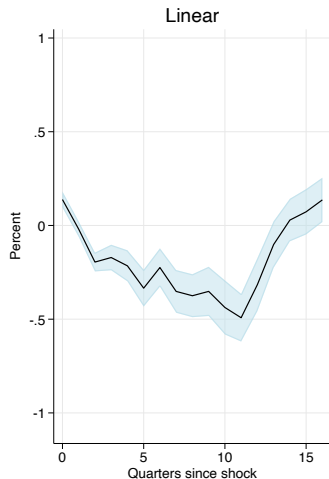
- $y_{i,t+8}$ is the 2 year ahead log of Personal Income
- $D_{i,t}$ share of working population

Regression table:

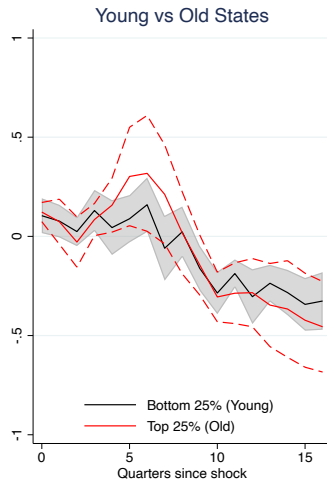
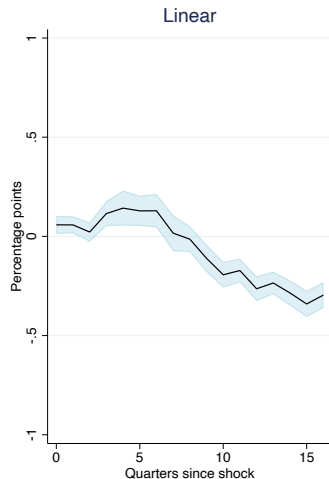
	(1)	(2)	(3)	(4)
	OLS	OLS	2SLS	2SLS
MP_t	-0.545*** (0.0905)	-0.473*** (0.0902)	-0.385 (0.252)	
$D_{i,t}$	-0.000767* (0.000434)	-0.000710* (0.000431)	-0.00747*** (0.00239)	-0.00697*** (0.00229)
$MP_t \times (D_{i,t} - \bar{D})$		0.227*** (0.0232)	0.649*** (0.216)	1.024* (0.526)
Observations	7701	7701	7392	7392
Controls	YES	YES	YES	YES
State FE	YES	YES	YES	YES
Time FE	NO	NO	NO	NO
First stage F stat.			17.74	22.60

Standard errors in parentheses

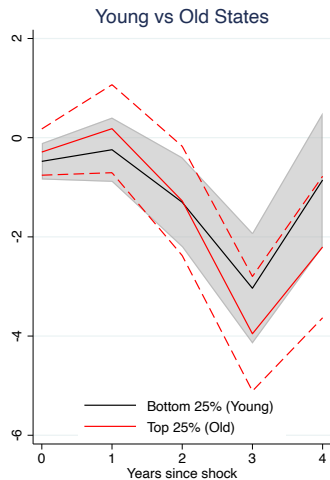
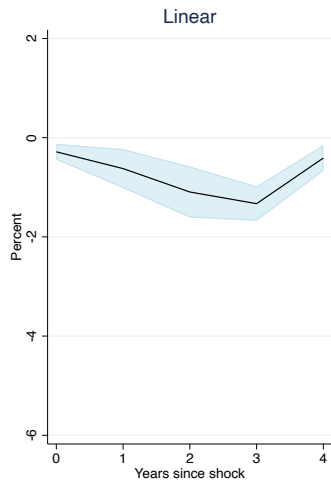
* $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$



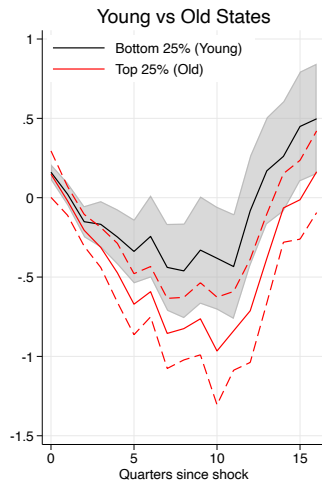
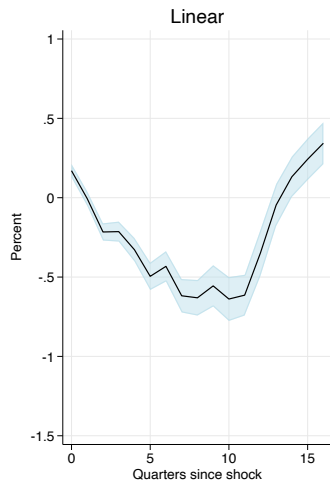
IRF Inflation rate, IV



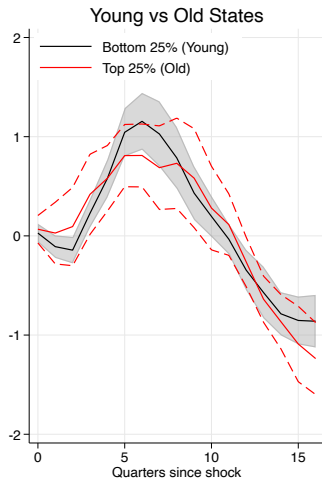
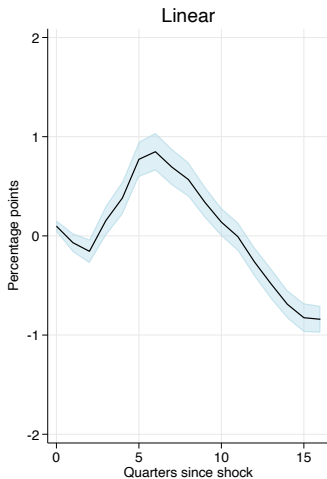
IRF Regional GDP, IV



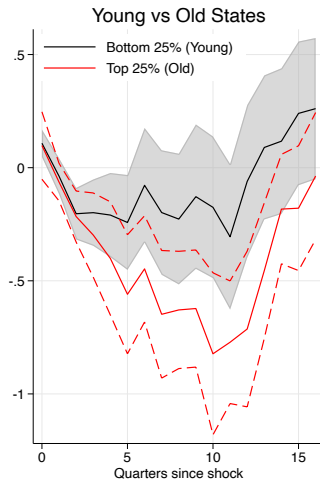
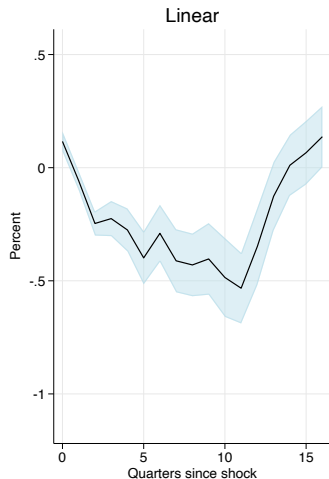
IRF Personal Income, No financial crisis



