# Automation and Human Capital: Accounting for Individual-Level Responses 

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## Motivation

1. Automation substitutes routine labour (production, clerks, sales) and complements abstract labour (managerial, professional)
2. Automation creates incentives.
for routine $(R)$ workers to accumulate HC , to join abstract occupations for abstract (A) workers to accumulate more of HC R workers can be limited in their mobility towards abstract occupations - lower learning ability/stock of HC.

How human capital responses to automation contribute to life-cycle earnings inequality?

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## What I do

- Empirical Analysis:
- NLSY79 data: show ability-based selection to A and R occupations
- CPS data: estimate price series for A and R labor over the last 4 decades
- Quantitative Analysis:
- Build life-cycle model with HC accumulation and occupational choice
- Calibrate it to NLSY79 cohort, using exogenous A and R price series (CPS)
- Run counterfactuals, fixing $A$ and $R$ prices on 1976 level
- See how HC responses to automation contribute to earnings inequality over the life cycle


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## What I find: Empirical Analysis

## From NLSY79 data:

1. Ability-based selection into $R$ and $A$ occupations
2. Over the life cycle, outflow of workers from $R$ and inflow to $A$ occupations
3. Probability of $R \rightarrow A \& A \rightarrow R$ switches is ability-dependent

From CPS data:
4. Price of $R$ labor $\downarrow$ and price of A labor $\uparrow$ over the last 4 decades

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From CPS data:
4. Price of $R$ labor $\downarrow$ and price of A labor $\uparrow$ over the last 4 decades

- Different from price series estimates for high-, mid-, and low-skilled labor


## What I find: Quantitative Analysis

1. Modest contribution of automation to log-earnings variance

- Up to $10.8 \%$ by the end of the working life cycle
- Mostly due to a change in prices for HC in $R$ and $A$

2. Significant contribution of automation to abstract wage premium $\frac{\operatorname{avg}\left(\text { wage }_{A}\right)}{\operatorname{avg}\left(\text { wage }_{R}\right)} \uparrow=\frac{\operatorname{avg}\left(\text { P Fice }_{A} \uparrow \times H C_{A}\right)}{\operatorname{avg}\left(\text { Price }_{R} \downarrow \times H C_{R}\right)}$

- Up to $28.6 \%$ of a rise is due to automation

3. HC responses and $\mathrm{R} \rightarrow \mathrm{A}$ switches dampen a rise in abstract wage premium: $\frac{\operatorname{avg}\left(\text { wage }_{A}\right)}{\operatorname{avg}\left(\text { wage }_{R}\right)} \downarrow=\frac{\operatorname{avg}\left(\text { Price }_{A} \times H C_{A} \downarrow\right)}{\operatorname{avg}\left(\text { Price }_{R} \times H C_{R} \uparrow\right)}$

- The premium would be 35.5 p.p. higher in the absence of HC responses


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## Related Literature

- Life-Cycle Inequality: Huggett, Ventura, and Yaron (2006, 2011), Storesletten, Telmer, and Yaron (2004)
- This paper: workers with different ability/HC respond differently to automation
- ...and this contributes to a rise in life-cycle earnings inequality
- Occupational Switching: Cortes (2016), Autor and Dorn (2009)
- This paper: models the reasons underlying selection into $A$ and $R$ occupations
- Quantity vs. Price of HC: Bowlus and Robinson (2012), Heckman, Lochner, and Taber (1998)
- This paper: shows that price series for HC in A and R diverge over time


## Outline

- Empirical Analysis
- Model Description
- Calibration and Model Fit
- Counterfactual Experiments


## Empirical Analysis

## Ability-Based Selection

- AFQT scores from NLSY79 data


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Occupational Distributions by Ability Quartiles

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Occupational Distributions by Ability Quartiles


- Ability predicts sorting across $A$ and $R$ occupations
- $45 \%$ of workers in A occupations are from the top ability quartile
- $32 \%$ of workers in R occupations are from the lowest ability quartile
- AFQT scores still predict allocation to occupations after 2.5 decades


## Occupational Switches

- Ability predicts switching between $A$ and $R$ occupations


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Occupational Switch Probabilities by Ability Quartiles

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Occupational Switch Probabilities by Ability Quartiles


- More able agents are more likely to go to A occupations
- Young least able are 7 times less likely to do RA switch than the most able ones
- Less able agents are more likely to fall to R occupations
- Switch probability decreases with age


## Outflow of Workers from R Occupations

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Occupational share of $R$ workers over the life cycle