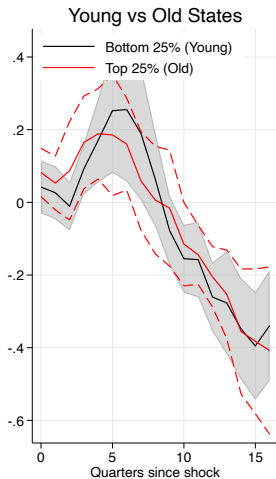
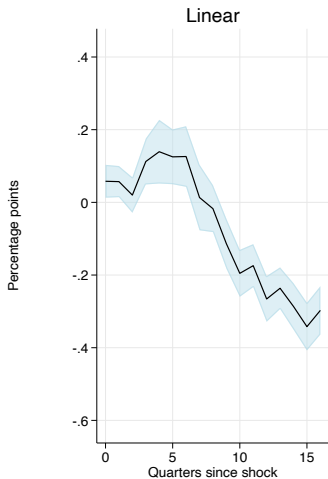
IRF Inflation rate, No financial crisis



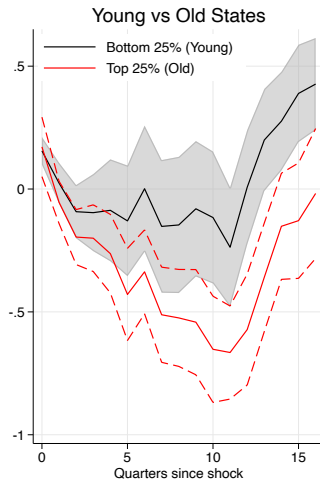
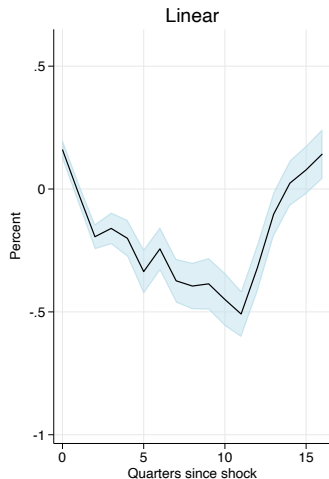
IRF Personal Income, Controlling for income



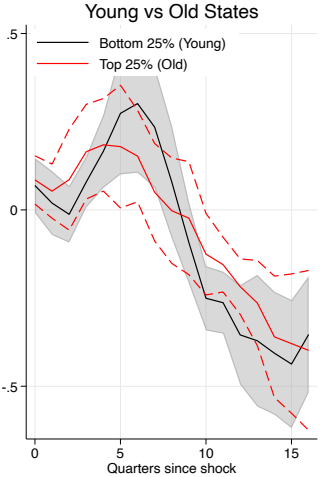
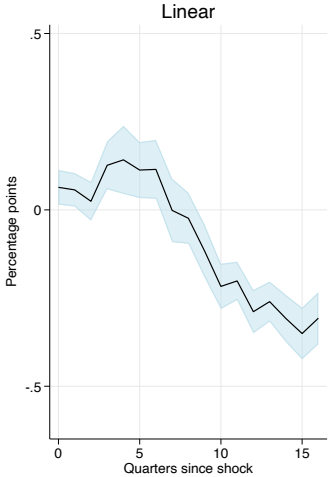
IRF Inflation rate, Controlling for income



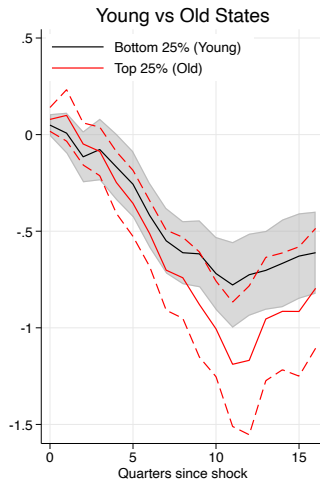
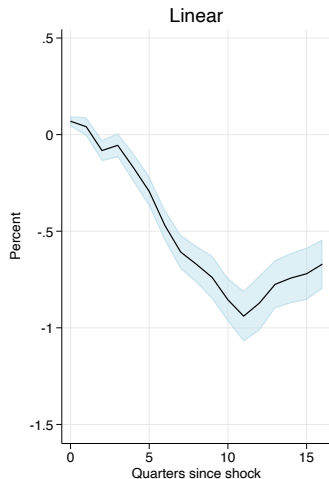
IRF Personal Income, no small states

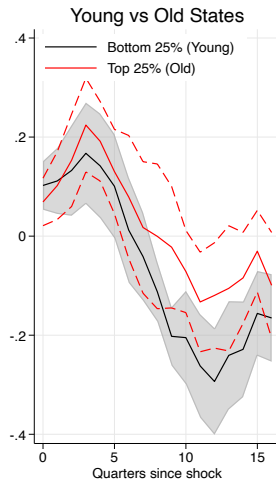
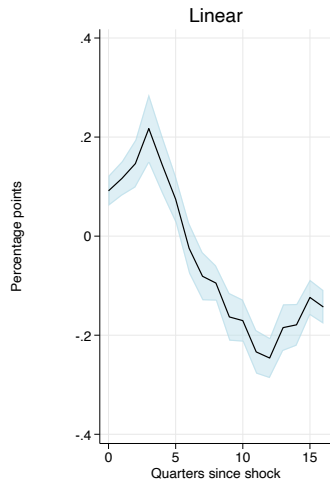


IRF Inflation rate, no small states

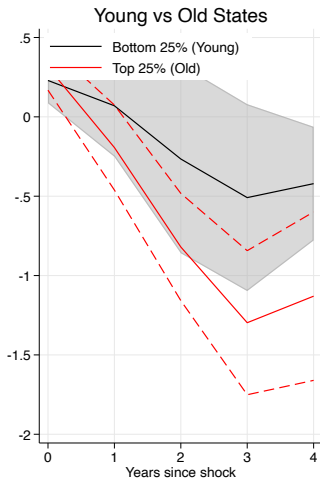
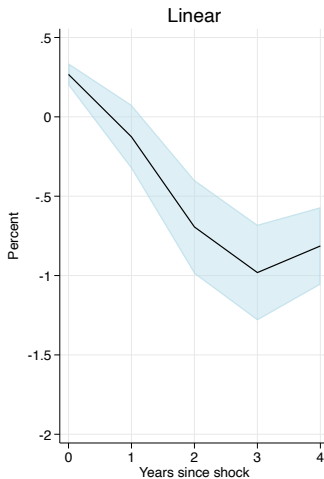


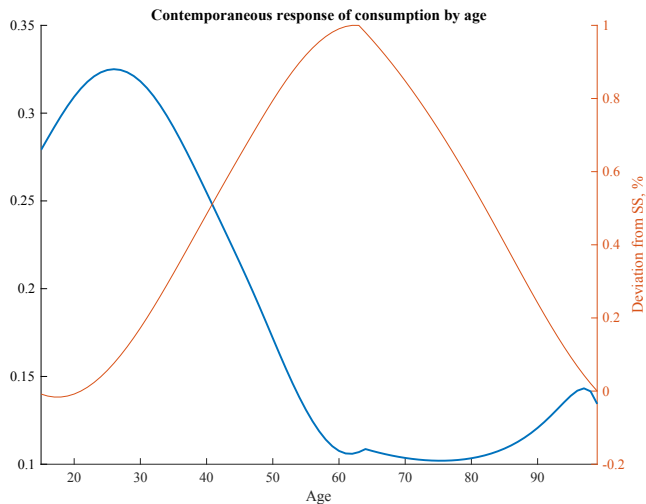
IRF Personal Income, IV-LP





IRF Regional GDP, IV-LP





Similarly to [Galesi and Rachedi \(2018\)](#) and [Ilzetzki et al. \(2013\)](#), we build a panel of quarterly data on inflation, output and interest rate from 1985:Q1 to 2015:Q4 for 44 countries