- $19 \%$ fall in R workers share offset by a rise in A workers share
- R to A switchers earn more in 10 years than those staying R

1. Ability predicts individuals' capacity to adjust to automation

- Less able agents are more limited in their upward mobility

2. High share of low-ability individuals in $\mathbf{R}$ occupations

- Significant share is unable to respond to automation

3. High share of high-ability individuals in $\mathbf{A}$ occupations

- Potentially accumulate more of human capital with automation

4. Workers upgrade from $\mathbf{R}$ to $\mathbf{A}$ occupations over the life cycle

- Potentially dampens the effect of automation on earnings inequality

How human capital responses to automation contribute to life-cycle earnings inequality?

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## Model Description

## Model Overview

- Partial equilibrium, perfect foresight
- Prices for R and A change exogenously
- Endogenous HC accumulation in A occupations, Ben-Porath type
- Agents living for J periods
- Ex-ante heterogeneous in:

1. Learning ability
2. Initial HC in A occupation
3. Productivity in R occupation

- Time and monetary investment into human capital in A occupation
- HC stock in $A$ and $R$ hit by the idiosyncratic shocks
- Working in either A or R occupation


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## Agent's Problem

- Heterogeneous in learning ability a, R productivity $\eta$, and initial HC in A $h_{A, 1}$


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- Maximize lifetime utility, linear in consumption:

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\begin{equation*}
\max _{\left\{c_{j}, o c c_{j}, l_{j}, n_{j}, d_{j}, h_{j+1}\right\}_{j=1}^{J}} E\left[\sum_{j=1}^{J} \beta^{j-1} c_{j}\right] \tag{1}
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- Labor earnings go to consumption and monetary investment into $h_{A, r}$, (BC):

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c_{j}+d_{j}=y_{j} \tag{2}
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- Work in either A or R occupation:

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\begin{equation*}
y_{j}=P_{k, t}\left(\exp \left(z_{k, j}\right) h_{k, j} l_{j}\right), \text { where } k \in\{A, R\} \tag{3}
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- Note 1: the model follows one cohort over its life cycle.
- Note 2: automation is introduced through:
- $P_{R, t} \downarrow$ and $P_{A, t} \uparrow$
- $P_{R, t}$ and $P_{A, t}$ are time-dependent, not age-dependent


## Occupational Choice and HC Accumulation

- The choice between occupations:

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o c c_{j}=A \text { if } h_{A, j} \geq \frac{P_{R, t} \exp \left(z_{R, j}\right)}{P_{A, t} \exp \left(z_{A, j}\right)} h_{R, j} \text { and occ } c_{j}=R \text { otherwise } \tag{4}
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- Unit endowment of time in each period $j$ :

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## Calibration and Model Fit

## R and A Prices

- Estimate $P_{A, t}$ and $P_{R, t}$ using a "flat spot" approach
- At older ages, changes in wages are due to changes in prices
$\operatorname{Mean}\left[\ln h_{k, j+1, t+1}\right]=\operatorname{Mean}\left[\ln h_{k, j, t}\right] \Longrightarrow \operatorname{Mean}\left[\ln P_{k, t+1} h_{k, j+1, t+1}\right]-\operatorname{Mean}\left[\ln P_{k, t} h_{k, j, t}\right]$ $=\ln P_{k, t+1}-\ln P_{k, t}$, where $k \in\{A, R\}$
- Applied to CPS data for 1976-2019; medians instead of means because of topcoding
- College grads aged 50-58 for A; high-school grads aged 46-55 for R


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## The Rest of Parameters

- HC accumulation in $\mathbf{R}$ occupations: $f(j)$ as age effect for $\mathbf{R}$ workers from PSID
- Number of lifetime periods $\mathrm{J}=41$, from real age of 18 to 58 and $\beta=0.96$
- Initial conditions: $\left(h_{0}, a, \eta\right) \sim \operatorname{LN}\left(\mu_{x}, \Sigma\right)$
- UC shock: i.i.d., $z_{j} \sim N\left(\mu, \sigma^{2}\right)$
- Calibrate initial distribution, shock, and $\alpha_{1}, \alpha_{2}$ to match the moments from NLSY79:

1. Variance of log-earnings
2. Abstract wage premium
3. Occupational distributions by ability quartiles
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## Earnings Stats and Ability Distributions



- U-shape of variance profile due to:

1. High ability workers accumulating HC at the beginning of the life cycle
2. High ability workers earning more later in the life cycle