Similarly to [Galesi and Rachedi \(2018\)](#) and [Ilzetzki et al. \(2013\)](#), we build a panel of quarterly data on inflation, output and interest rate from 1985:Q1 to 2015:Q4 for 44 countries

We run a structural VAR on $Y_t = [\Delta \log GDP_t, \Delta \log P_t, R_t]$ with four lags:

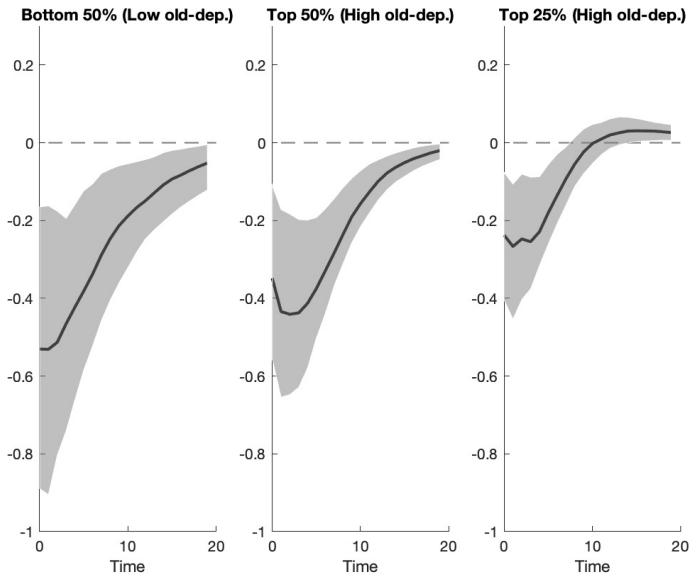
$$Y_t = A(L)Y_t + Bu_t$$

The monetary policy shocks are identified by **sign restrictions** on the impulse responses, a monetary policy shock:

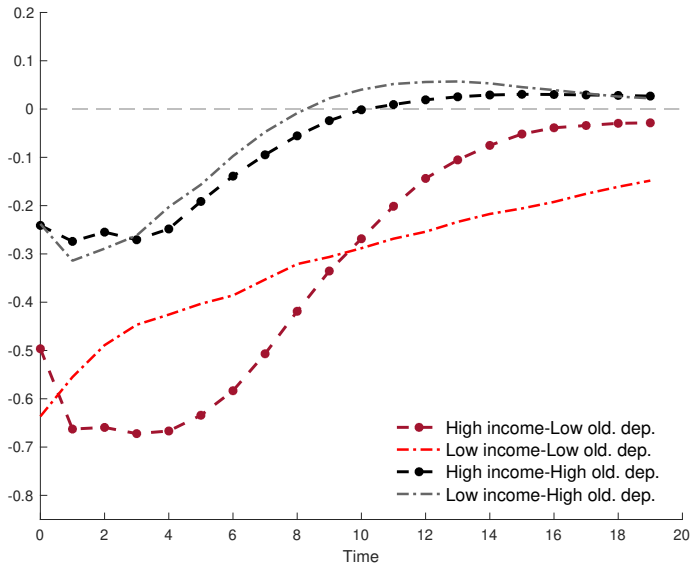
- Increases the nominal interest rate
- Reduces real output growth and the inflation rate on impact as well as in the following quarter.

▶ Country data

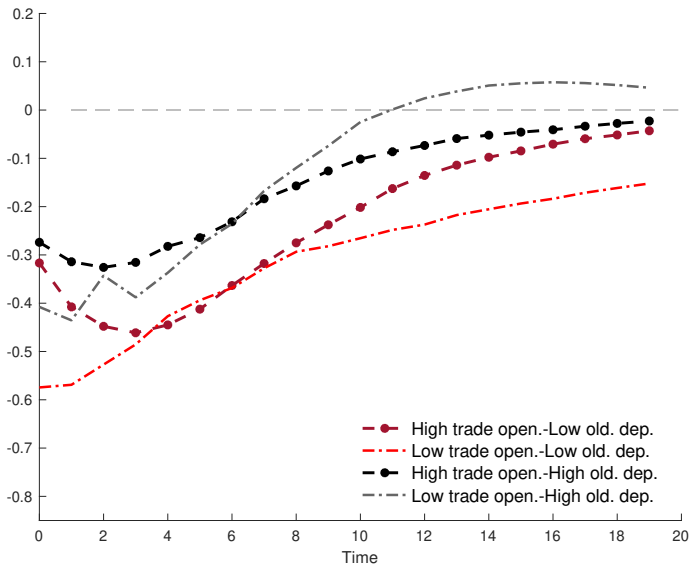
IRFs across countries' characteristics



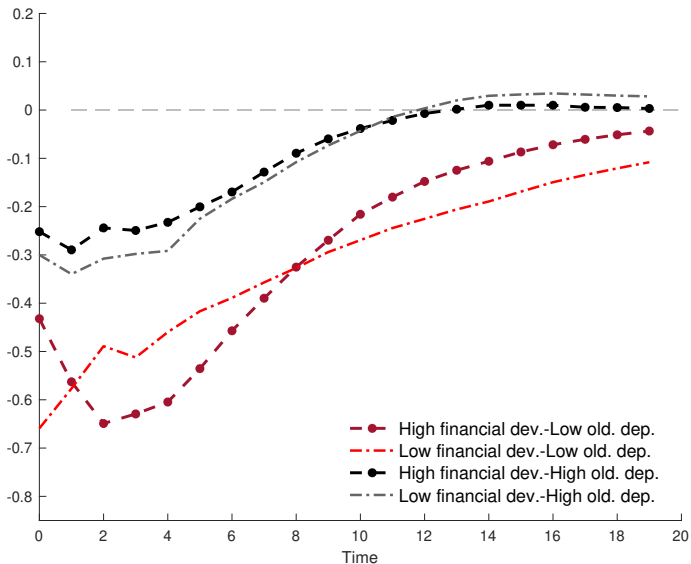
IRFs across countries' characteristics



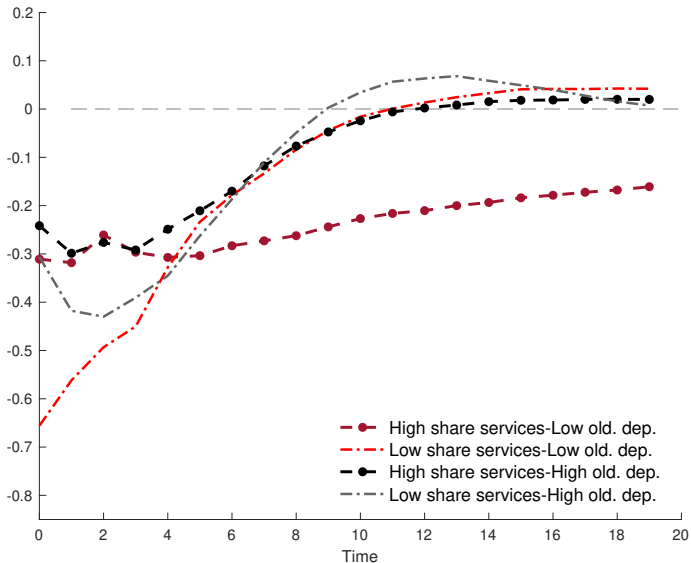
IRFs across countries' characteristics



IRFs across countries' characteristics



IRFs across countries' characteristics



We showed that:

- Older people purchase more from (relatively) **stickier-price sectors**
- Since the demographic structure of the US economy has significantly changed in the last 50 years, this result partially explains the observed **more dampened response of inflation** to monetary shocks

Based on this finding, **two hypothesis** can be tested:

- Inflation in countries with a higher old-age dependency ratio reacts less to MP shocks
- Inflation in Euro Area regions with a higher old-age dependency ratio reacts less to MP shocks.