## RA mobility





| Data |
| :---: |
| Model |

- RA mobility due to:

1. Medium to high a agents for whom $h_{A, j} \leq \frac{P_{R, t} h_{R, j}}{P_{A, t}}$ for the first $n$ periods, but who invest into $h_{A, j}$ and switch to A in $n+1$ period
2. Positive shocks to HC in A

## AR mobility





| $\square$ |
| :---: |
| $\square$ |
|  |
| $\square$ | Mata

- AR mobility due to negative realizations of HC shocks
- Ability-based selection due to strong positive $a$ and $h_{A, 1}$ correlation


## Counterfactual Exercises

## Changes in HC Prices



- No automation: setting $P_{A, t}=P_{A, 1976}$ and $P_{R, t}=P_{R, 1976}, \forall t$
- Change in HC prices contribution:
- Up to $10.8 \%$ to the variance
- Up to $28.6 \%$ to the abstract wage premium


## Contribution of HC responses



- No HC response:
- Using estimated $P_{A, t}, P_{R, t}$, but keeping policies optimal under no change in prices
- HC responses and RA switches dampen a rise in abstract wage premium
- Without HC responses, abstract wage premium would be up to 35.5 p.p. higher


## RA switches





| $\square$ |
| :--- |
| Full Model |
| $\square$ |
| $\square$ |

- With automation, RA mobility is higher across all ability quartiles
- Across all ability quartiles - more intensive accumulation of HC in A


## Conclusion

- Empirical analysis
- Ability-based selection into $R$ and $A$ occupations
- Probability of RA and AR switches is ability-dependent
- Price of $\mathrm{R} \downarrow$ labor and $\mathrm{A} \uparrow$ labor over the last 4 decades
- Quantitative analysis
- Modest contribution of automation to variance of log-earnings
- Significant contribution of automation to abstract wage premium
- HC responses and RA switches dampen a rise in the premium over life cycle


## Appendix 1: NLSY79 Sample

Table A1: NLSY79 Sample of Males by Age and Occupational Categories

| Observations/Age | $23-27$ | $28-32$ | $33-37$ | $38-42$ | $43-47$ | $48-52$ | $53-57$ | Total |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Total | 6,117 | 5,926 | 5,404 | 4,771 | 4,402 | 4,070 | 1,786 | 32,476 |
| By shares of |  |  |  |  |  |  |  |  |
| occ. categories        <br> Abstract 0.27 0.34 0.38 0.41 0.42 0.43 0.45 <br> Routine 0.63 0.58 0.54 0.50 0.48 0.47 0.46 <br> Service 0.10 0.08 0.08 0.09 0.10 0.10 0.09 |  |  |  |  |  |  |  |  |

Note: The table shows the number of observations and the shares of the three occupational categories by age groups for males from a cross-sectional sample of the NLSY79 data used for the analysis in this paper. Sample restrictions are: yearly working hours 260-5820 and yearly earnings at least $\$ 1000$ for those below 30 y.o., and yearly working hours 520-5820 and yearly earnings at least $\$ 1500$ for those above 30 y.o. (earnings are in 1979 dollars). Such restricted sample of males consists of 3,003 individual observations.

## Labor Income of Switchers to A vs. Stayers in R

Table A3.1: Labor Income across different Occupational Cycles

| Q1 | Q2 | Q3 | Q4 |
| :---: | :---: | :---: | :---: |

Panel 1: Routine Occupations
Occ. upgrading (RRA and RAA) vs. staying (RRR)

| Occ. | $0.226^{* * *}$ | 0.055 | $0.214^{* * *}$ | $0.247^{* * *}$ |
| :--- | :---: | :---: | :---: | :---: |
| upgrading | $(0.056)$ | $(0.042)$ | $(0.032)$ | $(0.038)$ |
| Age | $0.084^{* * *}$ | $0.035^{* * *}$ | $0.028^{* *}$ | -0.003 |
|  | $(0.027)$ | $(0.012)$ | $(0.012)$ | $(0.011)$ |
| Age ${ }^{2}$ | $-0.001^{* * *}$ | $-0.001^{* * *}$ | $-0.001^{* * *}$ | 0.000 |
|  | $(0.000)$ | $(0.000)$ | $(0.000)$ | $(0.000)$ |
| Year | -0.001 | $0.026^{* * *}$ | $0.030^{* * *}$ | $0.014^{* *}$ |
|  | $(0.007)$ | $(0.005)$ | $(0.004)$ | $(0.006)$ |
| Nonwhite | $-0.033^{* *}$ | -0.011 | 0.020 | -0.007 |
|  | $(0.015)$ | $(0.015)$ | $(0.026)$ | $(0.033)$ |
| Obs. | 1736 | 2173 | 2165 | 1427 |

Note: Columns Q1-Q4 show the estimated coefficients from a regression of log yearly labor income in $t+10$ on dummies for occupational upgrading and downgrading and a set of listed controls. Occ. upgrading dummy is defined as equal to 1 if individual follows RRA or RAA (upgrading) occupational cycle in $\mathrm{t}, \mathrm{t}+2$, and $\mathrm{t}+10$, respectively and as equal to 0 if individual follows RRR (staying); Occ. downgrading dummy is defined as equal to 1 if individual follows AAR or ARR (downgrading) occupational cycle in $t, t+2$, and $t+10$, respectively and as equal to 0 if individual follows AAA (staying). Robust s.e. in parentheses, ${ }^{*} p<0.1,{ }^{* *} p<0.05$,

$$
* * * p<0.01 \text { back }
$$

## Labor Income of Switchers to R vs. Stayers in A

Table A3.2: Labor Income across different Occupational Cycles
Q1

Panel 2: Abstract Occupations
Occ. downgrading (AAR and ARR) vs. staying (AAA)

| Occ. | $-0.327^{* * *}$ | $-0.267^{* * *}$ | $-0.285^{* * *}$ | $-0.475^{* * *}$ |
| :--- | :---: | :---: | :---: | :---: |
| downgrading | $(0.116)$ | $(0.064)$ | $(0.045)$ | $(0.050)$ |
| Age | -0.055 | $0.086^{* * *}$ | 0.026 | $0.043^{* * *}$ |
|  | $(0.097)$ | $(0.026)$ | $(0.017)$ | $(0.014)$ |
| Age ${ }^{2}$ | 0.001 | $-0.001^{* * *}$ | -0.000 | $-0.001^{* * *}$ |
|  | $(0.001)$ | $(0.000)$ | $(0.000)$ | $(0.000)$ |
| Year | -0.004 | $0.020^{* *}$ | $0.012^{*}$ | $0.024^{* * *}$ |
|  | $(0.023)$ | $(0.008)$ | $(0.007)$ | $(0.004)$ |
| Nonwhite | -0.047 | -0.021 | $0.061^{* * *}$ | -0.039 |
|  | $(0.043)$ | $(0.033)$ | $(0.019)$ | $(0.026)$ |
| Obs. | 223 | 612 | 1577 | 2947 |

Note: Columns Q1-Q4 show the estimated coefficients from a regression of log yearly labor income in $t+10$ on dummies for occupational upgrading and downgrading and a set of listed controls. Occ. upgrading dummy is defined as equal to 1 if individual follows RRA or RAA (upgrading) occupational cycle in $t, t+2$, and $t+10$, respectively and as equal to 0 if individual follows RRR (staying); Occ. downgrading dummy is defined as equal to 1 if individual follows AAR or ARR (downgrading) occupational cycle in $t, t+2$, and $t+10$, respectively and as equal to 0 if individual follows AAA (staying). Robust s.e. in parentheses, ${ }^{*} p<0.1,{ }^{* *} p<0.05$,

$$
* * * p<0.01 \text { back }
$$

## Outflow from A occupations

- Longitudinal ASEC CPS data
- The lowest outflow from $A$ is among college workers

- Share of A workers in $t-1$ observed out of A in $t$
- Out of A: R, S, unemployment, nilf

```
back
```


## Outflow from R occupations

- The lowest outflow from R is among high school workers

- Share of R workers in $t-1$ observed out of R in $t$
- Out of R: A, S, unemployment, nilf


## Inflow to A Occupations

- The lowest inflow to A is among college workers

- Share of A workers in $t$ observed out of A in $t-1$
- Out of $\mathrm{A}: \mathrm{R}, \mathrm{S}$, unemployment, nilf


## Inflow to R Occupations

- The lowest inflow to R is among high school workers

- Share of R workers in $t$ observed out of R in $t-1$
- Out of R: A, S, unemployment, nilf