▶ Euro Area

IRFs for the Euro Area

Dataset of CPI at NUTS-1/NUTS-2 level for 69 regions in the Euro Area (22 in Italy, 19 in Spain, 16 in Germany, 7 in Portugal and 5 in Finland) from 1999 to 2016

IRFs for the Euro Area

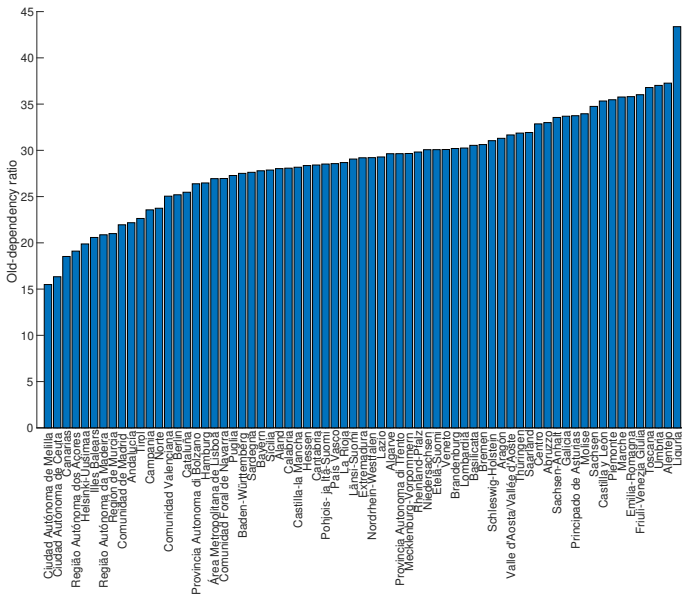
Dataset of CPI at NUTS-1/NUTS-2 level for 69 regions in the Euro Area (22 in Italy, 19 in Spain, 16 in Germany, 7 in Portugal and 5 in Finland) from 1999 to 2016

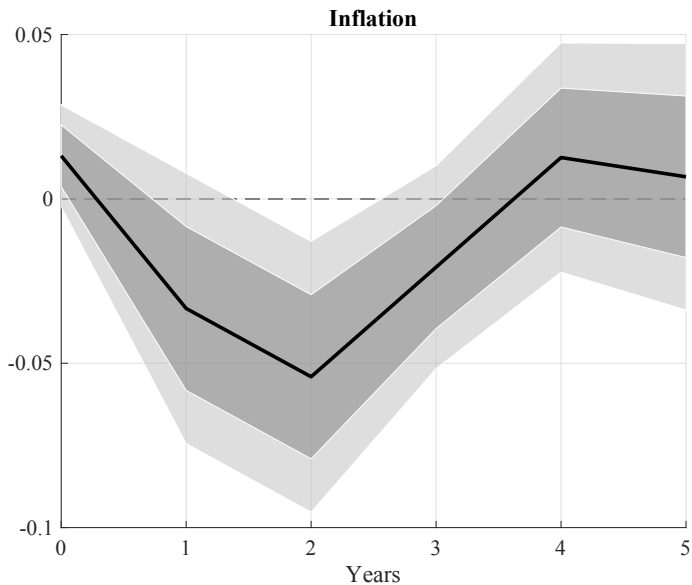
Estimate the empirical specification proposed by [Hauptmeier et al. \(2020\)](#):

$$y_{i,t+h} = \alpha_i + \beta_h \epsilon_t^{MP} + \gamma_h X_{i,t} + \delta_h X_{j,t} + \theta_h X_{k,t} + u_{i,t+h}$$

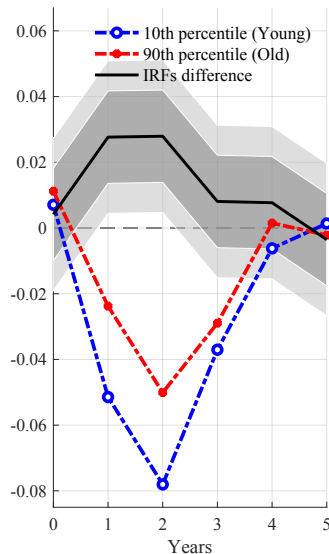
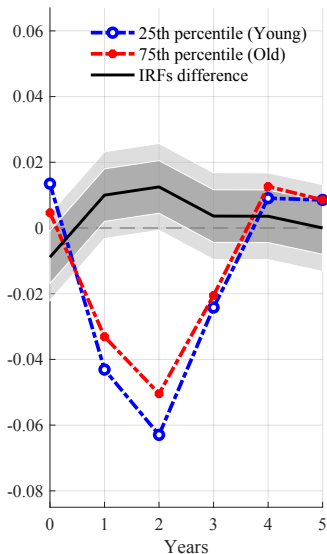
- ϵ_t^{MP} the high-frequency monetary policy shocks computed by [Jarociński and Karadi \(2020\)](#)
- $X_{i,t}$, $X_{j,t}$ and $X_{k,t}$ are vectors of time-variant control variables at the region-, country- and euro area-level (i.e. region fixed effects, population size, employment, share of services in gross value added, country and Euro Area GDP, Euro Area CPI).

old-age dependency ratio across European regions

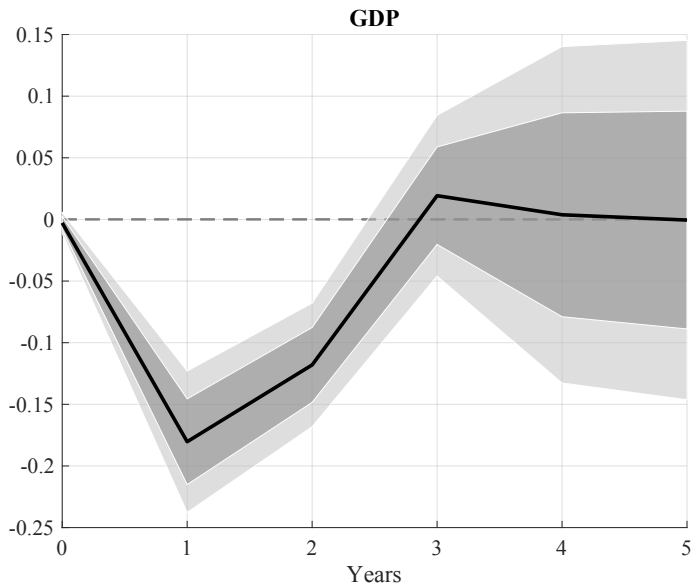




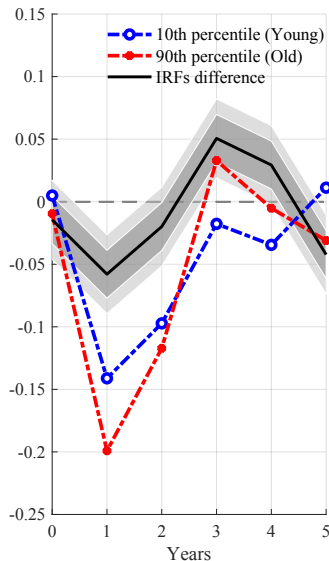
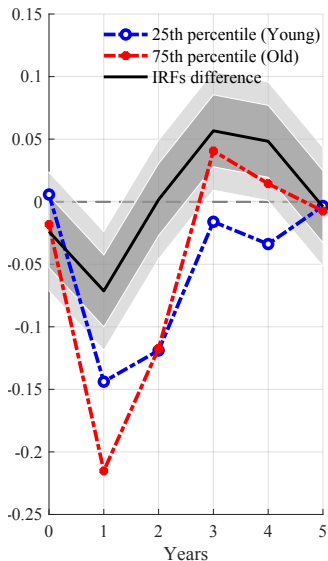
IRFs for the Euro Area



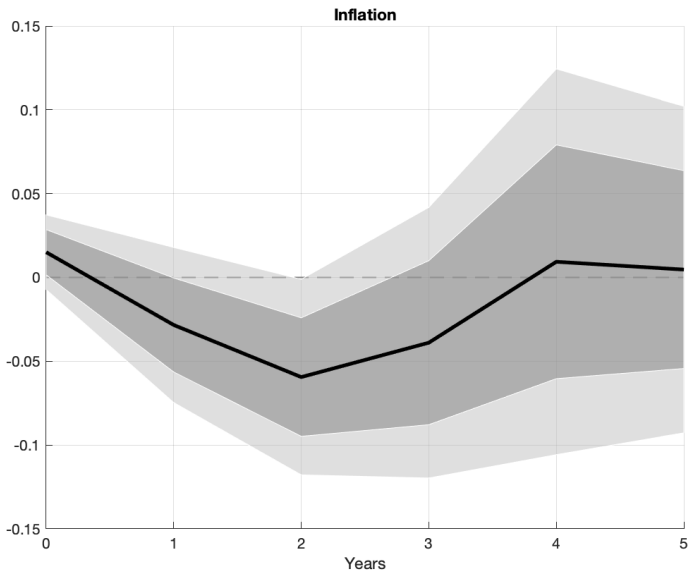
IRFs for the Euro Area



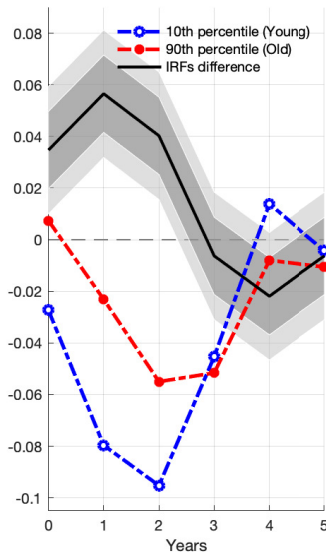
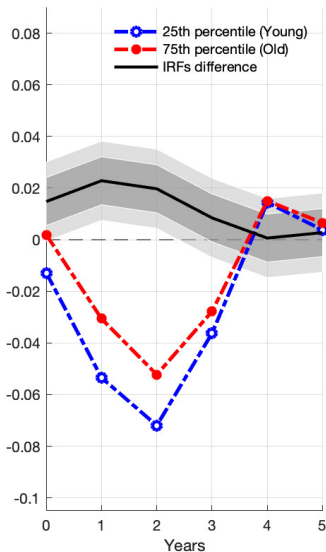
IRFs for the Euro Area



IRFs for the Euro Area, 3-month Eonia OIS

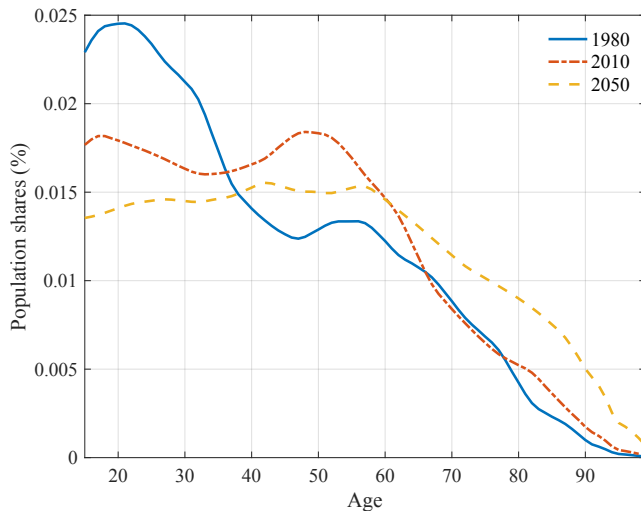


IRFs for the Euro Area, 3-month Eonia OIS



Population distribution

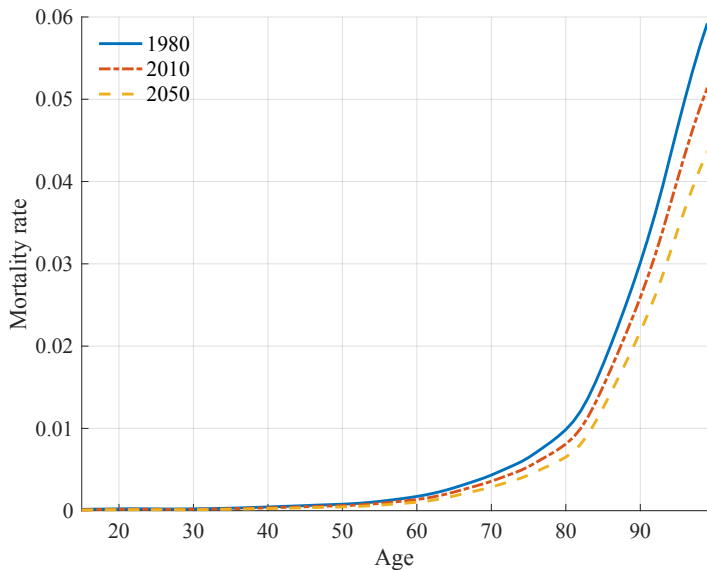
▶ Back



Source: UN (2017) World Population Prospects

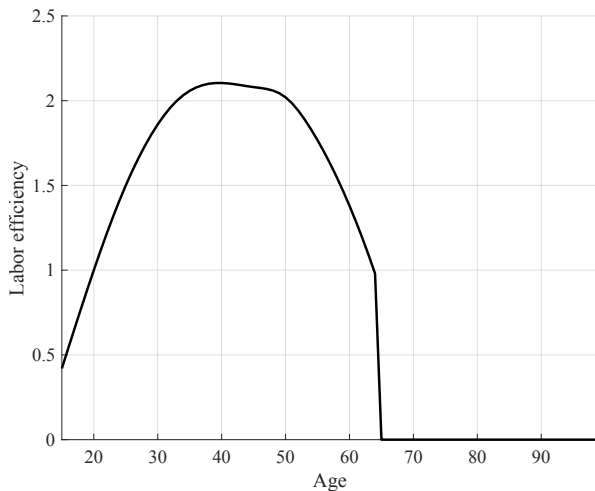
Mortality rate across age groups

▶ Back



Labor efficiency

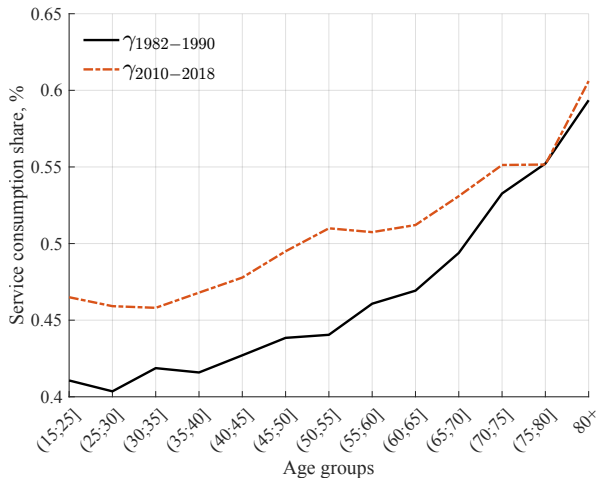
▶ Back



Source: Fullerton (1999)