## Wages of Stayers in A vs. Wages of those Joining/Leaving A

- Joining/Leaving A occupations do not statistically differ in wages from Stayers

| Dep.: Log Hourly Wage | Col | Some Col | HS | Col | Some Col | HS |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | $0.005^{* * *}$ | -0.000 | 0.000 | $0.005^{* * *}$ | $-0.002^{* * *}$ | $-0.002^{* *}$ |
|  | $(0.000)$ | $(0.001)$ | $(0.001)$ | $(0.000)$ | $(0.001)$ | $(0.001)$ |
| Joining A | -3.674 | -0.746 | $5.616^{*}$ |  |  |  |
|  | $(4.123)$ | $(3.541)$ | $(2.995)$ |  |  |  |
| Joining A $\times$ Year | 0.002 | 0.000 | $-0.003^{*}$ |  |  |  |
|  | $(0.002)$ | $(0.002)$ | $(0.001)$ |  |  |  |
| Age | -0.003 | 0.000 | 0.003 | -0.001 | $0.006^{* *}$ | $0.010^{* * *}$ |
|  | $(0.002)$ | $(0.003)$ | $(0.003)$ | $(0.002)$ | $(0.003)$ | $(0.003)$ |
| Leaving A |  |  |  | 2.208 | 3.429 | 0.182 |
|  |  |  |  | $(3.483)$ | $(3.639)$ | $(2.812)$ |
| Leaving A $\times$ Year |  |  |  | -0.001 | -0.002 | -0.000 |
|  | $-5.884^{* * *}$ | $3.645^{* *}$ | 2.485 | $-6.447^{* * *}$ | $6.7000^{* * *}$ | $5.485^{* * *}$ |
| Constant | $(0.892)$ | $(1.488)$ | $(1.615)$ | $(0.859)$ | $(1.420)$ | $(1.396)$ |
|  | 21,648 | 8,624 | 6,777 | 22,206 | 8,944 | 7,020 |
| Observations | 0.011 | 0.006 | 0.002 | 0.009 | 0.007 | 0.004 |
| $R^{2}$ |  |  |  |  |  |  |

- Longitudinal ASEC CPS data
- Joining $=1$ if out of A in $t-1$ and in A in $t$; Joining $=0$ if in A in $t-1$ and in A in $t$
- Leaving $=1$ if in A in $t-1$ and out of A in $t$; Leaving $=0$ if in A in $t-1$ and in A in $t$


## Wages of Stayers in $R$ vs. Wages of those Joining/Leaving $R$

- Joining/Leaving A occupations do not statistically differ in wages from Stayers

| Dep.: Log Hourly Wage | Col | Some Col | HS | Col | Some Col | HS |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | $0.002^{*}$ | $-0.003^{* * *}$ | $-0.006^{* * *}$ | 0.001 | $-0.003^{* * *}$ | $-0.007^{* * *}$ |
|  | $(0.001)$ | $(0.001)$ | $(0.000)$ | $(0.001)$ | $(0.001)$ | $(0.000)$ |
| Joining R | 2.274 | 1.582 | -0.422 |  |  |  |
|  | $(4.611)$ | $(2.969)$ | $(2.108)$ |  |  |  |
| Joining R $\times$ Year | -0.001 | -0.001 | 0.000 |  |  |  |
|  | $(0.002)$ | $(0.001)$ | $(0.001)$ |  |  | 0.001 |
| Age | $-0.011^{* *}$ | $-0.007^{* * *}$ | $0.003^{*}$ | $-0.018^{* * *}$ | -0.002 | $0.001)$ |
|  | $(0.005)$ | $(0.002)$ | $(0.001)$ | $(0.005)$ | $(0.002)$ | $(0.001)$ |
| Leaving R |  |  |  | -5.369 | 0.725 | -1.925 |
|  |  |  |  | $(5.414)$ | $(2.889)$ | $(2.272)$ |
| Leaving R $\times$ Year |  |  |  | 0.003 | -0.000 | 0.001 |
|  |  |  |  | $(0.003)$ | $(0.001)$ | $(0.001)$ |
| Constant | -0.975 | $8.906^{* * *}$ | $14.608^{* * *}$ | 2.059 | $9.481^{* * *}$ | $17.188^{* * *}$ |
|  | $(2.604)$ | $(1.312)$ | $(0.757)$ | $(2.821)$ | $(1.347)$ | $(0.755)$ |
| Observations | 4496 | 10944 | 22552 | 4292 | 10997 | 23048 |
| $R^{2}$ | 0.007 | 0.004 | 0.013 | 0.016 | 0.004 | 0.017 |

- Joining $=1$ if out of R in $t-1$ and in R in $t$; Joining $=0$ if in R in $t-1$ and in R in $t$
- Leaving $=1$ if in R in $t-1$ and out of R in $t$; Leaving $=0$ if in R in $t-1$ and in R in $t$ back