Service consumption share

▶ Back

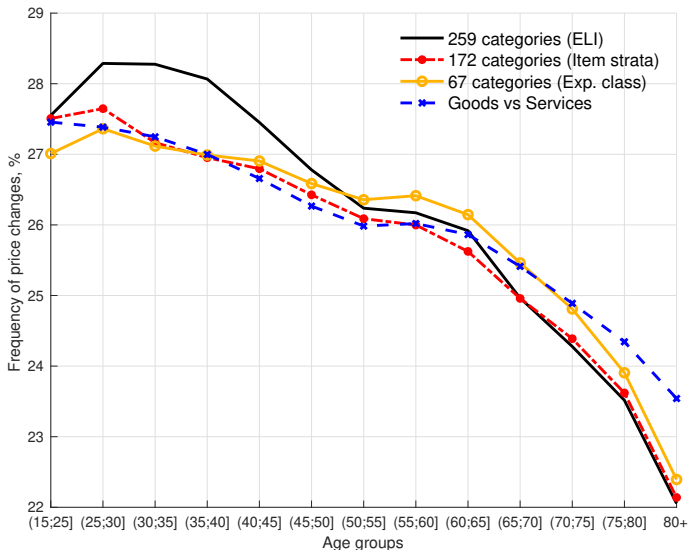


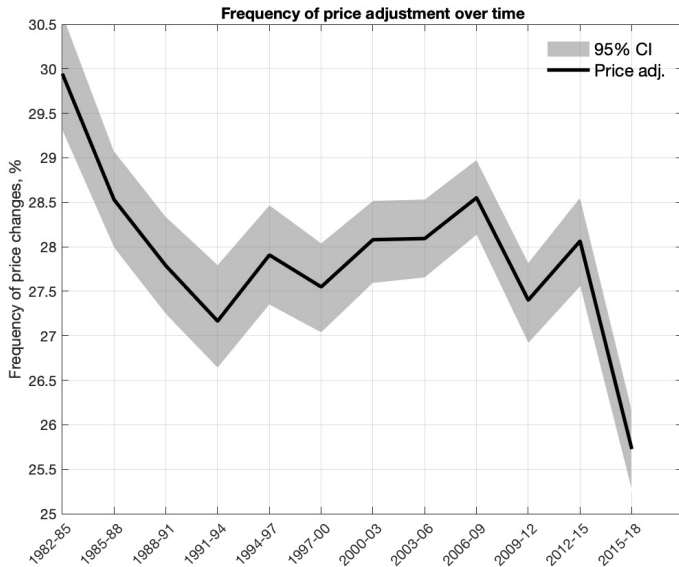
Source: Own calculation, CEX data

Summary results

Parameters	1980 → 2010	Parameters ↗	
		π^{IRF}	$Output^{IRF}$
Service preference α_j	↗	↘	↗
Survival probabilities s_j	↗	↗	↘
Retirement age jw	↗	↗	↘

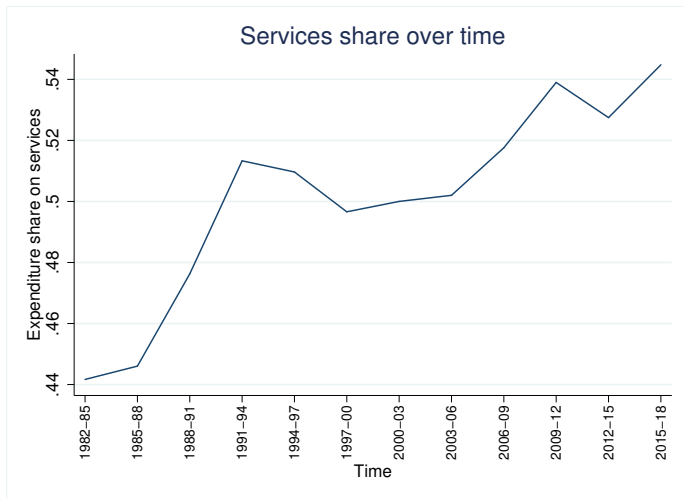
Frequency of price adjustment across age groups



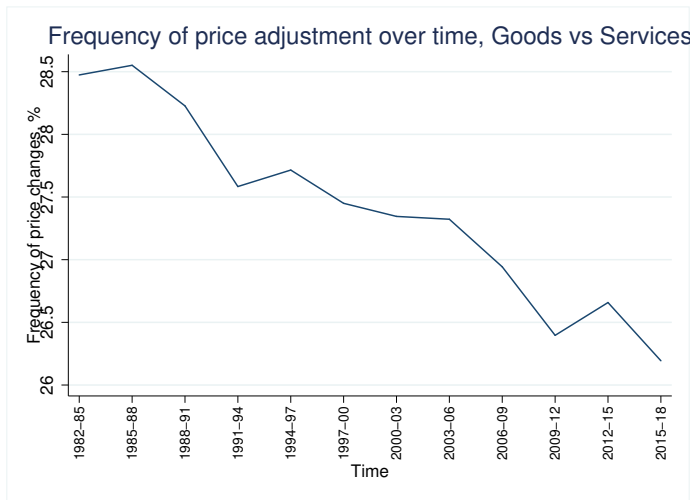


Services vs Goods, share

$\theta_{Services}$: 14, θ_{Goods} : 39



Services vs Goods, price stickiness



Age group level inflation rate

The BLS procedure to compute CPI consists in calculating:

$$X_j^a = \frac{\sum_i fwt_i^a \sum_t c_{i,j,t}^a}{\sum_i fwt_i^a MO_SCOPE_i^a} \times 12$$

where fwt_i^a is the frequency weight for household i at age group a , $c_{i,j,t}^a$ refers to the annual consumption on category j by household i at age group a and $MO_SCOPE_i^a$ identify the number of months per year household i reported its expenditures.

The age group level expenditure weight for category j can be then computed as:

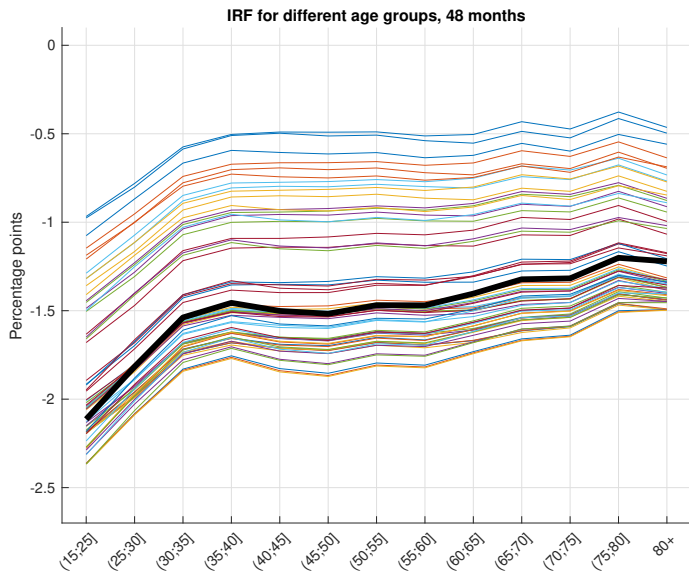
$$w_j^a = \frac{X_j^a}{\sum_j X_j^a}$$

Finally, the the age-group level price index is given by:

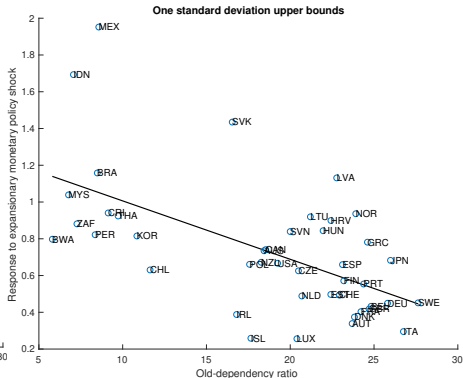
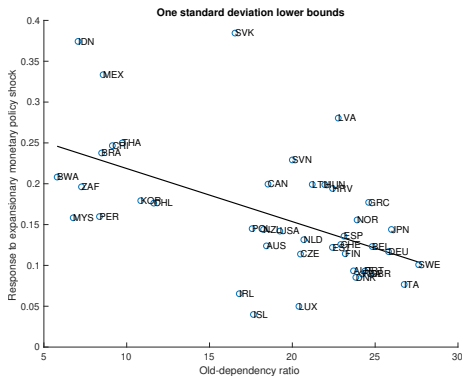
$$CPI_t^a = CPI_\nu^a \sum_j (w_{j,\beta}^a \times \frac{P_{j,t}}{P_{j,\nu}})$$

where ν is the pivot year and month prior to the month when expenditure weights from reference period β are first used, β is the predetermined expenditure reference period, $w_{j,\beta}^a$ is the age group level expenditure weight for category j during the predetermined expenditure reference period β and $P_{j,t}$ is the price of item j at time t . [▶ Back](#)

IRFs by age groups, robustness

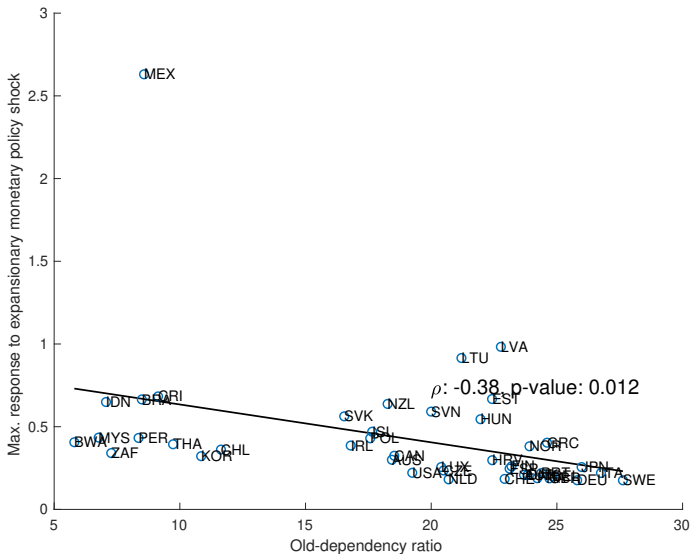


Contemporaneous response of inflation - Confidence bands

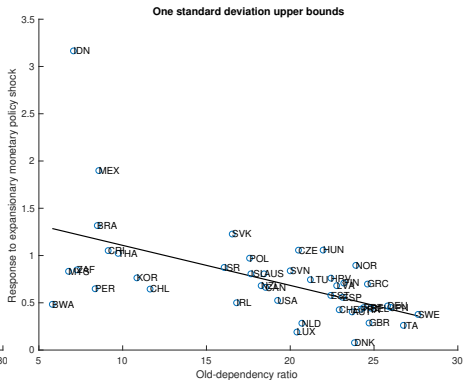
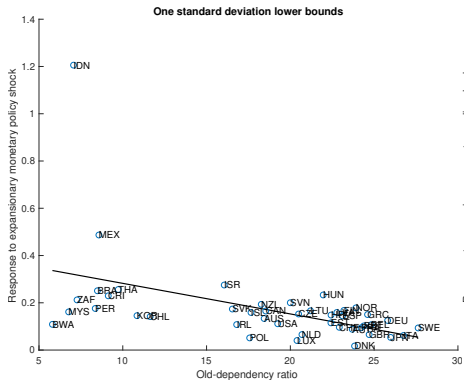


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Maximum response of inflation

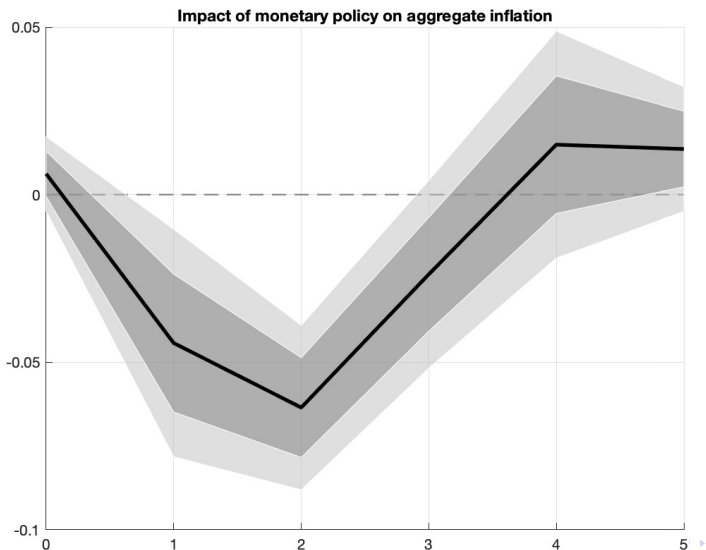


Contemporaneous response of inflation - Confidence bands

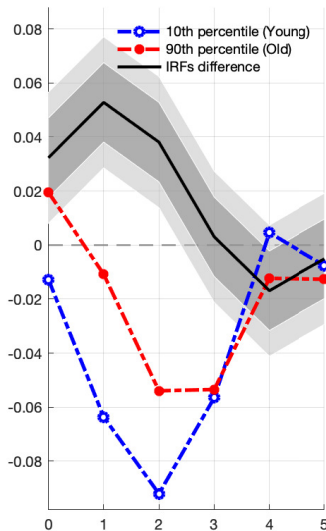
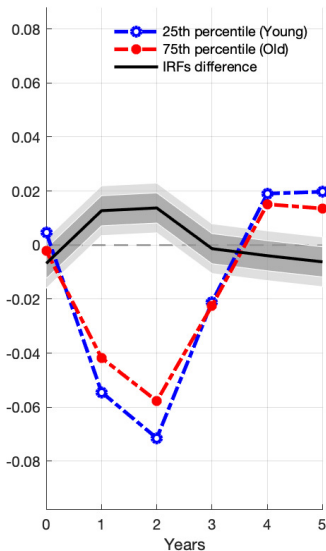


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IRFs for the Euro Area with additional controls, high-frequency shocks