## Age Profile in $R$ occupations


back

$$
\begin{gathered}
\log \left(y_{a g e, t}\right)=\beta_{0}+\beta_{1} \text { age }+\beta_{2} a g e^{2}+\gamma_{1} t+\gamma_{2} t^{2}+\epsilon_{j, t} \\
f(j)=\beta_{0}+\beta_{1} j+\beta_{2} j^{2}=-1.09+0.1523 j-0.0017 j^{2}, \text { where } j \in[18,68]
\end{gathered}
$$

## No shocks

RA mobility in the data vs. mobility in the model


## A and R Ability Distributions



## Variance Decomposition

> Table: Variance of log-Earnings in the Models with Different Sources of Earnings Variation

| Model | 25 | 35 | 45 | 55 |
| :--- | :---: | :---: | :---: | :---: |
|  | 0.64 | 0.36 | 0.59 | 0.69 |
| No growth in Prices |  |  |  |  |
|  | 0.59 | 0.38 | 0.54 | 0.62 |
| No shocks | $(0.92)$ | $(1.06)$ | $(0.91)$ | $(0.89)$ |
|  | 0.54 | 0.19 | 0.46 | 0.54 |
| No variation in initial conditions | $(0.85)$ | $(0.53)$ | $(0.79)$ | $(0.78)$ |
|  | 0.09 | 0.09 | 0.13 | 0.18 |
|  | $(0.14)$ | $(0.24)$ | $(0.2)$ | $(0.26)$ |

Note: Full model - the baseline calibration; No growth in prices - prices for human capital in abstract and routine occupations are fixed at the 1979 level; No shocks - the variance of shocks to human capital in abstract and routine occupations is set to 0 ; No variation in initial
conditions - $a, h_{A, j}$, and $\eta$ are set to the mean values of the calibrated distributions for all agents. Values in brackets show the share of the Full model variance produced by each model.

## Calibration: Empirical Moments

## Earnings Stats




Each line is using age effects $\beta_{j}$ from: stat $_{j, t}=\mu^{\text {stat }}+\alpha_{c}^{s t a t}+\beta_{j}^{\text {stat }}+\epsilon_{j, t}^{s t a t}$

## Calibration: Parameter Values

## Definition

Discount factor
Length of the life cycle
Abstract HC prices
Routine HC prices
Age premium
in routine job
HC elasticities

Initial conditions

Abstract HC shocks
Routine HC shocks
Price ratio in $\mathrm{j}=1$

Symbol
$\beta$
$J$
$P_{A, j}$
$P_{R, j}$
$f(j)$

$$
\begin{gathered}
\alpha_{1}, \alpha_{2} \\
\left(h_{0}, a, \eta\right) \sim L N\left(\mu_{x}, \Sigma\right)
\end{gathered}
$$

## Value

$$
0.9615
$$

$$
41
$$

$$
[1,1.18]
$$

$$
[0.80,1]
$$

$$
f(j)=-1.09+0.1523 j-0.0017 j^{2}
$$

$$
0.61,0.15
$$

$$
\left(\mu_{h}, \mu_{\mathrm{a}}, \mu_{\eta}\right)=(4.77,-1.50,5.23)
$$

$$
z \sim N\left(\mu_{A}, \sigma_{A}^{2}\right)
$$

$$
z \sim N\left(\mu_{R}, \sigma_{R}^{2}\right)
$$

$$
P_{R, 1976} / P_{A, 1976}
$$

$$
\begin{aligned}
{\left[\begin{array}{ccc}
\sigma_{h}^{2} & \sigma_{h a} & \sigma_{h \eta} \\
\sigma_{a h} & \sigma_{a}^{2} & \sigma_{a \eta} \\
\sigma_{\eta h} & \sigma_{\eta a} & \sigma_{\eta}^{2}
\end{array}\right] } & =\left[\begin{array}{lll}
0.62 & 0.19 & 0.33 \\
0.19 & 0.29 & 0.14 \\
0.33 & 0.14 & 0.55
\end{array}\right] \\
\left(\mu_{A}, \sigma_{A}\right) & =(0,0.07) \\
\left(\mu_{R}, \sigma_{R}\right) & =(0,0.09) \\
& 0.70
\end{aligned}
$$

## HC Responses Across Ability Quartiles



3rd Quartile


2nd Quartile


4th Quartile