IRFs for the Euro Area with additional controls, high-frequency shocks

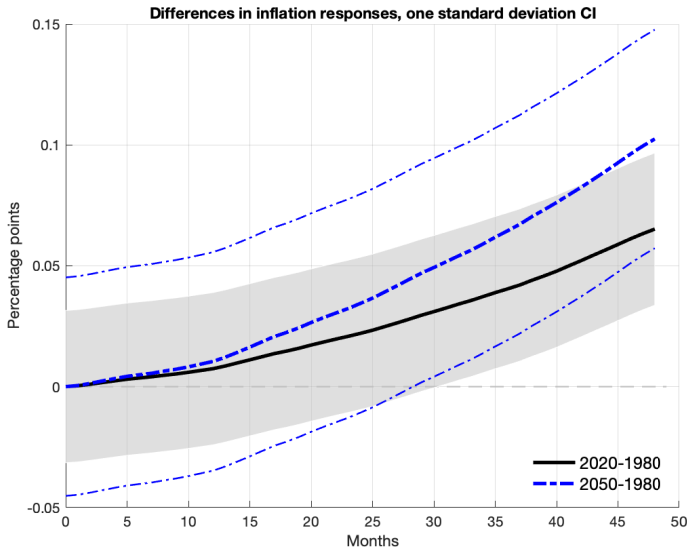


Expenditure weights

	Age groups						
	25-	(30,35]	(40,45]	(50,55]	(60,65]	(70,75]	80+
Alcohol	2.1	1.4	1.2	1.2	1.2	1.1	0.6
Apparel	5.1	4.8	4.7	4.2	3.8	3.1	2.3
Education	6.7	1.5	2.4	3.9	1.0	0.6	0.4
Energy	3.8	5.0	5.4	5.5	6.0	6.7	7.9
Entertainment	5.9	7.0	7.5	6.9	6.8	6.0	4.4
Food Away	6.1	5.6	5.8	5.8	5.6	5.1	4.1
Food at Home	11.4	12.5	13.0	12.1	12.3	12.9	13.5
Medical	3.4	5.4	6.4	7.6	10.7	15.1	19.0
Household F&O	6.4	9.9	9.1	9.0	9.8	10.1	11.1
Other Lodging	1.2	1.0	1.4	2.0	1.8	2.0	0.9
Owned Dwellings	1.8	6.5	7.5	7.7	8.1	7.6	5.9
Other Expenses	0.9	1.1	1.3	1.4	1.6	1.8	2.4
Personal Care	1.9	1.9	2.0	1.9	1.9	2.0	2.1
Private Transportation	20.5	21.8	21.7	21.6	20.8	17.5	11.3
Public Transportation	1.2	1.3	1.4	1.5	1.8	1.7	1.1
Reading	0.3	0.4	0.4	0.5	0.6	0.7	0.7
Rented Dwellings	19.4	10.8	6.4	4.4	3.7	3.9	10.2
Tobacco	1.3	1.0	1.1	1.2	1.1	0.8	0.4
Water	0.6	1.1	1.2	1.2	1.3	1.5	1.7

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Differences in IRFs



Summary statistics

	Mean	St. dev.	Min.	Max.	Obs.
Population	2.916	3.175	26	17.885	1.240
Employment	1.330	1.581	16	9.299	1.240
GDP	85.376	112.197	1.046	651.455	1.240
Share services	0,70	0,08	0,52	0,89	1.240

Note: The data are sourced from the European Regional Database of Cambridge Econometrics and cover the period 1999-2016. Population and employment are in thousands of people. The GDP and the share of services in GVA are valued at market prices before being deflated to 2015 constant price euros.

Finland (5 NUTS II Regions)

Regions: Etelä-Suomi, Helsinki-Uusimaa, Länsi-Suomi, Itä-Suomi, Åland

Data source: Statistics Finland

Germany (16 NUTS I Regions)

Regions: Baden-Württemberg, Bayern, Berlin, Brandenburg, Bremen, Hamburg, Hessen, Mecklenburg-Vorpommern, Niedersachsen, Nordrhein-Westfalen, Rheinland-Pfalz, Saarland, Sachsen, Sachsen-Anhalt, Schleswig-Holstein, Thüringen

Data source: Statistical offices of the individual German states

Italy (22 NUTS II Regions)

Regions: Abruzzo, Basilicata, Calabria, Campania, Emilia-Romagna, Friuli-Venezia Giulia, Lazio, Liguria, Lombardia, Marche, Molise, Piemonte, Provincia Autonoma di Bolzano, Provincia Autonoma di Trento, Puglia, Sardegna, Sicilia, Tirol, Toscana, Umbria, Valle d'Aosta, Veneto

Data source: Istituto Nazionale di Statistica (ISTAT)

Portugal (7 NUTS II Regions)

Regions: Acores, Algarve, Alentejo, Centro, Lisboa, Madeira, Norte

Data source: Instituto Nacional de Estatística (INE)

Spain (19 NUTS II Regions)

Regions: Andalucía, Aragón, Canarias, Cantabria, Castilla y León, Castilla La Mancha, Cataluña, Ciudad Autónoma de Ceuta, Ciudad Autónoma de Melilla, Comunidad Foral de Navarra, Comunidad Valenciana, Comunidad de Madrid, Extremadura, Galicia, Illes Balears, La Rioja, País Vasco, Principado de Asturias, Región de Murcia

Data source: Instituto Nacional de Estadística (INE)

Australia (1985Q1-2015Q4), Austria (1996Q1-2015Q4), Belgium (1985Q1-2015Q4), Canada (1985Q1-2015Q4), Costa Rica (1997Q1-2015Q4), Czech Republic (1996Q1-2015Q4), France (1985Q1-2015Q4), Germany (1991Q1-2015Q4), Greece (1995Q1-2015Q4), Hungary (1995Q1-2015Q4), Iceland (1995Q1-2015Q4), Portugal (1995Q1-2015Q4), Israel (1991Q1-2015Q4), Italy (1985Q1-2015Q4), Lithuania (1999Q1-2015Q4), Luxembourg (1999Q1-2015Q4), Latvia (1998Q1-2015Q4), Poland (1995Q1-2015Q4), Netherlands (1995Q1-2015Q4), New Zealand (1985Q1-2015Q4), Norway (1985Q1-2015Q4), Slovakia (1995Q1-2015Q4), Switzerland (1985Q1-2015Q4), United Kingdom (1985Q1-2015Q4), United States (1985Q1-2015Q4), South Africa (1993Q1-2015Q4): The interest rate is the short term rate from the OECD, the CPI is from the OECD, and the seasonally-adjusted real GDP comes from the IFS.

Mexico (1985Q1-2015Q4), Indonesia (2000Q1-2015Q4), Ireland (1995Q1-2015Q4), Denmark (1995Q1-2015Q4), Finland (1990Q1-2015Q4), Japan (1985Q1-2015Q4), Korea (1985Q1-2015Q4), Spain (1995Q1-2015Q4) and Sweden (1993Q1-2015Q4): The interest rate is the money market rate from the IFS, the CPI is from the OECD, and the seasonally-adjusted real GDP comes from the IFS.

Brasil (1996Q1-2015Q4): The interest rate is from the [Ilzetzki et al. \(2013\)](#) dataset. The CPI is from the OECD and the seasonally-adjusted real GDP comes from the IFS.

Data for Botswana (1993Q1-2015Q4), Chile (1996Q1-2015Q4), Croatia (1997Q1-2015Q4), Estonia (1999Q1-2015Q4), Peru (1995Q1-2015Q4), Slovenia (1992Q1-2015Q4) and Thailand (1993Q1-2015Q4) are taken from [Ilzetzki et al. \(2013\)](#).

Malaysia (1999Q1-2015Q4): The interest rate is the money market rate from the IFS. Real GDP and CPI from [Ilzetzki et al. \(2013\)](#).